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Working Paper 16-109

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# Creativity Under Fire:
# The Effects of Competition on Creative Production

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**Abstract:**

Though fundamental to innovation and essential to many industries and occupations, the creative act has received limited attention as an economic behavior and has historically proven difficult to study. This paper studies the incentive effects of competition on individuals' creative production. Using a sample of commercial logo design competitions, and a novel, content-based measure of originality, I find that competition has an inverted-U effect on creativity: some competition is necessary to induce agents to produce radically novel, untested ideas over incrementally tweaking their earlier work, but heavy competition drives them to stop investing altogether. The results are consistent with economic theory and reconcile conflicting evidence from an extensive literature on the effects of competition on innovation, with implications for R&D policy, competition policy, and organizations in creative or research industries.

**JEL Classification:** D81, D82, D83, L4, M52, M55, O31, O32

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The creative act is among the most important yet least understood phenomena in economics and the social and cognitive sciences. Technological progress at its heart consists of creative solutions to familiar problems. Millions of workers in the U.S. alone are employed in creative fields such as research, software development, and media and the arts, and surveys show that CEOs' top concerns consistently include creativity and innovation in the firm. Despite its importance, the creative act itself has received only limited attention in economics and has historically proven difficult to study, due to a dearth of data on creative behavior and exogenous variation in incentives that can be used to understand what drives it.

This paper studies the incentive effects of competition on creative production, exploiting a unique setting where creative activity and competition can be precisely measured and related: tournaments for the design of commercial graphics, logos, and branding. Using image comparison tools to measure originality, I provide causal evidence that competition can both create and destroy incentives for individuals to innovate. While some competition is necessary for high-performing agents to prefer developing novel, untested ideas over tweaking their existing work, heavy competition discourages effort of either kind. These patterns are driven by risk-return tradeoffs inherent to innovation, which I show to be high-risk, high-return. The implication of these results is an inverted-U shaped effect of competition on individual creativity, with incentives for producing new ideas greatest in the presence of one equal-quality competitor.

Over the last three decades, social scientists have coalesced around a definition of creativity as the act of producing ideas that are novel and appropriate to the goal at hand (Amabile 1996, Sternberg 2008).[1] The challenge of motivating this behavior is naturally cast as a principal-agent problem. Suppose a firm would like its workers to experiment with new, potentially better (lower cost or higher quality) product designs, but the firm cannot measure creative choices and can only reward workers on the quality of their output. In this setting, failed experiments are indistinguishable from shirking. Workers who are risk-averse or face decreasing returns to improvement, as they do in this paper, may then prefer exploiting existing solutions over exploring new ones if the existing method reliably yields an acceptable result.

To contextualize this problem in a competitive setting, the paper begins by developing a model of a winner-take-all innovation tournament.[2] In the model, a principal seeks a high-value product design from a pool of workers and solicits ideas through a tournament mechanism, awarding a prize to the best entry. Workers compete for the prize by entering designs in turns. At each turn, a worker must choose between developing an original design, tweaking an existing design, or declining to invest any further. Each submission receives immediate, public feedback on its quality, and at the end of the contest, the sponsor selects the winner.

[1] As Sternberg (2008, p. 468) writes, creativity is "the process of producing something that is both original and worthwhile." This process requires not only cognitive effort, but also a deliberate choice to be original, which is the focus of this paper.

[2] The model in this paper is most closely related to Taylor (1995), Che and Gale (2003), Fullerton and McAfee (1999), and Terwiesch and Xu (2008) but differs in that it injects an explore-exploit dilemma into the agents' choice set: whereas existing work models competing agents who must choose how much effort to exert, the agents in this paper must also choose whether to build off of an old idea or try a new one, much like a choice between incremental versus radical innovation. The model also has ties to recent work on tournaments with feedback (e.g., Ederer 2010), bandit problems in single-agent settings (Manso 2011), and models of competing firms' choice over R&D project risk (Cabral 2003, Rosen 1991).

The model establishes an inverted-U effect of competition on creative behavior, with agents' incentives for generating original ideas maximized at intermediate levels of competition.[3]

I bring the theoretical intuition to a sample of logo design competitions similar to the model's setting, drawing a sample of contests from a popular online platform.[4] In these contests, a firm ("sponsor") solicits custom designs from a community of freelance designers ("players") in exchange for a winner-take-all prize. The contests in the sample offer prizes of a few hundred dollars and on average attract around 35 players and 100 designs. An important feature of this setting is that the sponsor can provide interim feedback on players' designs in the form of 1- to 5-star ratings. These ratings allow players to gauge the quality of their own work and of competing submissions. Most importantly, the dataset also includes the designs themselves, allowing me to study creative choices over the course of a contest: I use image comparison algorithms similar to those used by commercial content-based image retrieval software (e.g., Google Image Search) to calculate similarity scores between pairs of images in a contest, which I then use to quantify the originality of each design relative to prior submissions both by that player and by her competitors.

This setting presents a unique opportunity to directly observe the creative process, and its interaction with incentives, in the field. Though commercial advertising is important in its own right – advertising is a $165 billion industry in the U.S. and a $536 billion industry worldwide[5] – the design process observed here is similar to that in other settings where new products are developed, tested, and refined. The setting can be interpreted as a competitive workplace, but the mechanism itself is fundamentally a request for proposals (RFP), a procurement device widely used by government agencies and firms to solicit ideas, systems, and technologies, often in multiple rounds, with typically only the top bid rewarded.

The sponsors' ratings are critical in this paper as a source of variation in the information that both I and the players have about the state of the competition. Using these ratings, I am able to directly estimate a player's probability of winning, and the results establish that ratings are meaningful: a five-star design generates the same increase in a player's win probability as 10 four-star designs, 100 three-star designs, and nearly 2,000 one-star designs. Data on the time at which designs are entered by players and rated by sponsors enables me to determine what every participant knows at each point in time – and what they have yet to find out. To obtain causal estimates of the effects of feedback and competition, I exploit naturally-occurring, quasi-random variation in the timing of sponsors' ratings, and identify the effects of ratings that the designers can observe. The empirical strategy effectively compares players' responses to information they observe at the time of design against that which is absent or has not yet been provided.

[3] These results concord with previous results from the tournament literature finding that asymmetries discourage effort (e.g., Baik 1994, Brown 2011). The contribution of this model is to inject an explore-exploit problem into the problem, effectively adding a new margin along which effort may vary. Interestingly, the empirical evidence breaks slightly with this prediction: while distant laggards do reduce their investment, distant leaders are more likely to tweak than abandon.

[4] The empirical setting is conceptually similar to coding competitions studied by Boudreau et al. (2011), Boudreau et al. (2016), and Boudreau and Lakhani (2015), though the opportunity to measure originality is unique. Ulrich and Wooten (2013, 2014) have also studied graphic design competitions, focusing on the effects of visibility and feedback on performance.

[5] Data from Magna Global Advertising Forecast for 2015, available at http://news.magnaglobal.com/.

I find that feedback and competition have large effects on creative choices. In the absence of competition, positive feedback causes players to cut back sharply on experimentation: players with the top rating enter designs that are one full standard deviation more similar to their previous work than those who have only low ratings. The effect is strongest when a player receives her first five-star rating in a contest – her next design will be a near replica of the high-rated design, on average *three* standard deviations more similar to it – and attenuates at each rung down the ratings ladder. But these effects are significantly reversed (by half or more) when high-quality competition is present. Intense competition and negative feedback also drive players to stop investing in a contest, as their work is unlikely to win: the probability of doing so increases with each high-rated competitor. In both reduced-form regressions and a choice model, I find that high-performers are most likely to experiment when facing one high-quality competitor.

For players with poor ratings, the data show that starting down a new path clearly dominates imitation of their poor-performing work. But why would a top contender ever deviate from her winning formula? The model suggests that even top contenders may wish to do so when competition is present, provided original designs will in expectation outperform tweaks. To evaluate whether it pays to be creative, I recruit a panel of professional designers to provide independent ratings on all five-star designs in my sample and correlate their responses with these designs' originality. I find that original designs on average receive higher ratings than incremental tweaks but that the distribution of opinion also has higher variance. The results reinforce one of the standard assumptions in the innovation literature – that innovation is high-risk and high-reward – which is the required condition for competition to motivate creativity.

To my knowledge, this paper provides the most direct view into the creative act to-date in the economics literature. The phenomenon is a classic example of a black box: we can see the inputs and outputs, but we have little evidence or understanding of what happens in between. Reflecting these data constraints, empirical research has opted to measure innovation in terms of inputs (R&D spending) and outcomes (patents), when innovation is at heart about what goes on in between: individual acts of exploration and discovery. Because experimentation choices cannot be inferred from R&D inputs alone, and because patent data only reveal the successes – and only the subset that are patentable and its owners are willing to disclose – we may know far less about innovation than commonly believed. This paper is an effort to fill this gap.

Although creativity has only recently begun to receive attention in economics,[6] social psychology has a rich tradition in studying the effects of intrinsic and extrinsic motivators on creativity. The consensus from this literature is that creativity is inspired by intrinsic "enjoyment, interest, and personal challenge" (Hennessey and Amabile 2010), and that extrinsic pressures of reward, supervision, evaluation, and competition tend to undermine intrinsic motivation by causing workers to "feel controlled by the situation." The implication

[6] E.g., Azoulay et al. (2011), or Charness and Grieco (2014). Akcigit and Liu (2015) and Halac et al. (2015) study problems more similar to the one in this paper: Akcigit and Liu embed an explore-exploit problem into a two-player patent race and study the efficiency consequences of private information; Halac et al. study the effects of assorted disclosure and prize-sharing policies on experimentation in contests to achieve successful innovation.

is that creativity cannot be managed: any attempts to incentivize creative effort will backfire. Although intrinsic motivation is undoubtedly important to creativity, this paper challenges the rejection of extrinsic motivators with evidence that creative choices respond strongly to incentives.

The evidence that incentives for exploring new ideas are greatest under moderate competition has broader implications for R&D policy, competition policy, and management of organizations in creative and research industries, which I discuss at the end of the paper. The results also provide a partial resolution to the long-standing debate on the effects of competition on innovation, which is summarized by Gilbert (2006) and Cohen (2010). Since Schumpeter's contention that monopoly is most favorable to innovation, researchers have produced explanations for and evidence of positive, negative, and inverted-U relationships between competition and innovation in a variety of markets. The confusion results from disagreements of definition and measurement; ambiguity in the type of competition being studied; problems with econometric identification; and institutional differences, such as whether innovation is appropriable. This paper addresses these issues by establishing clear and precise measures of competition and innovation, identifying the causal effects of *information* about competition on innovation, and focusing the analysis on a winner-take-all market with copyright protections. Moreover, as Gilbert (2006) notes, the literature has largely ignored that individuals are the source of innovation ("discoveries come from creative people"), even if patents get filed by corporations. It is precisely this gap that I seek to fill with the present paper.

The paper proceeds as follows. Section 1 presents the model. Section 2 introduces the empirical setting and describes the identification strategy. Section 3 estimates the effects of competition on originality and participation. In Section 4, I unify these results and show that creativity is greatest with one high-quality competitor. Section 5 establishes that innovation in this setting is high-risk, high-return, confirming the driving assumption of the model. Section 6 concludes with discussion of the implications of these results for policy, management, and future research on creativity and innovation.

# 1 Creative Effort in Innovation Tournaments

Suppose a risk-neutral principal seeks a new product design. Because R&D is risky, and designs are difficult to value, the principal cannot contract on absolute performance. It instead sponsors a tournament to solicit prototypes from $J$ risk-neutral players, who enter designs sequentially and receive immediate, public feedback on their quality. Each design can be either an original submission or adapted from the blueprints of previous entries; players who choose to continue working on a given design post-feedback can re-use the blueprint to create variants, with the original version remaining in contention. At a given turn, the player must choose whether to continue participating and if so, whether to create an original design or tweak an earlier submission. At the end of the tournament, the sponsor awards a fixed, winner-take-all prize $P$ to its favorite entry. The sponsor seeks to maximize the value of the winning design.

To hone intuition, suppose each player enters at most two designs. Let each design be characterized by latent value $\nu_{jt}$, which is sponsor-specific and only the sponsor observes:

$$\nu_{jt} = \ln(\beta_{jt}) + \varepsilon_{jt}, \quad \varepsilon_{jt} \sim \text{i.i.d. Type-I E.V.} \tag{1}$$

where $j$ indexes players and $t$ indexes designs. In this model, $\beta_{jt}$ represents the design's quality, which may not be known ex-ante and is revealed by the sponsor's feedback. The design's value to the sponsor, $\nu_{jt}$, is increasing and concave in its quality, and the design with the highest $\nu$ wins the contest.[7] The $\varepsilon_{jt}$ term is an i.i.d. random shock (luck), which may arise due to idiosyncracies in the sponsor's tastes at the time a winner is chosen. Player $j$'s probability of winning takes the following form:

$$Pr(\text{player } j \text{ wins}) = \frac{\beta_{j1} + \beta_{j2}}{\sum_{k \neq j}(\beta_{k1} + \beta_{k2}) + \beta_{j1} + \beta_{j2}} = \frac{\beta_{j1} + \beta_{j2}}{\mu_j + \beta_{j1} + \beta_{j2}} \tag{2}$$

where $\mu_j \equiv \sum_{k \neq j}(\beta_{k1} + \beta_{k2})$ is the competition that player $j$ faces in the contest. This function is concave in the player's own quality and decreasing in the quality of her competition.

Players develop and submit designs one at a time, in turns, and immediately receive public feedback that reveals $\beta_{jt}$. It is assumed that property protections are in place to prevent idea theft by competitors. Every player's first design in the contest is thus novel to that contest, and at their subsequent turn, players have three options: they can exploit (tweak, or adapt) the existing design, explore (experiment with) a radically different design, or abandon the contest altogether. I elaborate on each option:

1. **Exploitation** costs $c > 0$ and yields a design of the same quality as the one being exploited. A player who exploits will tweak her first design, which has quality $\beta_{j1}$, resulting in a second-round design with $\beta_{j2} = \beta_{j1}$ and increasing her probability of winning accordingly.

2. **Exploration** costs $d \geq c$ and can yield a high- or low-quality design, each with positive probability. Define $\alpha \geq 1$ as the exogenous degree of experimentation under this option (conversely, $\frac{1}{\alpha} \in [0, 1]$ can be interpreted as the similarity of the exploratory design to the player's first design). With probability $q$, exploration will yield a high-quality design with $\beta_{j2}^H = \alpha\beta_{j1}$; with probability $(1 - q)$ it will yield a low-quality design with $\beta_{j2}^L = \frac{1}{\alpha}\beta_{j1}$.[8] I assume $q > \frac{1}{1+\alpha}$, which implies that a design's expected quality under exploration ($E[\beta_{j2}|\text{Explore}]$) is greater than that under exploitation ($\beta_{j1}$). As specified, exploitation is a special case of exploration, with $\alpha = 1$.

[7] The decision to model $\nu_{jt}$ as logarithmic in $\beta_{jt}$ is taken for analytical convenience but also supported by the intuition of decreasing returns to quality. $\nu_{jt}$ could also be linear in $\beta_{jt}$ and the core inverted-U result of this paper would obtain, since the feature of the model driving the results is the concavity of the success function.

[8] In this model, I assume $\alpha$ and $q$ are fixed. When $\alpha$ is endogenized and costless, the (risk-neutral) player's optimal $\alpha$ is infinite, since the experimental upside would then be unlimited and the downside bounded at zero. A natural extension to this model would be to relax experimentation costs $d(\cdot)$ and/or the probability of a successful experiment $q(\cdot)$ to vary with $\alpha$. Such a model is considerably more difficult to solve and beyond the scope of this paper.

While its quality may be uncertain, a novel design's content is not arbitrary: it is the result of a player's deliberate choices and discretion in her attempt to produce a winning idea. Players with a track record of success can thus leverage this experience into higher-quality original work (in expectation) – as they are observed to do in the data. This feature subtly differentiates the phenomenon modeled here from pure experimentation as previously studied in other settings.

3. **Abandonment** is costless: the player can always walk away. Doing so leaves the player's probability of winning unchanged, as her previous work remains in contention.

Under this specification, sponsor feedback has three effects: it informs each player about her first design's quality, helps her improve and set expectations over her second design, and reveals the level of competition she faces. Players use this information to decide (i) whether to continue participating and (ii) whether to do so by exploring a new design or re-using a previous one, which is a choice over which kind of effort to exert: creative or rote. The model thus characterizes incentives for creativity.

In the remainder of this section, I examine a player's incentives to explore, exploit, or abandon the competition. Section 1.1 studies the conditions required for the player to prefer exploration over the alternatives and shows that these conditions lead to an inverted-U relationship between competition and creativity (proofs in Appendix A). Section 1.2 contextualizes this result in the existing literature. To simplify the mathematics, I assume the level of competition $\mu_j$ is known to player $j$, though the results are general to other assumptions about players' beliefs over the competition they will face, including competitors' best responses. The model can also be extended to allow players to enter an arbitrary number of designs, and the results will hold as long as players do not choose exploration for its option value.

## 1.1 Incentives for Exploration

To simplify notation, let $F(\beta_2) = F(\beta_2|\beta_1, \mu)$ denote player $j$'s probability of winning with a second design of quality $\beta_2$, given $\beta_1$ and $\mu$ (omitting the $j$ subscript). The model permits four values of $\beta_2$: $\beta_2^H$, $\beta_2^L$, $\beta_1$, and 0. The first two values result from exploration, and the latter two from exploitation and abandonment, respectively. For player $j$ to develop a new design, she must prefer doing so over both tweaking an existing design (incentive compatibility) and abandoning (individual rationality):

$$\underbrace{\left[qF\left(\beta_2^H\right) + (1-q)\,F\left(\beta_2^L\right)\right] \cdot P - d}_{E[\pi|\text{explore}]} > \underbrace{F(\beta_1) \cdot P - c}_{E[\pi|\text{exploit}]} \tag{IC}$$

$$\underbrace{\left[qF\left(\beta_2^H\right) + (1-q)\,F\left(\beta_2^L\right)\right] \cdot P - d}_{E[\pi|\text{explore}]} > \underbrace{F(0) \cdot P}_{E[\pi|\text{abandon}]} \tag{IR}$$

These conditions can be rearranged to be written as follows:

$$qF\left(\beta_2^H\right) + (1-q)\,F\left(\beta_2^L\right) - F\left(\beta_1\right) > \frac{d-c}{P} \tag{IC}$$

$$qF\left(\beta_2^H\right) + (1-q)\,F\left(\beta_2^L\right) - F\left(0\right) > \frac{d}{P} \tag{IR}$$

In words, the probability gains from exploration over exploitation or no action must exceed the cost differential, normalized by the prize. These conditions are less likely to be met as the cost of exploration rises, but the consideration of cost in players' decision-making is mitigated in tournaments with large prizes that dwarf the cost of a new design. As written, they will generate open intervals for $\mu \in \mathbb{R}^+$ in which players will degenerately prefer one of exploration, exploitation, or abandonment. If costs were stochastic – taking a distribution, as is likely the case in practice – the conditions would similarly generate intervals in which one action is more likely than (rather than preferred to) the others.

### 1.1.1 Exploration versus Exploitation (IC)

At what values of $\mu$ are the payoffs to exploration greatest relative to exploitation? I answer this question with the following lemma that characterizes the shape of these payoffs as a function of $\mu$, and a proposition establishing the existence of a unique value that maximizes this function.

**Lemma 1.** Payoffs to exploration over exploitation. *When $q \in (\frac{1}{1+\alpha}, \frac{1}{2})$, the gains to exploration over exploitation are decreasing and convex in $\mu$ for small $\mu$, increasing and concave for intermediate $\mu$, and decreasing and convex for large $\mu$. When $q \in \left(\frac{1}{2}, \frac{3\alpha+1}{4\alpha+1}\right)$, they are increasing and convex for small $\mu$ and decreasing and convex for large $\mu$. When $q > \frac{3\alpha+1}{4\alpha+1}$, they are increasing and concave for small $\mu$ and decreasing and convex for large $\mu$. When $q < \frac{1}{1+\alpha}$, they are decreasing and convex for small $\mu$ and increasing and concave for large $\mu$. In every case, the gains are zero when $\mu = 0$; when $q > \frac{1}{1+\alpha}$ ($q < \frac{1}{1+\alpha}$), they approach zero from above (below) as $\mu \longrightarrow \infty$, holding $\beta_1$ fixed.*

**Proposition 1.** *When $q > \frac{1}{1+\alpha}$, there exists a unique level of competition $\mu_2^*$ at which the gains to exploration, relative to exploitation, are maximized.*

A player's incentive to explore over exploit depends on her relative position in the contest: provided $q > \frac{1}{1+\alpha}$, a player will prefer exploration when she lags sufficiently far behind her competition, and exploitation when she is sufficiently far ahead. This result leads to the first empirical prediction, which I examine in Section 3.1: conditional on participating, positive feedback will increase players' tendency to exploit their previous idea, but this effect will be offset by high-quality competition.

### 1.1.2 Exploration versus Abandonment (IR)

I now ask the question: at what values of $\mu$ are the payoffs to exploration greatest relative to abandonment? I answer this question with a similar lemma and proposition.

**Lemma 2.** Payoffs to exploration over abandonment. *The gains to exploration over abandonment are increasing and concave in $\mu$ when $\mu$ is small and decreasing and convex when $\mu$ is large. The gains are zero when $\mu = 0$ and approach zero from above as $\mu \longrightarrow \infty$, holding $\beta_1$ fixed.*

**Proposition 2.** *For all values of $q$, there exists a unique level of competition $\mu_1^*$ at which the gains to exploration, relative to abandonment, are maximized.*

A player is thus likely to abandon when there is either very little or very intense competition. This result constitutes the second empirical prediction of the model, which I examine in Section 3.2.

#### 1.1.3 Tying it together: Exploration vs. the next-best alternative

**Proposition 3.** *At very low and very high $\mu$, the IR constraint binds: the next-best alternative to exploration is abandonment. At intermediate $\mu$, the IC constraint binds: the next-best alternative is exploitation.*

As $\mu$ increases from zero to infinity, the player's preferred action will evolve from abandonment, to exploitation, to exploration, to abandonment. Figure 1 plots the absolute payoffs to each as the degree of competition increases for an example parametrization. Note that the region in which players stop investing due to a lack of competition is very narrow, and effectively occurs only under pure monopoly.

[Figure 1 about here]

Putting the first three propositions together, the implication is an inverted-U shaped effect of competition on creativity. Provided that exploration is on average higher-quality than exploitation, incentives for exploration are maximized at a unique, intermediate level of competition, and this optimum will be attainable as long as the cost of exploration is not so large as to make it completely infeasible for the player. This inverted-U pattern is plotted in Figure 2 for the same parametrization as in Figure 1.

**Proposition 4.** *When $q > \frac{1}{1+\alpha}$, there exists a unique level of competition $\mu^* \in [\mu_1^*, \mu_2^*]$ at which the gains to exploration are maximized relative to the player's* next-best *alternative.*

[Figure 2 about here]

The inverted-U result can be understood through two steps, as follows: (i) at lower levels of competition, incentive compatibility binds, and increasing competition makes experimentation more attractive relative to incremental tweaks; (ii) at high levels of competition, the participation constraint binds, and increasing competition drives the returns to effort of either kind to zero. Incentives for exploration will generally peak at the point where the participation constraint becomes binding.

At the heart of this model is the player's choice between a gamble and a safe outcome. The concavity of the success function implies that players may prefer the certain outcome to the gamble, forgoing expected quality improvements, even though they are risk-neutral. The inverted-U result is therefore robust to risk-aversion, which only increases the concavity of the payoffs. Note that while the inverted-U result for players' incentives for creativity is not sensitive to the precise specification of the model, the sponsor's desire for creativity is, as it depends on both the specification and the parametrization.

While these results speak most directly to the incentives of the player who moves last, Appendix A shows that they also apply to inframarginal players, either in partial equilibrium or internalizing competitors' best responses: a player with an inordinate lead or deficit has no reason to continue, one with a solid lead can compel her competitors to quit by exploiting, and one in a neck-and-neck state or slightly behind will be most inclined to explore a new idea to have a fighting chance at winning.

### 1.2 Remarks and Relation to Previous Literature

The inverted-U effect of competition on creative effort is intuitive. With minimal or extreme competition, the returns are too low to justify continued participation – a pattern which is consistent with existing theoretical and empirical results from the tournament literature, which has argued that asymmetries will reduce effort from both leaders and laggards (e.g. Baik 1994, Brown 2011). The contribution of the model is to consider participation jointly with the bandit dilemma, which introduces a new layer to the problem. At intermediate levels of competition, continued participation is justified, but developing new ideas may not be: with only limited competition, the player is sufficiently well-served by exploiting her previous work. Only at greater levels of competition will the player have an incentive to experiment.[9]

The model introduces a new explanation of an inverted-U pattern to the literature on competition and innovation, distinct from yet complementary to Aghion et al. (2005), who study the effects of product market competition on step-by-step innovation. The principal contribution of this section is to show that a similar result obtains (i) in a winner-take-all setting, which is a fundamentally different competitive environment; (ii) when innovation is continuous rather than step-by-step; and (iii) when the intensity of competition is determined by the technology rather than the market structure.

## 2 Graphic Design Contests

I collected a randomly-drawn sample of 122 logo design contests from a widely-used online platform to study how creative effort responds to competition.[10] The platform from which the data were collected hosts hundreds of contests each week in several categories of commercial graphic design, including logos, business cards, t-shirts, product packaging, book/magazine covers, website/app mockups, and others. Logo design is

[9] It is tempting to also draw comparisons against models of patent races, in which firms compete to be the first to arrive at a successful innovation, with the threshold for quality fixed and time of arrival unknown. In innovation contests such as the one modeled here, firms compete to create the highest-quality innovation prior to a deadline. Although Baye and Hoppe (2003) establish an isomorphism between the two, it requires that players are making i.i.d. draws with each experiment. A player's probability of winning in either model is then determined by the number of draws they make – their "effort." This assumption quite clearly does not carry over to the present setting, where designs are drawn from distributions varying across players and over time. Some of the intuition from patent race models nevertheless applies, such as predictions that firms that are hopelessly behind will abandon the competition (Fudenberg et al. 1983).

[10] The sample consists of all logo design contests with public bidding that began the week of Sept. 3-9, 2013 and every three weeks thereafter through the week of Nov. 5-11, 2013, excluding those with multiple prizes or mid-contest rule changes such as prize increases or deadline extensions. Appendix B describes the sampling procedures in greater detail.

the modal design category on this platform and is thus a natural choice for analysis. A firm's choice of logo is also nontrivial, since it is the defining feature of its brand, which can be one of the firm's most valuable assets and is how consumers will recognize and remember the firm for years to come.

In these contests, a firm (the sponsor; typically a small business or non-profit organization) solicits custom designs from a community of freelance designers (players) in exchange for a fixed prize awarded to its favorite entry. The sponsor publishes a design brief describing its business, its customers, and what it likes and seeks to communicate with its logo; specifies the prize structure; sets a deadline for submissions; and opens the contest to competition. While the contest is active, players can enter (and withdraw) as many designs as they want, at any time they want, and sponsors can provide players with private, real-time feedback on their submissions in the form of 1- to 5-star ratings and written commentary. Players see a gallery of competing designs and the distribution of ratings on these designs, but not the ratings on specific competing designs. Copyright is enforced.[11] At the end of the contest, the sponsor picks the winning design and receives the design files and full rights to their use. The platform then transfers payment to the winner.

For each contest in the sample, I observe the design brief, which includes a project title and description, the sponsor's industry, and any specific elements that must be included in the logo; the contest's start and end dates; the prize amount; and whether the prize is committed.[12] While multiple prizes are possible, the sample is restricted to contests with a single, winner-take-all prize. I also observe every submitted design, the identity of the designer, his or her history on the platform, the time at which the design was entered, the rating it received (if any), the time at which the rating was given, and whether it won the contest. I also observe when players withdraw designs from the competition, but I assume withdrawn entries remain in contention, as sponsors can request that any withdrawn design be reinstated. Since I do not observe written feedback, I assume the content of written commentary is fully summarized by the rating.[13]

The player identifiers allow me to track players' activity over the course of each contest. I use the precise timing information to reconstruct the state of the contest at the time each design is submitted. For every design, I calculate the number of preceding designs in the contest of each rating. I do so both in terms of the prior feedback available (observed) at the time of submission as well as the feedback eventually provided. To account for the lags required to produce a design, I define preceding designs to be those entered at least one hour prior to a given design, and I similarly require that feedback be provided at least one hour prior to the given design's submission to be considered observed at the time it is made.

[11] Though players can see competing designs, the site requires that all designs be original and actively enforces copyright protections. Players have numerous opportunities to report violations if they believe a design has been copied or misused in any way. Violators are permanently banned from the site. The site also prohibits the use of stock art and has a strict policy on the submission of overused design concepts. These mechanisms are effective at limiting abuses.

[12] The sponsor may optionally retain the option of not awarding the prize to any entries if none are to its liking.

[13] One of the threats to identification throughout the empirical section is that the effect of ratings may be confounded by unobserved, written feedback: what seems to be a response to a rating could be a reaction to explicit direction provided by the sponsor that I do not observe. This assumption is evaluated in more depth in Appendix D, which shows that instructive comments are (i) rare, and (ii) not disproportionately given to 5-star designs relative to 3- or 4-star designs, such that they cannot be responsible for the differences in behavior that will be found in later tables.

The dataset also includes the designs themselves. I invoke image comparison algorithms commonly used in content-based image retrieval software (similar to Google Image's Search by Image feature) to quantify the originality of each design entered into a contest relative to preceding designs by the same and other players. I use two mathematically distinct procedures to compute similarity scores for image pairs, one of which is a preferred measure (the "perceptual hash" score) and the other of which is reserved for robustness checks (the "difference hash" score). Appendix B explains exactly how they work. Each one takes a pair of digital images as inputs, summarizes them in terms of a specific, structural feature, and returns a similarity index in [0,1], with a value of one indicating a perfect match and a zero indicating total dissimilarity. This index effectively measures the absolute correlation of two images' structural content.

To make this discussion concrete, the inset below demonstrates an example application. The figure shows three designs, entered in the order shown, by the same player in a logo design competition that is similar to those in the sample, although not necessarily from the same platform.[14] The first two logos have some features in common (they both use a circular frame and are presented against a similar backdrop), but they also have some stark differences. The perceptual hash algorithm gives them a similarity score of 0.31, and the difference hash algorithm scores them 0.51. The latter two logos are more alike, and though differences remain, they are now more subtle and mostly limited to the choice of font. The perceptual hash algorithm gives these logos a similarity score of 0.71, while the difference hash scores them 0.89. The appendix provides additional examples on more recognizable brands (Volkswagen and Microsoft Windows).

Illustration of image comparison algorithms

(1) (2) (3)

Notes: Figure shows three logos entered in order by a single player in a single contest. The perceptual hash algorithm calculates a similarity score of 0.313 for logos (1) and (2) and a score of 0.711 for (2) and (3). The difference hash algorithm calculates similarity scores of 0.508 for (1) and (2) and 0.891 for (2) and (3).

For each design in a contest, I compute its maximal similarity to previous designs in the same contest by the same player. Subtracting this value from one yields an index of originality between 0 and 1. This index is my principal measure of originality, and it is an empirical counterpart to the parameter $1/\alpha$ in the model. I also make use of related measures: for some specifications, I compare each design against only the best previously-rated designs by the same player or against the best previously-rated designs by competing players. Since players tend to re-use only their highest-rated work, the maximal similarity of a given design

[14] To keep the platform from which the sample was collected anonymous, I omit identifying information.

to any of that player's previous designs and maximal similarity to her highest-rated previous designs are highly correlated in practice (0.88 for the preferred algorithm, 0.87 for the alternative).

Creativity can manifest in other ways. For example, players sometimes create and enter several designs at once, and when doing so they can make each one similar to or distinct from the others. To capture this phenomenon, I define "batches" of proximate designs entered into the same contest by a single player and compute the maximum intra-batch similarity as a measure of creativity in batch work. Two designs are proximate if they are entered within 15 minutes of each other, and a batch is a set of designs in which every design in the set is proximate to another in the same set. Intra-batch similarity is an alternative – and arguably better – measure of creative experimentation, reflecting players' tendency to try minor variants of the same concept or multiple concepts over a short period of time.

These measures are not without drawbacks or immune to debate. One drawback is that these algorithms require substantial dimensionality reduction and thus provide only a coarse comparison between designs. Concerns on this front are mitigated by the fact that the empirical results throughout the paper are similar in sign, significance, and magnitude under two distinct algorithms. One might also question how well these algorithms emulate human perception. The example here and in the appendix assuage this concern; more generally, I have found these algorithms to be especially good at detecting designs that are plainly tweaks to earlier work (by my perception) versus those that are not, which is the margin that matters most for this paper. Appendix B discusses these and other related issues in greater detail.

## 2.1 Characteristics of the Sample

The average contest in the data lasts eight days, offers a $250 prize, and attracts 96 designs from 33 players (Table 1). On average, 64 percent of designs are rated; less than three receive the top rating.

[Table 1 about here]

Among rated designs, and the median and modal rating is three stars (Table 2). Though fewer than four percent of rated designs receive a 5-star rating, over 40 percent of all winning designs are rated five stars, suggesting that these ratings convey substantial information about a design's quality and odds of success.[15] The website also provides formal guidance on the meaning of each star rating, which generates consistency in their interpretation and use across different sponsors and contests.

[Table 2 about here]

Table 3 characterizes the similarity measures used in the empirical analysis. For each design in the sample, I measure its maximal similarity to previous designs by the same player, previously-rated designs by the

[15] Another 33 percent of winning designs are rated 4 stars, and 24 percent are unrated.

same player, and previously-rated designs by the player's competitors (all in the same contest). For every design batch, I calculate the maximal similarity of any two designs in that batch. Note that the analysis of intra-batch similarity is restricted to batches that are not missing any image files.

[Table 3 about here]

The designs themselves are available for 96 percent of submissions in the sample. The table shows that new entries are on average more similar to that player's own designs than her competitors' designs, and that designs in the same batch tend to be more similar to each other than to previous designs by even the same player. But these averages mask more important patterns at the extremes. At the upper decile, designs can be very similar to previous work by the same player ($\approx$0.75 under the perceptual hash algorithm) or to other designs in the same batch (0.91), but even the designs most similar to competing work are not all that similar (0.27). At the lower end, designs can be original by all of these measures.

Appendix C explores other features of these contests which may be of interest, including the relationship between contest characteristics, such as prize value or feedback intensity, and contest-level outcomes. The appendix draws on data from Gross (2015), which studies a larger but distinct sample from the same platform that can provide more precise estimates of these gradients.

**Are ratings meaningful? Estimating the empirical success function**

Using these larger-sample data, a player's probability of winning can be directly and precisely estimated as a function of her ratings.[16] Recall from Section 1 that a design's latent value is a function of its rating and an i.i.d. extreme value error. In the data, there are five possible ratings. This latent value can be flexibly specified with fixed effects for each rating, and the success function then be estimated as a conditional logit, using the win-lose outcomes of every design in a large sample of contests. To formalize, let $R_{ijk}$ denote the rating on design $i$ by player $j$ in contest $k$, and (in a slight abuse of notation) let $R_{ijk} = \emptyset$ when design $ijk$ is unrated. The value of each design, $\nu_{ijk}$, can be expressed as follows:

$$\nu_{ijk} = \gamma_{\emptyset}\mathbb{1}(R_{ijk} = \emptyset) + \gamma_1\mathbb{1}(R_{ijk} = 1) + \ldots + \gamma_5\mathbb{1}(R_{ijk} = 5) + \varepsilon_{ijk} \equiv \psi_{ijk} + \varepsilon_{ijk} \tag{3}$$

The details of the estimation are provided in Appendix C, but here we can summarize the patterns that emerge from the exercise. The fixed effects are precisely estimated, and the estimated value of a design is monotonically increasing in its rating. Only a 5-star design is on average preferred to the outside option. To produce the same increase in a player's win probability generated by a five-star design, a player would need 12 four-star designs, 137 three-star designs, or nearly 2,000 one-star designs, such that competition at the

[16] Estimating the success function requires a larger sample of winners, and thus contests, than are in the primary sample of this paper. As Appendix C shows, the sample in Gross (2015) contains >4,000 contests, is empirically comparable, and includes all of the same variables except for the images themselves – sufficient for the exercise.

top effectively only comes from other five-star designs. As a measure of fit, the model correctly "predicts" the true winner relatively well, with the odds-on favorite winning almost half of all contests in the sample. These results demonstrate that this simple model fits the data quite well and in an intuitive way, suggesting that ratings provide considerable information about a player's probability of winning. The strong fit of the model also justifies the assumption that players can accurately assess these odds: though players do not observe the ratings on specific competing designs, they are provided with the distribution of ratings on their competitors' designs, which makes it possible for players to invoke a simple heuristic model such as the one estimated here in their decision-making.

## 2.2 Empirical Methods and Identification

I exploit variation in the level and timing of the sponsor's ratings to estimate the effects of competition on players' creative choices. With timestamps on all activity, I can determine exactly what a player knows at each point in time about the sponsor's preference for her work and the competition she faces, and identify the effects of ratings observed at the time of design. Identification is achieved by harnessing variation in the *information* that players possess about their own and their competitors' performance.

Formally, the identifying assumption is that there are no omitted factors correlated with observed feedback that also affect choices. This assumption is supported by two pieces of evidence. First, the arrival of ratings is unpredictable, such that the set of ratings observed at any point in time is effectively random: sponsors are erratic, and it is difficult to know exactly when or how often a sponsor will log onto the site to rate new entries, much less any single design. More importantly, players' choices are uncorrelated with ratings that were unobserved at the time, including forthcoming ratings and ratings on specific competing designs. The thought experiment is to compare the actions of a player with a 5-star design under her belt before learning the rating versus after, or with latent 5-star competition before finding out versus after – noting that empirically, undisclosed information is as good as no information.[17]

To establish that feedback provision is unpredictable, I explore the relationship between feedback lags and ratings. In concept, sponsors may be quicker to rate the designs they like the most, to keep these players engaged and improving their work, in which case players might be able to infer their eventual ratings from the time elapsed without any feedback. Empirical assessment of this question (Appendix D) confirms that this is not the case: whether measured in hours or as a percent of the total contest duration, the lag between when a design is entered and rated is unrelated to its rating. The probability that a rated design was rated before versus after the contest ends is similarly uncorrelated with the rating granted.

[17] Though this setting may seem like a natural opportunity for a controlled experiment, the variation of interest is in the 5-star ratings, which are sufficiently rare that a controlled intervention would require either unrealistic manipulation or an infeasibly large sample. I therefore exploit naturally-occurring variation for this study.

Evidence that choices are uncorrelated with unobserved feedback is presented in Section 3. As a first check, I estimate the effects of observed ratings on originality both with and without controls for the player's forthcoming ratings and find the results unchanged. I also estimate the relationship between forthcoming ratings and originality, finding that it is indistinguishable from zero. In the appendix, I examine players' tendency to imitate highly-rated competing designs and find no such patterns, ostensibly because players simply do not know which competing designs are highly rated (and thus which ones to imitate). The results confirm that players respond only to information available at the time of design and that there are no omitted factors correlated with feedback that would confound its effects.

# 3 Competition and Creative Choices

Recall that the inverted-U effect predicted by the theory is the convolution of two results: (i) as competition intensifies, players are more likely to experiment versus tweak, conditional on participating, but (ii) when competition grows too intense, they stop participating altogether. These patterns generate the upward- and downward-sloping ends of the inverted-U, respectively. In this section I provide reduced-form evidence of each of these effects. Section 3.1 (as well as Appendices E through I) provides a battery of evidence that conditional on continued participation, competition induces high-performing players to enter more original work than they otherwise would. The primary estimating equation in this part of the paper is the following specification, with variants estimated throughout this section:

$$\begin{aligned} Similarity_{ijk} \quad = \quad & \beta_0 + \beta_5 \cdot \mathbb{1}(\bar{R}_{ijk} = 5) + \beta_{5c} \cdot \mathbb{1}(\bar{R}_{ijk} = 5) \cdot \mathbb{1}(\bar{R}_{-ijk} = 5) + \beta_{5p} \cdot \mathbb{1}(\bar{R}_{ijk} = 5) \cdot P_k \\ & + \sum_{r=2}^{4} \beta_r \cdot \mathbb{1}(\bar{R}_{ijk} = r) + \gamma \cdot \mathbb{1}(\bar{R}_{-ijk} = 5) + T_{ijk}\lambda + X_{ijk}\theta + \zeta_k + \varphi_j + \varepsilon_{ijk} \end{aligned}$$

where $Similarity_{ijk}$ is the maximal similarity of design $ijk$ to any previous designs by player $j$ in contest $k$; $\bar{R}_{ijk}$ is the highest rating player $j$ has received in contest $k$ prior to design $ijk$; $\bar{R}_{-ijk}$ is the highest rating player $j$'s competitors have received prior to design $ijk$; $P_k$ is the prize in contest $k$ (measured in \$100s); $T_{ijk}$ is fraction of the contest elapsed at the time design $ijk$ is entered; $X_{ijk}$ is a vector of design-level controls, consisting of the number of previous designs by the same player and competing players, and days remaining in the contest; and $\zeta_k$ and $\varphi_j$ are contest and player fixed effects, respectively.

It may be helpful to provide a roadmap to this part of the analysis in advance. In the first set of regressions, I estimate the specification above. In the second set, I replace the dependent variable with the similarity to that player's best, previously-rated designs, and then within-batch similarity. The third set of regressions examines the change in similarity to previously-rated designs, as a function of newly-received feedback. The fourth set of regressions tests the aforementioned identifying assumption that players are not acting on forthcoming information. The appendix provides additional robustness checks.

Section 3.2 considers the counterpart analysis examining the effects of competition on players' tendency to continue investing in or abandon the contest. The evidence substantiates the model's second prediction: that increasing competition can drive players to quit. The specifications in this section are similar to those of the regressions testing originality. I estimate variants of the following model:

$$\begin{aligned} Abandon_{ijk} \quad = \quad & \beta_0 + \sum_{r=1}^{5} \beta_r \cdot \mathbb{1}(\bar{R}_{ijk} = r) + \sum_{r=1}^{5} \delta_r \cdot \mathbb{1}(\bar{R}_{ijk} = r) \cdot N_{-ijk} \\ & + N_{-ijk}\delta + T_{ijk}\lambda + X_{ijk}\theta + \zeta_k + \varphi_j + \varepsilon_{ijk} \end{aligned}$$

where $Abandon_{ijk}$ indicates that player $j$ entered no additional designs in contest $k$ after design $ijk$; $N_{-ijk}$ is the number of five-star designs by player $j$'s competitors in contest $k$ at the time of design $ijk$; and $\bar{R}_{ijk}$, $\bar{R}_{-ijk}$, $T_{ijk}$, $X_{ijk}$, $\zeta_k$, and $\varphi_j$ retain their previous definitions. The precise moment at which each player makes a decision to stop investing is impossible to measure, and I thus use inactivity as a proxy. In general, this measure does not distinguish between a "wait and see" approach that ends with abandonment versus abandonment immediately following design $ijk$. Since the end result is the same, the distinction is immaterial for the purposes of this paper. Standard errors throughout the following subsections are clustered by player to account for any within-player correlation in the error term.

## 3.1 Competition and Originality

### 3.1.1 Similarity of new designs to a player's previous designs

I begin by studying players' tendency to enter original versus derivative designs. Table 4 provides estimates from regressions of the maximal similarity of each design to previous designs by the same player on indicators for the highest rating that player had previously received. All specifications interact the indicator for the top rating with (i) the prize (in \$100s) and (ii) a variable indicating the presence of top-rated competition, and control for the fraction of the contest elapsed and contest and player fixed effects. Even-numbered columns add the remaining controls. Columns (3) and (4) additionally account for future feedback on the player's earlier work; if players have contest-specific ability or other information unobserved by the researcher (e.g., sponsors' written comments), it will be accounted for by these regressions.

[Table 4 about here]

The results are consistent across all specifications in the table. Players with the top rating enter designs that are 0.3 points, or roughly one full standard deviation, more similar to previous work than players who have only low feedback (or no feedback). Roughly one third of this effect is shaved off by the presence of high-rated competition. With a highest observed rating of four stars, new designs are on average around 0.1

points more similar to previous work. This effect further attenuates as the best observed rating declines, and it is indistinguishable from zero at a best observed rating of two stars.

In practice, players tend to tweak only their highest-rated designs. Table 5, columns (1) and (2) estimate a variant on the first two columns of Table 4, regressing each design's maximal similarity to the *highest-rated* preceding designs by the same player on the same set of explanatory variables. Columns (3) and (4) invoke the sample of design batches and the alternative measure of creativity: the maximal similarity of any two designs in each batch. Columns (5) and (6) repeat this latter exercise, weighting observations of batches by their size. All specifications control for contest and player fixed effects, and the table shows variants of the regressions with versus without design- and batch-level covariates.

[Table 5 about here]

The results for the design-level regressions (Columns 1 and 2) are similar to but slightly stronger than those of the previous table. Players with the top rating enter designs that are 0.35 points, or about 1.3 standard deviations, more similar to their highest-rated work in that contest, but this effect is reduced by more than half when there is top-rated competition. Players' tendency to make tweaks on their best designs is again monotonically decreasing in their highest prior rating.

Columns (3) to (6) demonstrate that competition has similar effects on originality within batches of designs. When entering multiple designs at one time, the maximal similarity of any two designs in the batch declines 0.3 points, or approximately one standard deviation, for players with a top rating who face top-rated competition, relative to those who do not. Top players facing competition are thus more likely to experiment not only across batches but also within them. The consistency of the results demonstrates that they are not sensitive to inclusion of controls or weighting batches by their size.

The regressions in Tables 4 and 5 use contest and player fixed effects to control for factors that are constant within individual contests or players, but they do not control for factors that are constant within a given contest for a given player, as doing so leaves too little variation to identify the focal effects. Such factors may nevertheless be confounding. For example, if players can sense their match to a particular contest, and self-select or change their behavior throughout that contest, the estimated effects may be confounded by both types of selection. The estimates in the previous tables additionally mask potential heterogeneity in players' reactions to feedback and competition over the course of a contest.

Table 6 addresses these issues with a model in first differences. The dependent variable is the change in designs' similarity to the player's best previously-rated work. This variable can take values in [-1,1], where a value of 0 indicates that the given design is as similar to the player's best preceding design as was the last one she entered; a value of 1 indicates that the player transitioned fully from innovating to recycling; and a value of -1, the converse. The independent variables are changes in indicators for the highest rating the

player has received, with the usual interactions of the top rating with the prize and the presence of top-rated competition. I estimate this model with assorted configurations of contest fixed effects, player fixed effects, and controls to account for other reasons why players' inclination to experiment may change over time, though the results are not statistically different across these specifications.

[Table 6 about here]

The results provide the strongest evidence thus far on the effects of feedback and competition on creative choices. When a player receives her first 5-star rating, her next design will be a near replica. The degree of similarity increases by nearly 0.9 points, or *three* standard deviations. Top-rated competition shaves nearly half of this effect. Given their magnitudes, these effects will be plainly visible to the naked eye (see the earlier inset for an example of what they look like in practice). The effects of a new, best rating of 4-, 3-, and 2-stars on originality attenuate monotonically, similar to earlier results.[18]

The appendix provides robustness checks and supplementary analysis. To confirm that these patterns are not an artifact of the perceptual hash algorithm, Appendix E re-estimates the regressions in the preceding tables using the difference hash algorithm to calculate similarity scores. The results are statistically and quantitatively similar. In Appendix F, I split out the effects of competition by the number of top-rated competing designs, finding no statistical differences between the effects of one versus more than one: all effects of competition on originality are achieved by one high-quality competitor.

This latter result is especially important for ruling out an information-based story. The fact that other designs received a 5-star rating might indicate that the sponsor has diverse preferences and that experimentation has a higher likelihood of success than the player might otherwise believe. If this were the case, then originality should continue to increase as 5-star competitors are revealed. That this is not the case suggests that the effect is in fact the result of variation in incentives from competition.

The model suggests that similar dynamics should arise for players with 4-star ratings facing 4-star competition when no higher ratings are granted, since only relative performance matters – though this result may only arise towards the end of a contest, when the absence of higher ratings approaches finality. Appendix G tests this prediction, finding similar patterns for 4-on-4 competition that strengthen over the course of a contest. In unreported regressions, I also look for effects of 5-star competition on originality for players with only 4-star designs, and find attenuated effects that are negative but not significantly different from zero. I also explore the effect of prize commitment on originality, since the sponsor's outside option of not awarding the prize is itself a competing alternative – one which according to the conditional logit estimates

[18] Interestingly, these regressions also find that new recipients of the top rating can also be induced to try new designs with larger prizes. The theory suggests a natural explanation: large prizes moderate the role of potentially higher costs in players' decision-making. If original designs are more costly (take more time or effort) than tweaks, they may only be worth doing when the prize is large. This is particularly the case for players with highly-rated work in the contest, given how the shape of and movement along a player's success function depends on the quality of her designs.

in Table C.2 is on average preferred to all but the highest-rated designs. The effect of prize commitment is not estimated to be different from zero. I similarly test for effects of the presence of four-star competition on originality for players with five-star designs, finding none. These results reinforce the perception that competition effectively comes from designs with a contest's highest rating.

For even more evidence that behavior is driven by information, Appendix H asks whether players imitate high-rated competitors, which in practice they should not be able to identify, and finds no evidence that they do so in any systematic way. As a final check, Appendix I looks for comparable behavior by entrants. The concept of originality in this case is not well-defined, as there is no obvious precedent against which to compare, but as a first-approximation I restrict to contest-players who enter with a batch of submissions and evaluate similarity within initial batches. I find that entrants' intra-batch similarity decreases monotonically in the quality of competition at entry, with effects of large magnitude, but the sample is sufficiently reduced that the standard errors do not allow us to rule out no effect.

#### 3.1.2 Placebo test: Similarity of new designs to a player's not-yet-rated designs

The identifying assumptions require that players are not acting on information that correlates with feedback but is unobserved in the data. As a simple validation exercise, the regressions in Table 7 perform a placebo test of whether originality is related to impending feedback. If an omitted determinant of creative choices is correlated with the feedback, then it would appear as if originality responds to forthcoming ratings, but if the identifying assumptions hold, we should only find zeros.

[Table 7 about here]

The specification in Column (1) regresses a design's maximal similarity to the player's best designs that have not yet been *but will eventually be* rated on indicators for the ratings they later receive. I find no evidence that designs' originality is related to forthcoming ratings. Because a given design's similarity to an earlier, unrated design can be incidental if both are tweaks on a third design, Column (2) adds controls for similarity to the best already-rated design. Column (3) allows these controls to vary with the rating received. As a final check, I isolate the similarity to the unrated design that cannot be explained by similarity to the third design in the form of a residual, and in Column (4) I regress these residuals on the same independent variables. In all cases, I find no evidence that players systematically tweak designs with higher forthcoming ratings. Feedback only relates to choices when observed in advance.

### 3.2 Competition and Abandonment

The natural next step is to examine the prediction that governs the right-hand side of the inverted-U: that intense competition drives players to quit, as the returns to effort decline to zero.

In Table 8 I estimate the probability that a given design is a player's final submission on the feedback and competition that was observed at the time. As previously discussed, this measure of abandonment could reflect either a simultaneous choice to abandon the project or a "wait and see" strategy that yields no further action – although according to one designer, it is often the case that players will enter their final design knowing it is their final design and never look back. The specifications in the table regress this measure of abandonment on indicators for the highest rating the player previously received, interactions with the number of 5-star competing designs, the latter as its own regressor, the fraction of the contest elapsed (a crucial control here), and other controls from previous tables.

[Table 8 about here]

Columns (1) to (3) estimate linear specifications with contest, player, and contest and player fixed effects. Linear specifications are used in order to control for these fixed effects (especially player fixed effects), which may not be estimated consistently in practice and could thus render the remaining estimates inconsistent in a binary outcome model. Column (4) estimates a logit model with only the contest fixed effects. The linear model with two-way fixed effects (in Column 3) is the preferred specification.

In all specifications, I find that players with higher ratings are more likely to continue investing than those with lower ratings, but that high-rated competition drives them away. In the preferred linear probability model in Column (3), I find that players with a top-rated design are more likely to continue investing, but the effect is offset by the presence of only a few top-rated competitors.

# 4 When is Creativity Most Likely?

The reduced-form results establish that while competition can motivate high performers to experiment with new ideas, too much competition will drive them out of the market altogether. How much is "too much"? Given that the full effect of competition on originality is achieved by a single, high-quality competitor, and that players are increasingly likely to stop investing as competition intensifies, it would be natural to conclude that incentives for continued creativity peak in the presence of one top-rated competitor – just enough to ensure that competition exists without further eroding returns to effort.

Because agents in this setting face a multi-choice menu, the question cannot be answered with a linear model. To formalize an answer, I estimate a multinomial logit choice model in which with each submission, a player selects from the three basic behaviors I observe in the data: (i) tweak and enter more designs, (ii) experiment and enter more designs, and (iii) do either and subsequently abandon the contest. For this exercise, I classify each design as a tweak if its similarity to any earlier design by the same player is 0.7 or higher and original if its maximal similarity to earlier designs by that player is 0.3 or lower.[19]

[19] A player's intentions are more ambiguous at intermediate values, which I accordingly omit from the exercise.

Each action in this choice set is assumed to have latent utility $u^a_{ijk}$, where $i$ indexes submissions by player $j$ in contest $k$. I model this latent utility as a function of the player's own ratings, her competitors' ratings, the fraction of the contest transpired, additional controls, and a logit error term:

$$\begin{aligned} u^a_{ijk} \quad = \quad & \beta^a_0 + \sum_{r=1}^{5} \beta^a_r \cdot \mathbb{1}(\bar{R}_{ijk} = r) + \sum_{r=1}^{5} \gamma^a_r \cdot \mathbb{1}(\bar{R}_{-ijk} = r) \\ & + \delta^a_1 \cdot \mathbb{1}(\bar{R}_{ijk} = 5)\mathbb{1}(N_{-ijk} = 1) \\ & + \delta^a_2 \cdot \mathbb{1}(\bar{R}_{ijk} = 5)\mathbb{1}(N_{-ijk} = 2) \\ & + \delta^a_3 \cdot \mathbb{1}(\bar{R}_{ijk} = 5)\mathbb{1}(N_{-ijk} \geq 3) \\ & + T_{ijk}\lambda^a + X_{ijk}\theta^a + \varepsilon^a_{ijk}, \quad \varepsilon^a_{ijk} \sim \text{i.i.d. Type-I E.V.} \end{aligned}$$

The explanatory variables are defined as before: $\bar{R}_{ijk}$ is the highest rating player $j$ has received in contest $k$ prior to $ijk$, $\bar{R}_{-ijk}$ is the highest rating on competing designs, $N_{-ijk}$ is the number of top-rated competing designs, $T_{ijk}$ is the fraction of the contest already elapsed, and $X_{ijk}$ are controls.

I estimate the parameters by maximum likelihood using observed behavior. I then use the results to compute the probability that a player with a 5-star design takes each of the three actions near the end of a contest, and to evaluate how these probabilities vary as the number of top-rated competitors increases from zero to three or more (for the case of no 5-star competitors, I assume the highest rating on any competing design is 4 stars). These probabilities are shown in Figure 3. Panel A plots the probability that the player tweaks; Panel B, the probability that she experiments; and Panel C, the probability that she abandons the contest. Bars around each point provide the associated 95 percent confidence interval.

[Figure 3 about here]

The probability that a high-performer tweaks one of her previous submissions (Panel A) peaks at 52 percent when there are no 5-star competitors and is significantly lower with non-zero competition, with all differences significant at the one percent level. The probability that the player actively experiments (Panel B) peaks at 52 percent with one 5-star competitor and is significantly lower with zero, two, or three 5-star competitors (differences against zero and three significant at the one percent level; difference against two significant at the ten percent level). Panel C shows that the probability of abandonment increases monotonically in the level of competition, approaching 80 percent under heavy competition.

The inverted-U pattern can also be unearthed by other means. Taking advantage of the richness of the data, an alternative is to model the same discrete choice as a function of a player's contemporaneous probability of winning, which can be computed using the conditional logit estimates in Table C.2, and the standard controls. Figure 4 plots the results. The probability that a player tweaks her previous work (Panel A) is maximized at around 0.7 when she is a strong favorite to win, and declines monotonically to zero as her

odds decline. The probability that the player experiments (Panel B) follows a distinct and highly significant inverted-U pattern, peaking at approximately a one-half odds of winning. Finally, the probability that she abandons (Panel C) increases from zero to around 0.8 as her odds of winning decline.

[Figure 4 about here]

Observed behavior thus appears to conform to the predictions of economic theory: when competition is low, players are on the margin between exploring and exploiting, whereas when competition is high, they straddle the margin between exploring and abandoning. The results also concord with the reduced-form evidence in finding that high-rated players are most creative when facing precisely one highly-rated competitor. Panel B of each figure directly illustrates the empirical inverted-U pattern.

# 5 Evaluating the Returns to Creativity

Why do players respond to competition by exploring new directions? In conversations with creative professionals, including the panelists hired for the exercise below, many have asserted that competition means that they need to "be bold" or "bring the 'wow' factor," and that it induces them to take creative risks. Gambling on a more radical, untested idea is thus a calculated and intentional choice. The implicit assumption motivating this choice is that innovation is high-risk but high-return – the upside is what makes it worthwhile. This assumption is common not only in research, but also in the public discourse on innovation and entrepreneurship. Whether or not it is true is ultimately an empirical question.

A natural approach to answering this question in the setting of this paper is to look at the distribution of sponsors' ratings on original versus tweaked designs in the sample. To do so, I categorize designs as tweaks if they have similarity to any earlier designs by the same player of 0.7 or higher and record the rating of the design they are most similar to; I classify designs as original if their maximal similarity to earlier designs by that player is 0.3 or below and record the highest rating the player had previously received. I then compare the distribution of sponsors' ratings on this subsample, conditioning on the rating of the tweaked design (for tweaks) or the highest rating previously given to that player (for originals). The evidence suggests that original designs outperform tweaks to designs with low ratings, but *underperform* tweaks of top-rated designs, raising the question of why a player would deviate from her top-rated work.

The problem with this approach is that the observed outcomes are censored: it is impossible to observe the fruits of a player's creativity beyond the 5-star rating. With this top-code in place, original designs will necessarily appear to underperform tweaks of 5-star designs – the sponsor's rating can only go down. The data are thus inadequate for evaluating the benefits to exploration for players at the top. To circumvent the top-code, I hired a panel of five professional graphic designers to independently assess all 316 designs

in my sample that were rated five stars by contest sponsors, and I look to the panelists' ratings to evaluate whether exploration is in fact high-risk, high-return.

## Results from a Panel of Professional Designers

For this exercise, I hired five professional graphic designers at their regular rates to evaluate each design on an extended scale. These ratings were collected though a custom web-based application in which designs were presented in random order and panelists were limited to 100 ratings per day. With each design, the panelist was provided the project title and client industry (excerpted from the source data) and instructed to rate the "quality and appropriateness" of the given logo on a scale of 1 to 10.

Appendix J provides more detail on the survey procedure and shows the distribution of ratings from all five panelists. One panelist ("Rater 5") was particularly critical with his/her ratings and frequently ran up against the lower bound. The mass around the lower bound was apparent after the first day of the survey (i.e., after 100 ratings), and though I provide the results from this panelist in the appendix for the sake of disclosure, the decision was made at that time to exclude these ratings from subsequent analysis. The results are nevertheless robust to including ratings from this panelist above the lower bound.

To account for differences in the remaining panelists' austerity, I first normalize their ratings by demeaning, in essence removing rater fixed effects. For each design, I then compute summary statistics of the panelists' ratings (mean, median, maximum, and s.d.). As an alternative approach to aggregating panelists' ratings, I also calculate each design's score along the first principal component generated by a principal component analysis. Collectively, these summary statistics characterize the distribution of opinion on a given design. One way to think about them is as follows: if contest sponsors were randomly drawn from this population, then the realized rating on the design would be a random draw from this distribution.

I identify designs as being tweaks or originals using the definitions above and then compare the level and heterogeneity in panelists' ratings on designs of each type. Table 9 provides the results. Designs classified as tweaks are typically rated below-average, while those classified as original are typically above-average. These patterns manifest for the PCA composite, mean, and median panelist ratings; the difference in all three cases is on the order of around half of a standard deviation and is significant at the one percent level. The maximum rating that a design receives from any of the panelists is also greater for originals, with the difference significant at the one percent level. Yet so is the level of disagreement: the standard deviation across panelists' ratings on a given design is significantly greater for original designs than for tweaks. The evidence thus appears to support the popular contention that radical innovation is both higher mean and higher variance than incremental innovation, even at the top.[20]

[20] If anything, these differences may be understated. If a player enters the same design twice, the first would first be classified in this exercise as an original, and the second as a tweak, but they would receive the same rating from panelists. Excluding designs that are either tweaks of or tweaked by others in the sample does not affect the results.

[Table 9 about here]

# 6 Implications and Conclusion

This paper combines theory, data from a sample of commercial logo design competitions, and new tools for measuring the content of ideas to show that individuals' incentives for creativity are greatest at intermediate levels of competition. The results tie together literatures in bandit decision models and tournament competition, and they provide what is to my knowledge the most direct evidence yet available on how incentives influence the intensity and direction of individuals' creative production.

The results have direct implications for policies and programs meant to incentivize creativity, both in and outside of the workplace. The foremost result is that the sharp incentives of prize competition can motivate creative effort in a work environment, but that doing so requires striking a delicate balance in the intensity of competition. In designing contracts for creative workers, managers would be keen to offer incentives for high-quality work relative to that of peers or colleagues, in addition to the traditional strategy of establishing a work environment with intrinsic motivators such as intellectual freedom, flexibility, and challenge. Another advantage of the tournament-style incentive structure is that it incorporates tolerance for failure by allowing players to recover from unsuccessful experimentation, which has been shown to be an important feature of contracts for motivating innovation (e.g., Manso 2011, Ederer and Manso 2013).

In practice, the 'Goldilocks' level of competition preferred by a principal may be difficult to achieve, much less determine. Finding it would likely require experimentation with the mechanism itself on the part of the principal, such as by varying the prize, subsidizing or restricting entry, or eliminating non-preferred players midway through the contest. In this paper, one high-quality competitor was found to be sufficient to induce another high-quality player to explore new directions, and further increases in competition have the effect of driving players away. As a rule of thumb for other settings, a good approximation may be to assume that one competitor of similar ability is enough to elicit fresh thinking, but that having more than a few such competitors is likely more harmful than helpful for motivating creative workers.

The results also have bearing on design of public incentives for R&D, which is itself a creative endeavor, and the implementation of other policies (such as antitrust policy) undertaken with the intent of incentivizing innovation. Although this paper is fundamentally about individuals, the theoretical framework can be interpreted as firms competing in a winner-take-all market. This interpretation is not without some peril, as markets are inherently more complex and dynamic than the model allows. The results nevertheless shed light on the forces that define the relationship between competition and innovation, particularly in settings where post-innovation rents are much larger than the pre-existing rents.

Three concrete policy implications follow. The first is support for maintaining balanced competition in winner-take-all procurement mechanisms for innovation. While the focus of this paper is graphic design for marketing materials, it is plausible that similar forces would be at work in other settings, such as R&D. This result is particularly relevant today, as governments, private foundations, and firms are increasingly contracting for R&D with prizes and institutionalizing prize competition.[21] The U.S. federal government now operates a platform (Challenge.gov) where agencies can solicit creative solutions to technical and non-technical problems from the public; as of February 2014, the site listed hundreds of open competitions with prizes ranging from status only (non-monetary) to upwards of 15 million dollars.

The second implication is an argument for monopoly and perfect competition being equally harmful to innovation in winner-take-all markets. According to the U.S. Horizontal Merger Guidelines (2010), "competition often spurs firms to innovate," and projected post-merger changes in the level of innovation is one of the government's criteria for evaluating mergers. The results of this paper suggest that a transition from no competition to some competition increases incentives for radical innovation over more modest, incremental improvements to existing technologies, but that returns to innovation can decline to zero in crowded or overly competitive markets, leaving participants content with the status quo.

A final implication of the results in this paper is that contrary to the conventional wisdom that duplicated R&D efforts are wasteful, simultaneous duplication may be ex-ante efficient: the competition of a horserace may induce greater ingenuity, whose fruits might compensate for the deadweight loss of duplicated effort. From a social welfare perspective, institutional policies prohibiting joint support of dueling research programs would then be doing more harm than good. This corollary requires testing, but if true, it suggests not only a fresh look at existing research on the welfare impacts of R&D, but potentially important changes to both R&D policy and strategies for managing innovation in the firm.

[21] See Williams (2012) for examples and an in-depth review of the literature on innovation inducement prizes.

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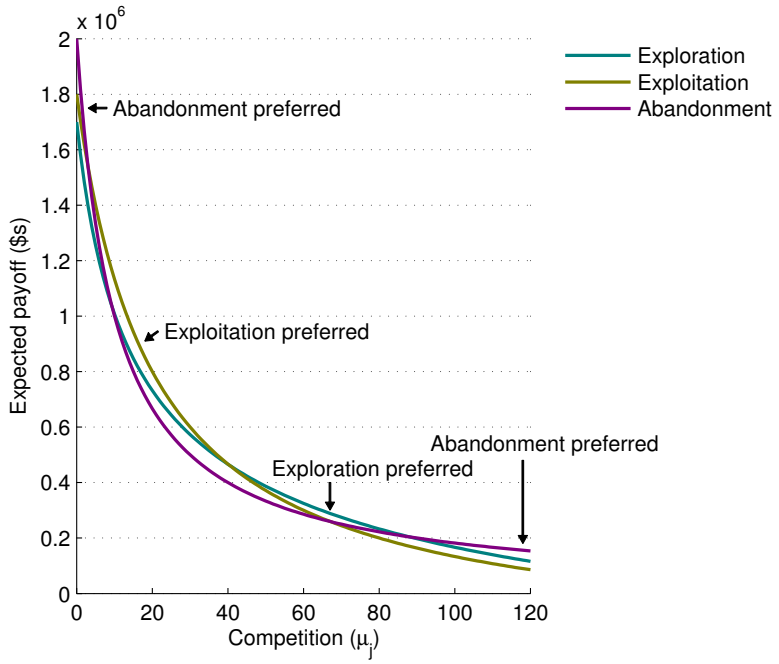
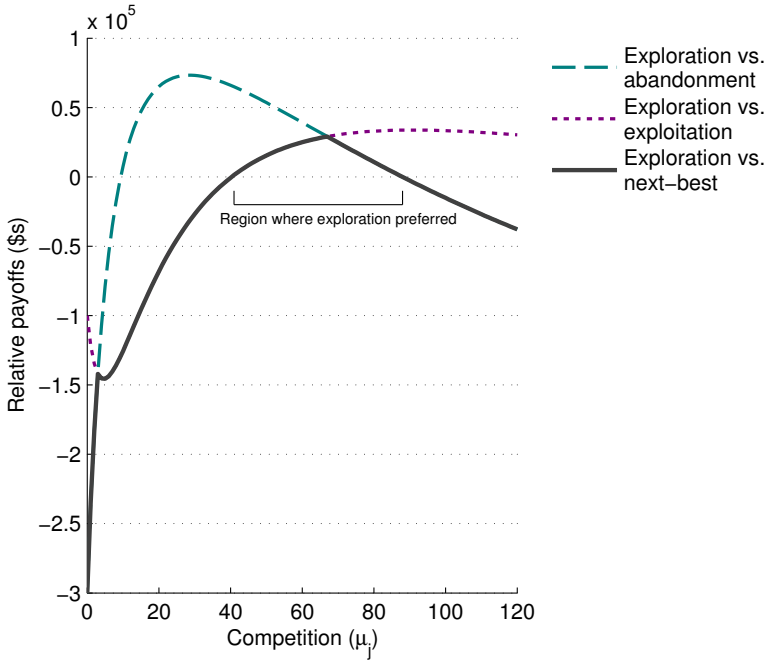
Figure 1: Payoffs to each of exploration, exploitation, and abandonment (example)



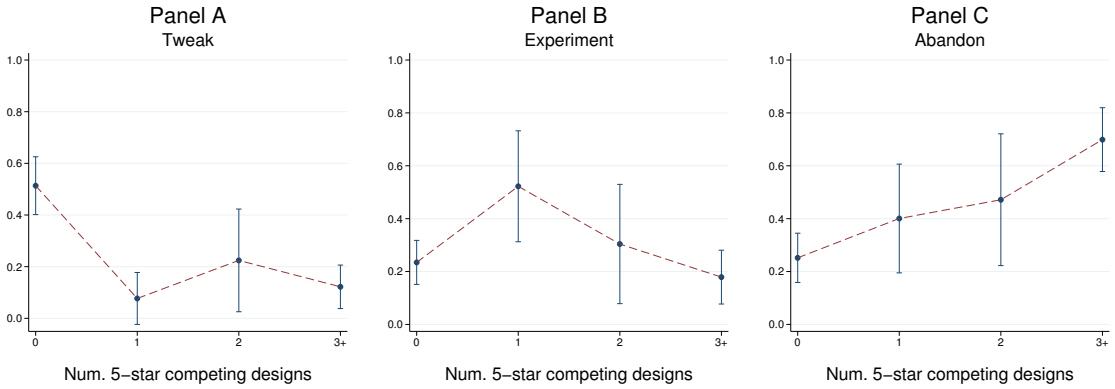

Notes: Figure plots the payoffs to each of exploration, exploitation, and abandonment as a function of $\mu_j$ for the following parametrization:
$\beta_{j1} = 10$, $q = 0.33$, $\alpha = 9$, $c = 200$k, $d = 300$k, $P = 2$m.

Figure 2: Payoff to exploration over next-best option (example)




Notes: Figure plots the incremental payoff to exploration over the next-best alternative at each value of $\mu_j$ for the same parametrization as the previous figure. The inverted-U effect shown here is the core theoretical result of the paper.

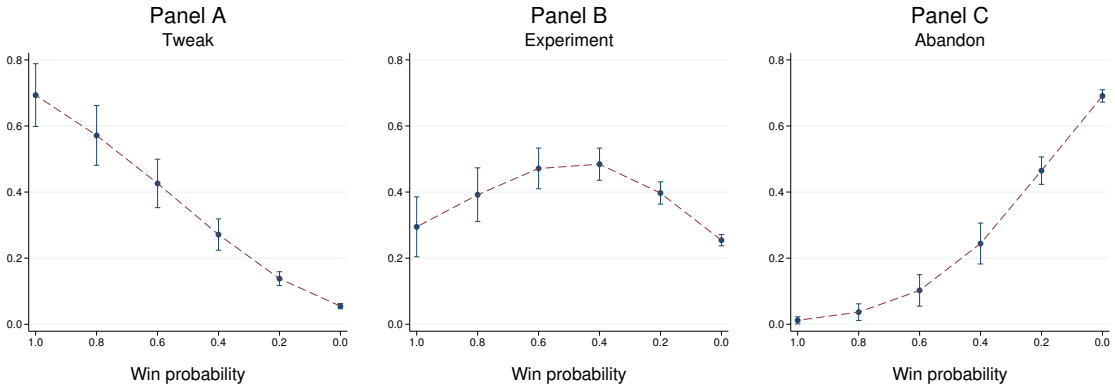
Figure 3: Probability of tweaking, experimenting, and abandoning as a function of 5-star competition




Notes: Figure plots the probability that a player who has at least one 5-star rating in a contest does one of the following on (and after) a given submission: tweaks and then enters more designs (Panel A), experiments and then enters more designs (Panel B), or abandons the contest (Panel C). The bars around each point provide the associated 95 percent confidence interval.

The figure establishes that experimentation follows an inverted-U pattern over competition in the data. This non-monotonicity appears to arise due to dynamics posited by the theory: when the probability of winning is high, players are on the margin between tweaks and experimentation (the incentive compatibility constraint, Panels A and B); when it is low, they are on the margin between experimentation and abandonment (participation constraint, Panels B and C).

Figure 4: Probability of tweaking, experimenting, and abandoning as a function of Pr(Win)




Notes: Figure plots the probability that a player does one of the following on (and after) a given submission, as a function of their contemporaneous win probability: tweaks and then enters more designs (Panel A), experiments and then enters more designs (Panel B), or abandons the contest (Panel C). These probabilities are estimated as described in the text, and the bars around each point provide the associated 95 percent confidence interval.

Table 1: Characteristics of contests in the sample

| Variable | N | Mean | SD | P25 | P50 | P75 |
|---|---|---|---|---|---|---|
| Contest length (days) | 122 | 8.52 | 3.20 | 7 | 7 | 11 |
| Prize value (US$) | 122 | 247.57 | 84.92 | 200 | 200 | 225 |
| No. of players | 122 | 33.20 | 24.46 | 19 | 26 | 39 |
| No. of designs | 122 | 96.38 | 80.46 | 52 | 74 | 107 |
| 5-star designs | 122 | 2.59 | 4.00 | 0 | 1 | 4 |
| 4-star designs | 122 | 12.28 | 12.13 | 3 | 9 | 18 |
| 3-star designs | 122 | 22.16 | 25.33 | 6 | 16 | 28 |
| 2-star designs | 122 | 17.61 | 25.82 | 3 | 10 | 22 |
| 1-star designs | 122 | 12.11 | 25.24 | 0 | 2 | 11 |
| Unrated designs | 122 | 29.62 | 31.43 | 7 | 19 | 40 |
| Number rated | 122 | 66.75 | 71.23 | 21 | 50 | 83 |
| Fraction rated | 122 | 0.64 | 0.30 | 0.4 | 0.7 | 0.9 |
| Prize committed | 122 | 0.56 | 0.50 | 0.0 | 1.0 | 1.0 |
| Prize awarded | 122 | 0.85 | 0.36 | 1.0 | 1.0 | 1.0 |

Notes: Table reports descriptive statistics for the contests. "Fraction rated" refers to the fraction of designs in each contest that gets rated. "Prize committed" indicates whether the contest prize is committed to be paid (vs. retractable). "Prize awarded" indicates whether the prize was awarded. The fraction of contests awarded awarded subsumes the fraction committed, since committed prizes are always awarded.

Table 2: Distribution of ratings (rated designs only)

| | 1-star | 2-star | 3-star | 4-star | 5-star | Total |
|---|---|---|---|---|---|---|
| **Count** | 1,478 | 2,149 | 2,703 | 1,498 | 316 | **8,144** |
| **Percent** | 18.15 | 26.39 | 33.19 | 18.39 | 3.88 | **100** |

Notes: Table tabulates rated designs by rating. 69.3 percent of designs in the sample are rated by sponsors on a 1-5 scale. The site provides guidance on the meaning of each rating, which introduces consistency in the interpretation of ratings across contests.

Table 3: Similarity to preceding designs by same player and competitors, and intra-batch similarity

| **Panel A. Using preferred algorithm: Perceptual Hash** | | | | | | |
|---|---|---|---|---|---|---|
| **Variable** | **N** | **Mean** | **SD** | **P10** | **P50** | **P90** |
| Max. similarity to any of own preceding designs | 5,075 | 0.32 | 0.27 | 0.05 | 0.22 | 0.77 |
| Max. similarity to best of own preceding designs | 3,871 | 0.28 | 0.27 | 0.03 | 0.17 | 0.72 |
| Max. similarity to best of oth. preceding designs | 9,709 | 0.14 | 0.1 | 0.04 | 0.13 | 0.27 |
| Maximum intra-batch similarity | 1,987 | 0.45 | 0.32 | 0.05 | 0.41 | 0.91 |
| **Panel B. Using alternative algorithm: Difference Hash** | | | | | | |
| **Variable** | **N** | **Mean** | **SD** | **P10** | **P50** | **P90** |
| Max. similarity to any of own preceding designs | 5,075 | 0.58 | 0.28 | 0.16 | 0.62 | 0.94 |
| Max. similarity to best of own preceding designs | 3,871 | 0.52 | 0.3 | 0.09 | 0.54 | 0.93 |
| Max. similarity to best of oth. preceding designs | 9,709 | 0.33 | 0.21 | 0.09 | 0.29 | 0.63 |
| Maximum intra-batch similarity | 1,987 | 0.69 | 0.28 | 0.23 | 0.77 | 0.98 |

Notes: Table reports summary statistics on designs' similarity to previously entered designs (both own and competing). Pairwise similarity scores are calculated as described in the text and available for all designs whose digital image could be obtained (96% of entries; refer to the text for an explanation of missing images). The "best" preceding designs are those with the most positive feedback provided prior to the given design. Intra-batch similarity is calculated as the similarity of designs in a given batch to each other, where a design batch is defined to be a set of designs entered by a single player in which each design was entered within 15 minutes of another design in the set. This grouping captures players' tendency to submit multiple designs at once, which are often similar with minor variations on a theme.

Table 4: Similarity to any of player's previous designs

| | (1) | (2) | (3) | (4) |
|---|---|---|---|---|
| Player's prior best rating==5 | 0.286*** | 0.277*** | 0.288*** | 0.276*** |
| | (0.085) | (0.087) | (0.084) | (0.087) |
| * 1+ competing 5-stars | -0.110* | -0.118** | -0.105* | -0.109* |
| | (0.058) | (0.059) | (0.058) | (0.058) |
| * prize value ($100s) | -0.011 | -0.021 | -0.013 | -0.025 |
| | (0.028) | (0.028) | (0.028) | (0.028) |
| Player's prior best rating==4 | 0.100*** | 0.077*** | 0.100*** | 0.076*** |
| | (0.017) | (0.017) | (0.017) | (0.018) |
| Player's prior best rating==3 | 0.039*** | 0.029** | 0.039*** | 0.027* |
| | (0.014) | (0.014) | (0.014) | (0.014) |
| Player's prior best rating==2 | -0.004 | -0.009 | -0.003 | -0.011 |
| | (0.020) | (0.020) | (0.020) | (0.020) |
| One or more competing 5-stars | -0.011 | -0.016 | -0.010 | -0.017 |
| | (0.021) | (0.022) | (0.021) | (0.022) |
| Pct. of contest elapsed | 0.037 | -0.097 | 0.037 | -0.096 |
| | (0.028) | (0.073) | (0.028) | (0.073) |
| Constant | 0.295 | 0.414** | 0.293 | 0.407** |
| | (0.179) | (0.195) | (0.179) | (0.196) |
| N | 5075 | 5075 | 5075 | 5075 |
| $R^2$ | 0.47 | 0.47 | 0.47 | 0.47 |
| Controls | No | Yes | No | Yes |
| Contest FEs | Yes | Yes | Yes | Yes |
| Player FEs | Yes | Yes | Yes | Yes |
| Forthcoming ratings | No | No | Yes | Yes |

Notes: Table shows the effects of feedback on originality. Observations are designs. Dependent variable is a continuous measure of a design's maximal similarity to previous entries in the same contest by the same player, taking values in [0,1], where a value of 1 indicates the design is identical to another. The mean value of this variable is 0.32 (s.d. 0.27). Columns (2) and (4) control for days remaining and number of previous designs by the player and her competitors. Columns (3) and (4) additionally control for the best *forthcoming* rating on the player's not-yet-rated designs. Similarity scores in this table are calculated using a perceptual hash algorithm. Preceding designs/ratings are defined to be those entered/provided at least 60 minutes prior to the given design. *, **, *** represent significance at the 0.1, 0.05, and 0.01 levels, respectively. Standard errors clustered by player in parentheses.

Table 5: Similarity to player's best previously-rated designs & intra-batch similarity

| | Designs | | Batches (uwtd.) | | Batches (wtd.) | |
|---|---|---|---|---|---|---|
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Player's prior best rating==5 | 0.357*** | 0.362*** | 0.223 | 0.238 | 0.254 | 0.285 |
| | (0.097) | (0.102) | (0.311) | (0.304) | (0.304) | (0.296) |
| * 1+ competing 5-stars | -0.206*** | -0.208*** | -0.308* | -0.305* | -0.303* | -0.295* |
| | (0.070) | (0.071) | (0.163) | (0.162) | (0.171) | (0.168) |
| * prize value ($100s) | -0.014 | -0.018 | 0.016 | 0.015 | 0.010 | 0.009 |
| | (0.031) | (0.033) | (0.099) | (0.097) | (0.096) | (0.093) |
| Player's prior best rating==4 | 0.121*** | 0.116*** | 0.054* | 0.065* | 0.064** | 0.086** |
| | (0.031) | (0.032) | (0.032) | (0.037) | (0.032) | (0.038) |
| Player's prior best rating==3 | 0.060** | 0.056** | 0.055 | 0.062* | 0.052 | 0.065* |
| | (0.028) | (0.028) | (0.035) | (0.037) | (0.035) | (0.037) |
| Player's prior best rating==2 | 0.026 | 0.024 | 0.021 | 0.027 | 0.007 | 0.018 |
| | (0.030) | (0.030) | (0.050) | (0.051) | (0.047) | (0.047) |
| One or more competing 5-stars | 0.004 | 0.001 | 0.025 | 0.027 | 0.027 | 0.027 |
| | (0.023) | (0.024) | (0.048) | (0.049) | (0.053) | (0.054) |
| Pct. of contest elapsed | -0.018 | -0.103 | -0.023 | -0.093 | -0.010 | -0.056 |
| | (0.034) | (0.084) | (0.049) | (0.114) | (0.050) | (0.111) |
| Constant | 0.421** | 0.487*** | 0.400*** | 0.507*** | 0.391*** | 0.459*** |
| | (0.168) | (0.187) | (0.066) | (0.148) | (0.060) | (0.146) |
| N | 3871 | 3871 | 1987 | 1987 | 1987 | 1987 |
| $R^2$ | 0.53 | 0.53 | 0.57 | 0.57 | 0.58 | 0.58 |
| Controls | No | Yes | No | Yes | No | Yes |
| Contest FEs | Yes | Yes | Yes | Yes | Yes | Yes |
| Player FEs | Yes | Yes | Yes | Yes | Yes | Yes |

Notes: Table shows the effects of feedback on originality. Observations in Columns (1) and (2) are designs, and dependent variable is a continuous measure of a design's similarity to the highest-rated preceding entry by the same player, taking values in [0,1], where a value of 1 indicates the design is identical to another. The mean value of this variable is 0.28 (s.d. 0.27). Observations in Columns (3) to (6) are design batches, which are defined to be a set of designs by a single player entered into a contest in close proximity (15 minutes), and dependent variable is a continuous measure of intra-batch similarity, taking values in [0,1], where a value of 1 indicates that two designs in the batch are identical. The mean value of this variable is 0.48 (s.d. 0.32). Columns (5) and (6) weight the batch regressions by batch size. All columns control for the time of submission and number of previous designs by the player and her competitors. Similarity scores in this table are calculated using a perceptual hash algorithm. Preceding designs/ratings are defined to be those entered/provided at least 60 minutes prior to the given design. *, **, *** represent significance at the 0.1, 0.05, and 0.01 levels, respectively. Standard errors clustered by player in parentheses.

Table 6: Change in similarity to player's best previously-rated designs

| | (1) | (2) | (3) | (4) | (5) | (6) |
|---|---|---|---|---|---|---|
| Δ(Player's best rating==5) | 0.878*** | 0.927*** | 0.914*** | 0.885*** | 0.929*** | 0.924*** |
| | (0.170) | (0.203) | (0.205) | (0.171) | (0.202) | (0.205) |
| * 1+ competing 5-stars | -0.412*** | -0.418*** | -0.427*** | -0.414*** | -0.418*** | -0.429*** |
| | (0.125) | (0.144) | (0.152) | (0.125) | (0.144) | (0.152) |
| * prize value ($100s) | -0.094** | -0.114** | -0.107** | -0.096** | -0.114** | -0.110** |
| | (0.039) | (0.049) | (0.047) | (0.040) | (0.049) | (0.048) |
| Δ(Player's best rating==4) | 0.282*** | 0.268*** | 0.276*** | 0.283*** | 0.270*** | 0.279*** |
| | (0.065) | (0.073) | (0.079) | (0.065) | (0.073) | (0.079) |
| Δ(Player's best rating==3) | 0.151*** | 0.135** | 0.137** | 0.151*** | 0.136** | 0.138** |
| | (0.058) | (0.065) | (0.069) | (0.058) | (0.065) | (0.069) |
| Δ(Player's best rating==2) | 0.082* | 0.063 | 0.059 | 0.082* | 0.063 | 0.059 |
| | (0.046) | (0.052) | (0.056) | (0.046) | (0.053) | (0.057) |
| One or more competing 5-stars | -0.003 | -0.003 | 0.004 | -0.002 | -0.004 | 0.003 |
| | (0.015) | (0.014) | (0.025) | (0.015) | (0.014) | (0.026) |
| Pct. of contest elapsed | 0.009 | 0.017 | 0.004 | -0.035 | 0.007 | -0.048 |
| | (0.018) | (0.024) | (0.030) | (0.043) | (0.048) | (0.074) |
| Constant | -0.017* | -0.022 | 0.029 | 0.025 | -0.015 | 0.060 |
| | (0.010) | (0.014) | (0.059) | (0.040) | (0.037) | (0.077) |
| N | 2694 | 2694 | 2694 | 2694 | 2694 | 2694 |
| $R^2$ | 0.05 | 0.11 | 0.14 | 0.05 | 0.11 | 0.14 |
| Controls | No | No | No | Yes | Yes | Yes |
| Contest FEs | Yes | No | Yes | Yes | No | Yes |
| Player FEs | No | Yes | Yes | No | Yes | Yes |

Notes: Table shows the effects of feedback on originality. Observations are designs. Dependent variable is a continuous measure of the *change* in designs' similarity to the highest-rated preceding entry by the same player, taking values in [-1,1], where a value of 0 indicates that the player's current design is as similar to her best preceding design as was her previous design, and a value of 1 indicates that the player transitioned fully from experimenting to copying (and a value of -1, the converse). The mean value of this variable is 0.00 (s.d. 0.23). Columns (4) to (6) control for days remaining and number of previous designs by the player and competitors. Similarity scores in this table are calculated using a perceptual hash algorithm. Preceding designs/ratings are defined to be those entered/provided at least 60 minutes prior to the given design. *, **, *** represent significance at the 0.1, 0.05, and 0.01 levels, respectively. Standard errors clustered by player in parentheses.

Table 7: Similarity to player's best not-yet-rated designs (placebo test)

| | Similarity to forthcoming | | | Residual |
|---|---|---|---|---|
| | (1) | (2) | (3) | (4) |
| Player's best forthcoming rating==5 | 0.007 | -0.084 | -0.105 | -0.113 |
| | (0.169) | (0.136) | (0.151) | (0.122) |
| * 1+ competing 5-stars | -0.094 | 0.032 | 0.027 | 0.035 |
| | (0.099) | (0.056) | (0.066) | (0.062) |
| * prize value ($100s) | -0.003 | 0.015 | 0.021 | 0.018 |
| | (0.031) | (0.025) | (0.027) | (0.025) |
| Player's best forthcoming rating==4 | 0.039 | 0.051 | 0.049 | 0.034 |
| | (0.066) | (0.096) | (0.094) | (0.095) |
| Player's best forthcoming rating==3 | 0.080 | 0.049 | 0.051 | 0.036 |
| | (0.052) | (0.088) | (0.088) | (0.088) |
| Player's best forthcoming rating==2 | 0.030 | -0.010 | -0.007 | -0.014 |
| | (0.049) | (0.093) | (0.094) | (0.095) |
| One or more competing 5-stars | -0.080 | -0.013 | -0.010 | -0.013 |
| | (0.097) | (0.110) | (0.117) | (0.119) |
| Pct. of contest elapsed | 0.016 | -0.502 | -0.466 | -0.468 |
| | (0.242) | (0.478) | (0.462) | (0.497) |
| Constant | 0.223 | 0.637 | 0.716 | 0.299 |
| | (0.321) | (0.559) | (0.619) | (0.532) |
| N | 1147 | 577 | 577 | 577 |
| $R^2$ | 0.68 | 0.83 | 0.83 | 0.67 |
| Controls | Yes | Yes | Yes | Yes |
| Contest FEs | Yes | Yes | Yes | Yes |
| Player FEs | Yes | Yes | Yes | Yes |

Notes: Table provides a placebo test of the effects of future feedback on originality. Observations are designs. Dependent variable in Columns (1) to (3) is a continuous measure of a design's similarity to the best design that the player has previously entered that has yet to *but will eventually be* rated, taking values in [0,1], where a value of 1 indicates that the two designs are identical. The mean value of this variable is 0.26 (s.d. 0.25). Under the identifying assumption that future feedback is unpredictable, current choices should be unrelated to forthcoming ratings. Note that a given design's similarity to an earlier, unrated design can be incidental if they are both tweaks on a rated third design. To account for this possibility, Column (2) controls for the given and unrated designs' similarity to the best previously-rated design. Column (3) allows these controls to vary with the highest rating previously received. Dependent variable in Column (4) is the residual from a regression of the dependent variable in the previous columns on these controls. These residuals will be the subset of a given design's similarity to the unrated design that is not explained by jointly-occurring similarity to a third design. All columns control for days remaining and number of previous designs by the player and her competitors. Similarity scores in this table are calculated using a perceptual hash algorithm. Preceding designs/ratings are defined to be those entered/provided at least 60 minutes prior to the given design. *, **, *** represent significance at the 0.1, 0.05, and 0.01 levels, respectively. Standard errors clustered by player in parentheses.

Table 8: Abandonment after a given design, as function of feedback and competition

| | Dependent variable: Abandon after given design | | | |
|---|---|---|---|---|
| | (1) | (2) | (3) | (4) |
| | Linear | Linear | Linear | Logit |
| Player's prior best rating==5 | -0.013 | -0.169*** | -0.146*** | -0.062 |
| | (0.034) | (0.055) | (0.044) | (0.175) |
| * competing 5s | 0.012* | 0.023*** | 0.023*** | 0.059** |
| | (0.007) | (0.009) | (0.008) | (0.030) |
| Player's prior best rating==4 | -0.056*** | -0.114*** | -0.091*** | -0.233*** |
| | (0.016) | (0.023) | (0.023) | (0.082) |
| * competing 5s | 0.015*** | 0.022*** | 0.021*** | 0.087*** |
| | (0.005) | (0.008) | (0.008) | (0.026) |
| Player's prior best rating==3 | -0.026* | -0.029 | -0.004 | -0.060 |
| | (0.015) | (0.019) | (0.019) | (0.071) |
| * competing 5s | 0.014** | 0.022*** | 0.018** | 0.066** |
| | (0.007) | (0.008) | (0.008) | (0.028) |
| Player's prior best rating==2 | -0.023 | 0.011 | 0.029 | -0.052 |
| | (0.022) | (0.026) | (0.027) | (0.107) |
| * competing 5s | 0.009 | 0.031** | 0.026* | 0.044 |
| | (0.010) | (0.014) | (0.013) | (0.044) |
| Player's prior best rating==1 | -0.027 | 0.050 | 0.059 | -0.077 |
| | (0.035) | (0.038) | (0.038) | (0.168) |
| * competing 5s | -0.025* | -0.009 | -0.018 | -0.113* |
| | (0.013) | (0.019) | (0.019) | (0.065) |
| Competing 5-star designs | 0.009** | -0.006* | 0.003 | 0.046** |
| | (0.004) | (0.003) | (0.005) | (0.018) |
| Pct. of contest elapsed | 0.214*** | 0.401*** | 0.383*** | 1.022*** |
| | (0.049) | (0.042) | (0.058) | (0.229) |
| Constant | 0.222*** | 0.085** | 0.021 | -1.423*** |
| | (0.044) | (0.034) | (0.100) | (0.334) |
| N | 11758 | 11758 | 11758 | 11758 |
| $R^2$ | 0.07 | 0.26 | 0.28 | |
| Controls | Yes | Yes | Yes | Yes |
| Contest FEs | Yes | No | Yes | Yes |
| Player FEs | No | Yes | Yes | No |

Notes: Table shows the effects of feedback and competition at the time a design is entered on the probability that player subsequently continues to participate (by entering more designs) versus abandons the contest. Observations are designs. Dependent variable in all columns is an indicator for whether the player stops participating. Columns (1) to (3) estimate linear models with fixed effects; Column (4) estimates a logit model without player fixed effects, which could render the results inconsistent. All columns control for days remaining and number of previous designs by the player and her competitors. Results are qualitatively similar under a proportional hazards model. *, **, *** represent significance at the 0.1, 0.05, and 0.01 levels, respectively. Standard errors clustered by player in parentheses.

Table 9: Normalized panelist ratings on tweaks vs. original designs

| Metric | Outcomes for: Tweaks | Originals | Diff. in means |
|---|---|---|---|
| PCA score of panelist ratings | -0.45 (0.21) | 0.18 (0.15) | 0.64*** $p$=0.008 |
| Average rating by panelists | -0.45 (0.20) | 0.22 (0.14) | 0.67*** $p$=0.004 |
| Median rating by panelists | -0.46 (0.21) | 0.23 (0.15) | 0.69*** $p$=0.005 |
| Max rating by panelists | 1.08 (0.22) | 1.99 (0.17) | 0.91*** $p$=0.001 |
| Disagreement (s.d.) among panelists | 1.34 (0.10) | 1.59 (0.07) | 0.25** $p$=0.019 |

Notes: Table compares professional graphic designers' ratings on tweaks and original designs that received a top rating from contest sponsors. Panelists' ratings were demeaned prior to analysis. The PCA score refers to a design's score along the first component from a principal component component analysis of panelists' ratings. The other summary measures are the mean, median, max, and s.d. of panelists' ratings on a given design. A design is classified as a tweak if its maximal similarity to any previous design by that player is greater than 0.7 and as original if it is less than 0.3. Standard errors in parentheses are provided below each mean, and results from a one-sided test of equality of means is provided to the right. *, **, *** indicate significance at the 0.1, 0.05, and 0.01 levels, respectively. Similarity scores calculated using perceptual hash algorithm. Results are robust to both algorithms and alternative cutoffs for originality.

# Appendix for Online Publication

# A Proofs of Theorems and Additional Figures

**Lemma 1:** When $q \in (\frac{1}{1+\alpha}, \frac{1}{2})$, the gains to exploration over exploitation are decreasing and convex in $\mu$ for small $\mu$, increasing and concave for intermediate $\mu$, and decreasing and convex for large $\mu$. When $q \in \left(\frac{1}{2}, \frac{3\alpha+1}{4\alpha+1}\right)$, they are increasing and convex for small $\mu$ and decreasing and convex for large $\mu$. When $q > \frac{3\alpha+1}{4\alpha+1}$, they are increasing and concave for small $\mu$ and decreasing and convex for large $\mu$. When $q < \frac{1}{1+\alpha}$, they are decreasing and convex for small $\mu$ and increasing and concave for large $\mu$. In every case, the gains are zero when $\mu = 0$; when $q > \frac{1}{1+\alpha}$ ($q < \frac{1}{1+\alpha}$), they approach zero from above (below) as $\mu \to \infty$, holding $\beta_1$ fixed.

**Proof:**

Part (i): Low $\mu$. As $\mu \longrightarrow 0$:

$$\begin{aligned} qF\left(\beta_2^H\right) + (1-q)F\left(\beta_2^L\right) - F\left(\beta_1\right) &= q\left(\frac{(1+\alpha)\beta_1}{(1+\alpha)\beta_1+\mu}\right) + (1-q)\left(\frac{\left(1+\frac{1}{\alpha}\right)\beta_1}{\left(1+\frac{1}{\alpha}\right)\beta_1+\mu}\right) - \left(\frac{2\beta_1}{2\beta_1+\mu}\right) \\ &\longrightarrow q\left(\frac{(1+\alpha)\beta_1}{(1+\alpha)\beta_1}\right) + (1-q)\left(\frac{\left(1+\frac{1}{\alpha}\right)\beta_1}{\left(1+\frac{1}{\alpha}\right)\beta_1}\right) - \left(\frac{2\beta_1}{2\beta_1}\right) \\ &= q + (1-q) - 1 = 0 \end{aligned}$$

$$\begin{aligned} \frac{\partial}{\partial\mu}\left[qF\left(\beta_2^H\right) + (1-q)F\left(\beta_2^L\right) - F\left(\beta_1\right)\right] &= q\left(\frac{-(1+\alpha)\beta_1}{\left((1+\alpha)\beta_1+\mu\right)^2}\right) + (1-q)\left(\frac{-\left(1+\frac{1}{\alpha}\right)\beta_1}{\left(\left(1+\frac{1}{\alpha}\right)\beta_1+\mu\right)^2}\right) + \frac{2\beta_1}{(2\beta_1+\mu)^2} \\ &\longrightarrow q\left(\frac{-1}{(1+\alpha)\beta_0}\right) + (1-q)\left(\frac{-1}{\left(1+\frac{1}{\alpha}\right)\beta_0}\right) + \frac{1}{2\beta_0} \\ &= \frac{-2q\left(1+\frac{1}{\alpha}\right) - 2(1-q)(1+\alpha) + (1+\alpha)\left(1+\frac{1}{\alpha}\right)}{2(1+\alpha)\left(1+\frac{1}{\alpha}\right)\beta_1} \\ &= \frac{-2\left(q+\frac{1}{\alpha}q\right) - 2(1-q+\alpha-\alpha q) + \left(1+\alpha+\frac{1}{\alpha}+1\right)}{2(1+\alpha)\left(1+\frac{1}{\alpha}\right)\beta_1} \\ &= \frac{2\alpha q - 2\frac{1}{\alpha}q - \alpha + \frac{1}{\alpha}}{2(1+\alpha)\left(1+\frac{1}{\alpha}\right)\beta_1} = \frac{\left(\alpha^2-1\right)(2q-1)}{2(1+\alpha)^2\beta_1} \longrightarrow \begin{cases} 0^+ & \text{if } q > \frac{1}{2} \\ 0^- & \text{if } q < \frac{1}{2} \end{cases} \end{aligned}$$

$$\frac{\partial^2}{\partial\mu^2}\left[qF\left(\beta_2^H\right) + (1-q)F\left(\beta_2^L\right) - F\left(\beta_1\right)\right] = q\left(\frac{2(1+\alpha)\beta_1}{\left((1+\alpha)\beta_1+\mu\right)^3}\right) + (1-q)\left(\frac{2\left(1+\frac{1}{\alpha}\right)\beta_1}{\left(\left(1+\frac{1}{\alpha}\right)\beta_1+\mu\right)^3}\right) - \frac{4\beta_1}{(2\beta_1+\mu)^3},$$

$$\text{numerator} \longrightarrow \left(q\cdot 4(1+\alpha)\left(1+\frac{1}{\alpha}\right)^3 + (1-q)\cdot 4\left(1+\frac{1}{\alpha}\right)(1+\alpha)^3 - (1+\alpha)^3\left(1+\frac{1}{\alpha}\right)^3\right)4\beta_0^7 \begin{cases} > 0 & \text{if } q < \frac{3\alpha+1}{4(\alpha+1)} \\ < 0 & \text{if } q > \frac{3\alpha+1}{4(\alpha+1)} \end{cases}$$

Part (ii): High $\mu$. As $\mu \longrightarrow \infty$:

$$\begin{aligned} qF\left(\beta_2^H\right) + (1-q)\,F\left(\beta_2^L\right) - F\left(\beta_1\right) &= q\left(\frac{(1+\alpha)\,\beta_1}{(1+\alpha)\,\beta_1+\mu}\right) + (1-q)\left(\frac{\left(1+\frac{1}{\alpha}\right)\beta_1}{\left(1+\frac{1}{\alpha}\right)\beta_1+\mu}\right) - \left(\frac{2\beta_1}{2\beta_1+\mu}\right) \\ &\longrightarrow \frac{1}{\mu}\left(q\,(1+\alpha)\,\beta_1 + (1-q)\left(1+\frac{1}{\alpha}\right)\beta_1 - 2\beta_1\right) \\ &= \frac{1}{\mu}\beta_1\left(\alpha q - \frac{1}{\alpha}q - 1 + \frac{1}{\alpha}\right) \\ &= \frac{1}{\alpha\mu}\beta_1\left(\left(\alpha^2-1\right)q - (\alpha-1)\right) \\ &= \frac{\alpha-1}{\alpha\mu}\beta_1\left((1+\alpha)\,q-1\right) \longrightarrow \begin{cases} 0^+ & \text{if } q > \frac{1}{1+\alpha} \\ 0^- & \text{if } q < \frac{1}{1+\alpha} \end{cases} \end{aligned}$$

$$\begin{aligned} \frac{\partial}{\partial\mu}\left[qF\left(\beta_2^H\right) + (1-q)\,F\left(\beta_2^L\right) - F\left(\beta_1\right)\right] &= q\left(\frac{-(1+\alpha)\,\beta_1}{\left((1+\alpha)\,\beta_1+\mu\right)^2}\right) + (1-q)\left(\frac{-\left(1+\frac{1}{\alpha}\right)\beta_1}{\left(\left(1+\frac{1}{\alpha}\right)\beta_1+\mu\right)^2}\right) + \frac{2\beta_1}{\left(2\beta_1+\mu\right)^2} \\ &\longrightarrow \frac{1}{\mu^2}\left(-q\,(1+\alpha)\,\beta_1 - (1-q)\left(1+\frac{1}{\alpha}\right)\beta_1 + 2\beta_1\right) \\ &= \frac{1}{\mu^2}\beta_1\left(-\left(q+\alpha q\right) - \left(1-q+\frac{1}{\alpha}-\frac{1}{\alpha}q\right) + 2\right) \\ &= \frac{1}{\mu^2}\beta_1\left(-\alpha q + \frac{1}{\alpha}q + 1 - \frac{1}{\alpha}\right) \\ &= \frac{1}{\alpha\mu^2}\beta_1\left(-\left(\alpha^2-1\right)q + (\alpha-1)\right) \\ &= \frac{\alpha-1}{\alpha\mu^2}\beta_1\left(-\left(1+\alpha\right)q+1\right)\beta_1 \longrightarrow \begin{cases} 0^+ & \text{if } q < \frac{1}{1+\alpha} \\ 0^- & \text{if } q > \frac{1}{1+\alpha} \end{cases} \end{aligned}$$

$$\frac{\partial^2}{\partial\mu^2}\left[qF\left(\beta_2^H\right) + (1-q)\,F\left(\beta_2^L\right) - F\left(\beta_1\right)\right] = q\left(\frac{2\,(1+\alpha)\,\beta_1}{\left((1+\alpha)\,\beta_1+\mu\right)^3}\right) + (1-q)\left(\frac{2\left(1+\frac{1}{\alpha}\right)\beta_1}{\left(\left(1+\frac{1}{\alpha}\right)\beta_1+\mu\right)^3}\right) - \frac{4\beta_1}{\left(2\beta_1+\mu\right)^3},$$

$$\text{numerator} \longrightarrow \left(q\cdot(1+\alpha) + (1-q)\cdot\left(1+\tfrac{1}{\alpha}\right) - 2\right)2\beta_1\mu^6 \begin{cases} > 0 & \text{if } q > \frac{1}{1+\alpha} \\ < 0 & \text{if } q < \frac{1}{1+\alpha} \end{cases}$$

Taken together, these asymptotics generate a curve with the shape described.

**Proposition 1:** When $q > \frac{1}{1+\alpha}$, there exists a unique level of competition $\mu_2^*$ at which the gains to exploration, relative to exploitation, are maximized.

**Proof:** Existence follows from lemma and continuity of the success function. Since the difference of the success function under exploration and exploitation is quadratic in $\mu$, it has at most two real roots, one of which is shown below to be zero, the other of which is shown to be positive if $q \in \left(\frac{1}{1+\alpha}, \frac{1}{2}\right)$ and negative otherwise. Given the shape described by the lemma, the value at which this difference is maximized must be unique.

To find the roots, set $qF\left(\beta_2^H\right) + (1-q)F\left(\beta_2^L\right) - F\left(\beta_1\right) = 0$ and solve for $\mu$:

$$
\begin{aligned}
0 &= q\left(\frac{(1+\alpha)\beta_1}{(1+\alpha)\beta_1+\mu}\right) + (1-q)\left(\frac{\left(1+\frac{1}{\alpha}\right)\beta_1}{\left(1+\frac{1}{\alpha}\right)\beta_1+\mu}\right) - \left(\frac{2\beta_1}{2\beta_1+\mu}\right) \\
&= q(1+\alpha)\beta_1\left[\left(\left(1+\frac{1}{\alpha}\right)\beta_1+\mu\right)(2\beta_1+\mu)\right] \\
&\quad + (1-q)\left(1+\frac{1}{\alpha}\right)\beta_1\left[((1+\alpha)\beta_1+\mu)(2\beta_1+\mu)\right] \\
&\quad - 2\beta_1\left[((1+\alpha)\beta_1+\mu)\left(\left(1+\frac{1}{\alpha}\right)\beta_1+\mu\right)\right] \\
&= \mu^2\beta_1\left[q(1+\alpha) + (1-q)\left(1+\frac{1}{\alpha}\right) - 2\right] \\
&\quad + \mu\beta_1^2\left[q\left(2+\alpha+\frac{1}{\alpha}\right) + (1-q)\left(2+\alpha+\frac{1}{\alpha}\right) + 2q(1+\alpha) + 2(1-q)\left(1+\frac{1}{\alpha}\right) - 2\left(2+\alpha+\frac{1}{\alpha}\right)\right] \\
&\quad + \beta_1^3\left[2q\left(2+\alpha+\frac{1}{\alpha}\right) + 2(1-q)\left(2+\alpha+\frac{1}{\alpha}\right) - 2\left(2+\alpha+\frac{1}{\alpha}\right)\right] \\
&= a\beta_1\mu^2 + b\beta_1^2\mu + c\beta_1^3\,,
\end{aligned}
$$

where

$$
a = q + \alpha q + (1-q) + \frac{1}{\alpha}(1-q) - 2 = \alpha q + \frac{1}{\alpha}(1-q) - 1 = \frac{1}{\alpha}(\alpha-1)((1+\alpha)q-1) \begin{cases} > 0 & \text{if } q > \frac{1}{1+\alpha} \\ < 0 & \text{if } q < \frac{1}{1+\alpha} \end{cases}
$$

$$
b = 2q + 2\alpha q + 2(1-q) + 2\frac{1}{\alpha}(1-q) - \left(2+\alpha+\frac{1}{\alpha}\right) = \frac{1}{\alpha}\left(\alpha^2-1\right)(2q-1) \begin{cases} > 0 & \text{if } q > \frac{1}{2} \\ < 0 & \text{if } q < \frac{1}{2} \end{cases}
$$

$$
c = 0
$$

By the quadratic formula, the roots are thus:

$$
\frac{-\left(b\beta_1^2\right) \pm \sqrt{\left(b\beta_1^2\right)^2 - 0}}{2\left(a\beta_1\right)} = \frac{-\left(b\beta_1\right) \pm -\left(b\beta_1\right)}{2a} = \begin{cases} -\beta_1\frac{b}{a} \\ 0 \end{cases}
$$

When $q < \frac{1}{1+\alpha}$, $a < 0$ and $b < 0$, and the non-zero root is negative. When $q \in \left(\frac{1}{1+\alpha}, \frac{1}{2}\right)$, $a > 0$ and $b < 0$, and the non-zero root is positive. When $q > \frac{1}{2}$, $a > 0$ and $b > 0$, and the non-zero root is negative.

**Corollary:** When $q < \frac{1}{1+\alpha}$, exploration will never be preferred to exploitation.

**Proof:** Follows from lemma, continuity of the success function, and results from the previous proof showing that when $q < \frac{1}{1+\alpha}$, there is no positive root for the difference of the success function for exploration and exploitation, such that this difference never becomes positive.

**Lemma 2:** The agent's gains to exploration over abandonment are increasing and concave in $\mu$ when $\mu$ is small and decreasing and convex when $\mu$ is large. The gains are zero when $\mu = 0$ and approach zero from above as $\mu \to \infty$, holding $\beta_1$ fixed.

**Proof:**

Part (i): Low $\mu$. As $\mu \longrightarrow 0$:

$$\begin{aligned} qF\left(\beta_2^H\right) + (1-q)F\left(\beta_2^L\right) - F(0) &= q\left(\frac{(1+\alpha)\beta_1}{(1+\alpha)\beta_1+\mu}\right) + (1-q)\left(\frac{\left(1+\frac{1}{\alpha}\right)\beta_1}{\left(1+\frac{1}{\alpha}\right)\beta_1+\mu}\right) - \left(\frac{\beta_1}{\beta_1+\mu}\right) \\ &\longrightarrow q\left(\frac{(1+\alpha)\beta_1}{(1+\alpha)\beta_1}\right) + (1-q)\left(\frac{\left(1+\frac{1}{\alpha}\right)\beta_1}{\left(1+\frac{1}{\alpha}\right)\beta_1}\right) - \left(\frac{\beta_1}{\beta_1}\right) \\ &= q + (1-q) - 1 = 0 \end{aligned}$$

$$\begin{aligned} \frac{\partial}{\partial\mu}\left[qF\left(\beta_2^H\right) + (1-q)F\left(\beta_2^L\right) - F(0)\right] &= q\left(\frac{-(1+\alpha)\beta_1}{((1+\alpha)\beta_1+\mu)^2}\right) + (1-q)\left(\frac{-\left(1+\frac{1}{\alpha}\right)\beta_1}{\left(\left(1+\frac{1}{\alpha}\right)\beta_1+\mu\right)^2}\right) + \frac{\beta_1}{(\beta_1+\mu)^2} \\ &\longrightarrow q\left(\frac{-1}{(1+\alpha)\beta_0}\right) + (1-q)\left(\frac{-1}{\left(1+\frac{1}{\alpha}\right)\beta_0}\right) + \frac{1}{\beta_0} \\ &= \frac{-q\left(1+\frac{1}{\alpha}\right) - (1-q)(1+\alpha) + (1+\alpha)\left(1+\frac{1}{\alpha}\right)}{(1+\alpha)\left(1+\frac{1}{\alpha}\right)\beta_1} \\ &= \frac{-\left(q+\frac{1}{\alpha}q\right) - (1-q+\alpha-\alpha q) + \left(1+\alpha+\frac{1}{\alpha}+1\right)}{(1+\alpha)\left(1+\frac{1}{\alpha}\right)\beta_1} \\ &= \frac{\alpha q - \frac{1}{\alpha}q + \alpha + \frac{1}{\alpha}}{(1+\alpha)\left(1+\frac{1}{\alpha}\right)\beta_1} = \frac{\left(\alpha^2-1\right)q + \left(1+\alpha^2\right)}{(1+\alpha)^2\beta_1} \longrightarrow 0^+ \end{aligned}$$

$$\frac{\partial^2}{\partial\mu^2}\left[qF\left(\beta_2^H\right) + (1-q)F\left(\beta_2^L\right) - F(0)\right] = q\left(\frac{2(1+\alpha)\beta_1}{((1+\alpha)\beta_1+\mu)^3}\right) + (1-q)\left(\frac{2\left(1+\frac{1}{\alpha}\right)\beta_1}{\left(\left(1+\frac{1}{\alpha}\right)\beta_1+\mu\right)^3}\right) - \frac{2\beta_1}{(\beta_1+\mu)^3},$$

$$\text{numerator} \longrightarrow \left(q\cdot(1+\alpha)\left(1+\tfrac{1}{\alpha}\right)^3 + (1-q)\cdot\left(1+\tfrac{1}{\alpha}\right)(1+\alpha)^3 - (1+\alpha)^3\left(1+\tfrac{1}{\alpha}\right)^3\right)2\beta_1^7 < 0$$

Part (ii): High $\mu$. As $\mu \longrightarrow \infty$:

$$\begin{aligned} qF\left(\beta_2^H\right) + (1-q)F\left(\beta_2^L\right) - F(0) &= q\left(\frac{(1+\alpha)\beta_1}{(1+\alpha)\beta_1+\mu}\right) + (1-q)\left(\frac{\left(1+\frac{1}{\alpha}\right)\beta_1}{\left(1+\frac{1}{\alpha}\right)\beta_1+\mu}\right) - \left(\frac{\beta_1}{\beta_1+\mu}\right) \\ &\longrightarrow \frac{1}{\mu}\left(q(1+\alpha)\beta_1 + (1-q)\left(1+\frac{1}{\alpha}\right)\beta_1 - \beta_1\right) \\ &= \frac{1}{\mu}\beta_1\left(\alpha q - \frac{1}{\alpha}q + \frac{1}{\alpha}\right) \\ &= \frac{1}{\alpha\mu}\beta_1\left(\left(\alpha^2-1\right)q+1\right) \longrightarrow 0^+ \end{aligned}$$

$$\begin{aligned} \frac{\partial}{\partial\mu}\left[qF\left(\beta_2^H\right) + (1-q)F\left(\beta_2^L\right) - F(0)\right] &= q\left(\frac{-(1+\alpha)\beta_1}{((1+\alpha)\beta_1+\mu)^2}\right) + (1-q)\left(\frac{-\left(1+\frac{1}{\alpha}\right)\beta_1}{\left(\left(1+\frac{1}{\alpha}\right)\beta_1+\mu\right)^2}\right) + \frac{\beta_1}{(\beta_1+\mu)^2} \\ &\longrightarrow \frac{1}{\mu^2}\left(-q(1+\alpha)\beta_1 - (1-q)\left(1+\frac{1}{\alpha}\right)\beta_1 + \beta_1\right) \\ &= \frac{1}{\mu^2}\beta_1\left(-(q+\alpha q) - \left(1-q+\frac{1}{\alpha}-\frac{1}{\alpha}q\right)+1\right) \\ &= \frac{1}{\mu^2}\beta_1\left(-\alpha q + \frac{1}{\alpha}q - \frac{1}{\alpha}\right) \\ &= \frac{1}{\alpha\mu^2}\beta_1\left(-\left(\alpha^2-1\right)q-1\right) \longrightarrow 0^- \end{aligned}$$

$$\frac{\partial^2}{\partial\mu^2}\left[qF\left(\beta_2^H\right) + (1-q)F\left(\beta_2^L\right) - F(0)\right] = q\left(\frac{2(1+\alpha)\beta_1}{((1+\alpha)\beta_1+\mu)^3}\right) + (1-q)\left(\frac{2\left(1+\frac{1}{\alpha}\right)\beta_1}{\left(\left(1+\frac{1}{\alpha}\right)\beta_1+\mu\right)^3}\right) - \frac{2\beta_1}{(\beta_1+\mu)^3},$$

$$\text{numerator} \longrightarrow \left(q\cdot(1+\alpha) + (1-q)\cdot\left(1+\tfrac{1}{\alpha}\right) - 1\right)2\beta_1\mu^6 > 0$$

Taken together, these asymptotics generate a curve with the shape described.

**Proposition 2:** For all values of $q$, there exists a unique level of competition $\mu_1^*$ at which the gains to exploration, relative to abandonment, are maximized.

**Proof:** Existence follows from lemma and continuity of the success function. Since the difference of the success function under exploration and abandonment is quadratic in $\mu$, it has at most two real roots, one of which is shown below to be zero, the other of which is shown to be negative. Given the shape described by the lemma, the value at which this difference is maximized must be unique.

To find the roots, set $qF\left(\beta_2^H\right) + (1-q)\,F\left(\beta_2^L\right) - F(0) = 0$ and solve for $\mu$:

$$
\begin{aligned}
0 &= q\left(\frac{(1+\alpha)\,\beta_1}{(1+\alpha)\,\beta_1+\mu}\right) + (1-q)\left(\frac{\left(1+\frac{1}{\alpha}\right)\beta_1}{\left(1+\frac{1}{\alpha}\right)\beta_1+\mu}\right) - \left(\frac{\beta_1}{\beta_1+\mu}\right) \\
&= q\,(1+\alpha)\,\beta_1\left[\left(\left(1+\frac{1}{\alpha}\right)\beta_1+\mu\right)(\beta_1+\mu)\right] \\
&\quad + (1-q)\left(1+\frac{1}{\alpha}\right)\beta_1\left[((1+\alpha)\,\beta_1+\mu)\,(\beta_1+\mu)\right] \\
&\quad - \beta_1\left[((1+\alpha)\,\beta_1+\mu)\left(\left(1+\frac{1}{\alpha}\right)\beta_1+\mu\right)\right] \\
&= \mu^2\beta_1\left[q\,(1+\alpha)+(1-q)\left(1+\frac{1}{\alpha}\right)-1\right] \\
&\quad + \mu\beta_1^2\left[q\left(2+\alpha+\frac{1}{\alpha}\right)+(1-q)\left(2+\alpha+\frac{1}{\alpha}\right)+q\,(1+\alpha)+(1-q)\left(1+\frac{1}{\alpha}\right)-\left(2+\alpha+\frac{1}{\alpha}\right)\right] \\
&\quad + \beta_1^3\left[q\left(2+\alpha+\frac{1}{\alpha}\right)+(1-q)\left(2+\alpha+\frac{1}{\alpha}\right)-\left(2+\alpha+\frac{1}{\alpha}\right)\right] \\
&= a\beta_1\mu^2 + b\beta_1^2\mu + c\beta_1^3\,,
\end{aligned}
$$

where

$$
a = q+\alpha q+(1-q)+\frac{1}{\alpha}(1-q)-1 = \alpha q+\frac{1}{\alpha}(1-q) = \frac{1}{\alpha}\left(\left(\alpha^2-1\right)q+1\right)\begin{cases} >0 & \text{if } q<\frac{1}{1-\alpha^2} \\ <0 & \text{if } q>\frac{1}{1-\alpha^2}\end{cases}
$$

$$
b = q+\alpha q+(1-q)+\frac{1}{\alpha}(1-q) = \alpha q+\frac{1}{\alpha}(1-q)+1 = \frac{1}{\alpha}(\alpha+1)\left((\alpha-1)\,q+1\right)\begin{cases} >0 & \text{if } q<\frac{1}{1-\alpha} \\ <0 & \text{if } q>\frac{1}{1-\alpha}\end{cases}
$$

$$
c = 0
$$

By the quadratic formula, the roots are thus:

$$
\frac{-\left(b\beta_1^2\right)\pm\sqrt{\left(b\beta_1^2\right)^2-0}}{2\,(a\beta_1)} = \frac{-\left(b\beta_1\right)\pm-\left(b\beta_1\right)}{2a} = \begin{cases} -\beta_1\frac{b}{a} \\ 0\end{cases}
$$

Since $\alpha$ is greater than one, $a<0$ and $b<0$. Thus the non-zero root is negative.

**Proposition 3:** At very low and very high $\mu$, the IR constraint binds: the next-best alternative to exploration is abandonment. At intermediate $\mu$, the IC constraint binds: the next-best alternative is exploitation.

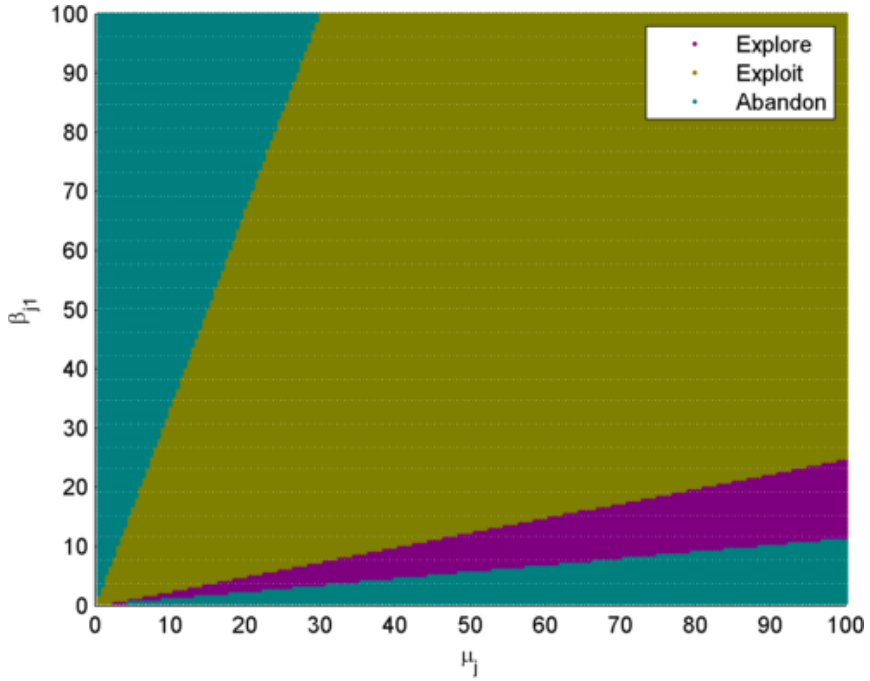
**Proof:** Lemma 1 can be used to characterize the shape of the gains to exploration versus abandonment and exploitation versus abandonment, since in this model, exploitation is a special case of exploration, with $\alpha = 1$. The proof to Lemma 1 establishes that the gains to exploitation are zero when $\mu = 0$, increasing for small $\mu$, decreasing for large $\mu$, and approach zero from above as $\mu \to \infty$. Provided the prize-normalized cost of exploitation is not greater than the maximum of this function, the payoffs to exploitation will begin negative, turn positive, and finish negative, implying that abandonment (the IR constraint) is binding to exploration at low and high $\mu$ and exploitation (the IC constraint) is binding at intermediate $\mu$.

**Proposition 4:** When $q > \frac{1}{1+\alpha}$, there exists a unique level of competition $\mu^* \in [\mu_1^*, \mu_2^*]$ at which the gains to exploration are maximized relative to the player's next-best alternative.

**Proof:** Result follows from the first three propositions.

Figure A.1 shows the regions in $(\mu_j, \beta_{j1})$-space where a player prefers exploring, exploiting, and abandoning, for the parametrization used throughout the paper. The figure shows that for any fixed, own first-round performance $(\beta_{j1})$, the player's preferred action goes from abandonment, to exploitation, to exploration, back to abandonment as competition $(\mu_j)$ grows from zero to infinity (left to right). The example can be interpreted alternatively as applying to (i) the marginal player, (ii) any player with a fixed belief over the final level of competition, or (iii) any player, in partial equilibrium – the key assumption being an exogenous $\mu_j$.

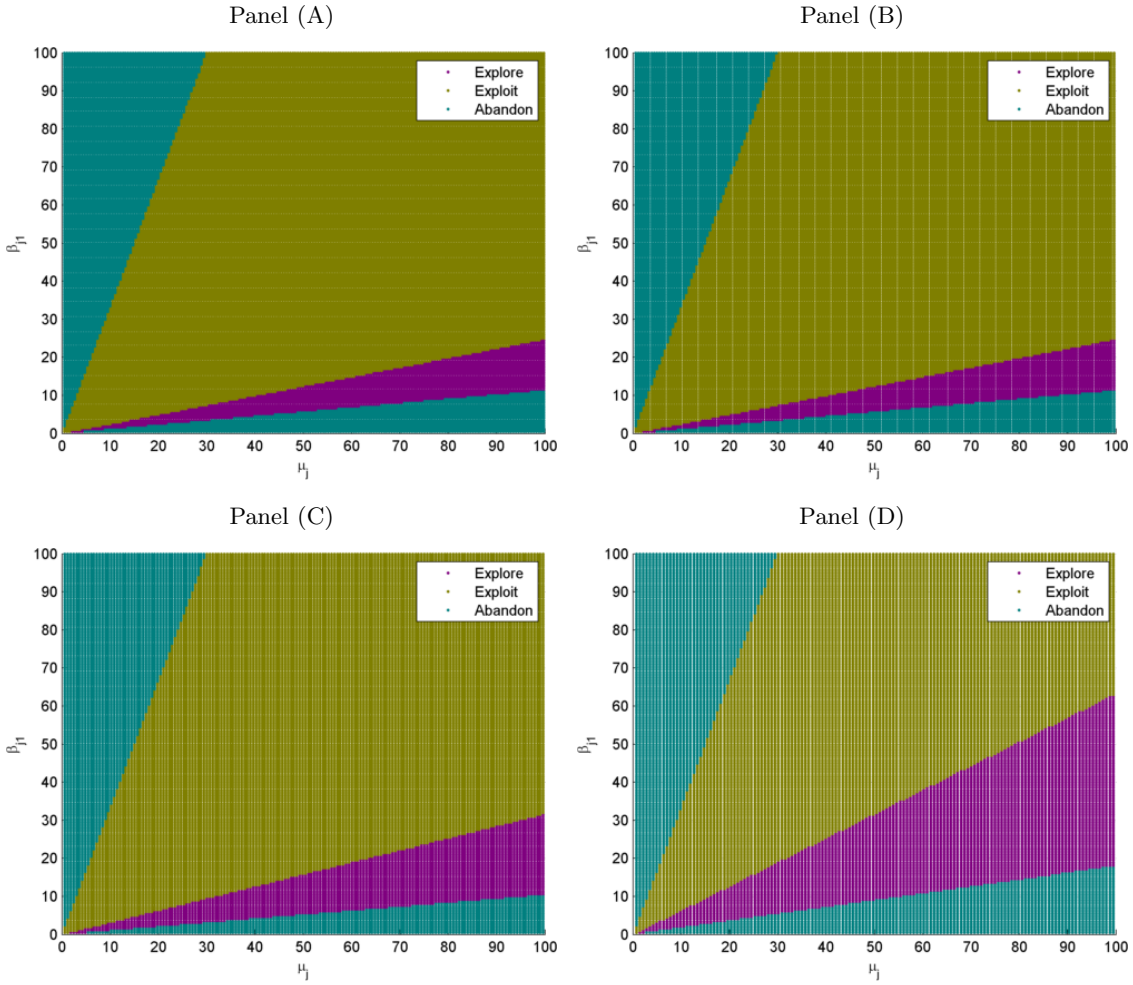
Figure A.1: Exploration, exploitation, and abandonment regions: marginal player




Notes: Figure plots the regions in $(\mu_j, \beta_{j1})$-space where the marginal player prefers each action for the parametrization used in the paper: $q = 0.33$, $\alpha = 9$, $c = 200$k, $d = 300$k, $P = 2$m. For any fixed, own first-round performance $(\beta_{j1})$, the player's preferred action changes from abandonment, to exploitation, to exploration, to abandonment as the cumulative competition grows from zero to infinity.

Figure A.2 performs a similar exercise for a strategic second-to-last player who takes into account the next player's best response. Panel (A) shows a simulation where the remaining competitors first-round performance comprises 5 percent of cumulative competition; Panel (B), 10 percent; Panel (C), 20 percent; and Panel (D), 40 percent. The sequence of preferred actions as competition increases is the same as that in the previous igure, provided that the remaining competitor is a fraction of the cumulative competition.

Figure A.2: Exploration, exploitation, and abandonment regions: inframarginal player




Notes: Figure plots the regions in $(\mu_j, \beta_{j1})$-space where the player with the second-to-last move prefers each action for the parametrization used in the paper ($q = 0.33$, $\alpha = 9$, $c = 200$k, $d = 300$k, $P = 2$m), taking into account the next player's best response. Panel (A) plots the case where the remaining competitor's first-round performance is five percent of the cumulative competition; Panel (B), 10 percent; Panel (C), 20 percent; and Panel (D), 40 percent. The incentives are similar to those for marginal or non-strategic players provided that the remaining competitor is a fraction of the cumulative competition.

# B Dataset Construction

Data were collected on all logo design contests with open (i.e., public) bidding that launched the week of September 3 to 9, 2013, and every three weeks thereafter through the week of November 5 to 11, 2013. Conditional on open bidding, this sample is effectively randomly drawn. The sample used in the paper is further restricted to contests with a single, winner-take-all prize and with no mid-contest rule changes such as prize increases, deadline extensions, and early endings. The sample also excludes one contest that went dormant and resumed after several weeks, as well as a handful of contests whose sponsors simply stopped participating and were never heard from again. These restrictions cause 146 contests to be dropped from the sample. The final dataset includes 122 contests, 4,050 contest-players, and 11,758 designs.[1]

To collect the data, I developed an automated script to scan these contests once an hour for new submissions, save a copy of each design for analysis, and record their owners' identity and performance history from a player profile. I successfully obtained the image files for 96 percent of designs in the final sample. The remaining designs were entered and withdrawn before they could be observed (recall that players can withdraw designs they have entered into a contest, though this option is rarely exercised and can be reversed at the request of a sponsor). All other data were automatically acquired at the conclusion of each contest, once the prize was awarded or the sponsor exercised its outside option of a refund.

## B.1 Variables

The dataset includes information on the characteristics of contests, contest-players, and designs:

- Contest-level variables include: the contest sponsor, features of the project brief (title, description, sponsor industry, materials to be included in logo), start and end dates, the prize amount (and whether committed), and the number of players and designs of each rating.
- Contest-player-level variables include: the player's self-reported country, his/her experience in previous contests on the platform (number of contests and designs entered, contests won), and that player's participation and performance in the given contest.
- Design-level variables include: the design's owner, its submission time and order of entry, the feedback it received, the time at which this feedback was given, and whether it was eventually withdrawn. For designs with images acquired, I calculate originality using the procedures described in the next section. The majority of the analysis occurs at the design level.

[1] While 268 contests were originally sampled, only 122 of these survived the filters described above. The number of contests sampled was constrained by high costs of data collection, which was performed in real-time by making round-the-clock hourly scans for new activity without putting a heavy strain on the platform's servers.

Note that designs are occasionally re-rated: five percent of all rated designs are re-rated an average of 1.2 times each. Of these, 14 percent are given their original rating, and 83 percent are re-rated within 1 star of the original rating. I treat the first rating on each design to be the most informative, objective measure of quality, since research suggests first instincts tend to be most reliable and ratings revisions are likely made relative to other designs in the contest rather than an objective benchmark.

## B.2 Image Comparison Algorithms

This paper uses two distinct algorithms to calculate pairwise similarity scores. One is a perceptual hash algorithm, which creates a digital signature (hash) for each image from its lowest frequency content. As the name implies, a perceptual hash is designed to imitate human perception. The second algorithm is a difference hash, which creates the hash from pixel intensity gradients.

I implement the perceptual hash algorithm and calculate pairwise similarity scores using a variant of the procedure described by the Hacker Factor blog.[2] This requires six steps:

1. Resize each image to 32x32 pixels and convert to grayscale.
2. Compute the discrete cosine transform (DCT) of each image. The DCT is a widely-used transform in signal processing that expresses a finite sequence of data points as a linear combination of cosine functions oscillating at different frequencies. By isolating low frequency content, the DCT reduces a signal (in this case, an image) to its underlying structure. The DCT is broadly used in digital media compression, including MP3 and JPEG formats.
3. Retain the upper-left 16x16 DCT coefficients and calculate the average value, excluding first term.
4. Assign 1s to grid cells with above-average DCT coefficients, and 0s elsewhere.
5. Reshape to 256 bit string; this is the image's digital signature (hash).
6. Compute the Hamming distance between the two hashes and divide by 256.

The similarity score is obtained by subtracting this fraction from one. In a series of sensitivity tests, the perceptual hash algorithm was found to be strongly invariant to transformations in scale, aspect ratio, brightness, and contrast, albeit not rotation. As described, the algorithm will perceive two images that have inverted colors but are otherwise identical to be perfectly dissimilar. I make the algorithm robust to color inversion by comparing each image against the regular and inverted hash of its counterpart in the pair,

[2]See http://www.hackerfactor.com/blog/archives/432-Looks-Like-It.html.

taking the maximum similarity score, and rescaling so that the scores remain in [0,1]. The resulting score is approximately the absolute value correlation of two images' content.

I follow a similar procedure outlined by the same blog[3] to implement the difference hash algorithm and calculate an alternative set of similarity scores for robustness checks:

1. Resize each image to 17x16 pixels and convert to grayscale.
2. Calculate horizontal gradient as the change in pixel intensity from left to right, returning a 16x16 grid (note: top to bottom is an equally valid alternative)
3. Assign 1s to grid cells with positive gradient, 0s to cells with negative gradient.
4. Reshape to 256 bit string; this is the image's digital signature (hash).
5. Compute the Hamming distance between the two hashes and divide by 256.

The similarity score is obtained by subtracting this fraction from one. In sensitivity tests, the difference hash algorithm was found to be highly invariant to transformations in scale and aspect ratio, potentially sensitive to changes in brightness and contrast, and very sensitive to rotation. I make the algorithm robust to color inversion using a procedure identical to that described for the perceptual hash.

Though the perceptual and difference hash algorithms are both conceptually and mathematically distinct, and the resulting similarity scores are only modestly correlated ($\rho = 0.38$), the empirical results of Section 3 are qualitatively and quantitatively similar under either algorithm. This consistency is reassurance that the patterns found are not simply an artifact of an arcane image processing algorithm; rather, they appear to be generated by the visual content of the images themselves.

### B.3 Why use algorithms?

There are two advantages to using algorithms over human judges. The first is that the algorithms can be directed to evaluate specific features of an image and thus provide a consistent, objective measure of similarity, whereas individuals may be attuned to different features and can have different perceptions of similarity in practice (Tirilly et al. 2012). This argument is supported by a pilot study I attempted using Amazon Mechanical Turk, in which I asked participants to rate the similarity of pairs of images they were shown; the results (not provided here) were generally very noisy, except in cases of nearly identical images, in which case the respondents tended to agree that they were similar. The second advantage of algorithms is more obvious: they are cheap, taking only seconds to execute a comparison.

[3]See http://www.hackerfactor.com/blog/archives/529-Kind-of-Like-That.html.

The obvious concern is that computer vision may not be comparable to human perception, in which case the measures are not relevant to human behavior. This concern is mitigated by the fact that they are used to approximate human perception in commercial software, as well as two pieces of evidence from the above studies: (i) when two images are similar, humans and algorithms tend to agree, and (ii) when two images are dissimilar, neither humans nor algorithms find them similar, but they may also disagree on the degree of dissimilarity. Thus, human and computer judgment tend to align within coarse categories, especially at extremes – which is the margin of variation that matters most for this paper.

The evidence of disagreement in subjects' assessments of similarity nevertheless raises a deeper question: is it sensible to apply a uniform similarity measure at all? Squire and Pun (1997) find that *expert* subjects' assessments of similarity tend to agree at all levels, and the designers in this paper could reasonably be classified as visual experts. Even so, it is reassuring that the results throughout this paper are similar in sign, significance, and magnitude under two fundamentally different algorithms, and thus emerge no matter which features we choose to focus on when measuring similarity.

## B.4 How do the algorithms perform?

In my own experience browsing the designs in the dataset, images that look similar to the naked eye tend to have a high similarity score, particularly under the perceptual hash algorithm. But as Tirilly et al. (2012) show, similarity is in the eye of the beholder – particularly at intermediate levels and when it is being assessed by laypersons. Figure B.1 illustrates the performance of the algorithms for three logos entered in the order shown by the same player in one contest (not necessarily from the sampled platform):

Figure B.1: Performance of image comparison algorithms

(1) (2) (3)

Notes: Figure shows three logos entered in order by a single player in a single contest. The perceptual hash algorithm calculates a similarity score of 0.313 for logos (1) and (2) and a score of 0.711 for (2) and (3). The difference hash algorithm calculates similarity scores of 0.508 for (1) and (2) and 0.891 for (2) and (3).

The first two images have several features in common but also have some notable differences. Each is centered, defined by a circular frame with text underneath, and presented against a similar backdrop. However the content of the circular frame and the font of the text below are considerably different, and the first logo is

in black and white while the second one is in color. The perceptual hash algorithm assigns these two logos a similarity score of 31 percent, while the difference hash gives them 51 percent.

In contrast, the second two images appear much more similar. They again have similar layouts, but now they share the same color assignments and the same content in the frame. Lesser differences remain, primarily with respect to the font style, but the logos appear broadly similar. The perceptual hash algorithm assigns these two logos a similarity score of 71 percent; the difference hash, 89 percent.

The algorithms thus pass the gut check in this example, which is not particularly unique: further examples using better-known brands are provided below. In light of this evidence, and the consistency of the paper's results, I believe that these algorithms provide empirically valid measures of experimentation.

Figure B.2: Volkswagen logo in 1937, 1967, 1995, 1999

(1937) (1967) (1995) (1999)

Notes: Figure shows the evolution of Volkswagen logos since 1937. The perceptual hash algorithm calculates similarity scores of 0.055 for the 1937 and 1967 logos, 0.430 for the 1967 and 1995 logos, and 0.844 for the 1995 and 1999 logos. The difference hash algorithm calculates similarity scores of 0.195, 0.539, and 0.953, respectively.

Figure B.3: Microsoft Windows 95, XP, 7, and 8 logos

(Windows 95) (Windows XP) (Windows 7) (Windows 8)

Notes: Figure shows a sequence of Windows logos. The perceptual hash algorithm calculates similarity scores of 0.195 for the Windows 95 and XP logos, 0.531 for the Windows XP and 7 logos, and 0.148 for the Windows 7 and 8 logos. The difference hash algorithm calculates similarity scores of 0.055, 0.563, and 0.117, respectively. The reason why the similarity of the Windows XP and 7 logos is not evaluated to be even higher is because the contrast generated by the latter's spotlight and shadow changes the structure of the image (for example, it changes the intensity gradient calculated by the difference hash algorithm).

Appendix References:

[1] Tirilly, Pierre, Chunsheng Huang, Wooseob Jeong, Xiangming Mu, Iris Xie, and Jin Zhang. 2012. "Image Similarity as Assessed by Users: A Quantitative Study." *Proceedings of the American Society for Information Science and Technology*, 49(1), pp. 1-10.

[2] Squire, David and Thierry Pun. 1997. "A Comparison of Human and Machine Assessments of Image Similarity for the Organization of Image Databases." *Proceedings of the Scandinavian Conference on Image Analysis*, Lappeenranta, Finland.

# C Contest Characteristics (evidence from Gross (2015) sample)

To highlight some of the basic relationships in the contests on this platform, I reproduce a subset of results from Gross (2015), which studies a larger sample from the same setting. Table C.1 estimates the relationship between contest characteristics such as the prize or frequency of feedback and key outcomes, and Table C.2 estimates the relationship between a design's rating and its probability of being selected.

**Correlation of contest characteristics and outcomes**

The estimates in Table C.1 suggest that an extra $100 in prize value on average attracts an additional 13.3 players, 47.7 designs, and 0.1 designs per player and increases the odds that a retractable prize will be awarded by 1.6 percent at the mean of all covariates. There is only a modest incremental effect of committed prize dollars, likely because the vast majority of uncommitted prizes are awarded anyway. The effects of feedback are also powerful: a sponsor who rates a high fraction of the designs in the contest will typically see fewer players enter but receive more designs from the participating players and have a much higher probability of finding a design it likes enough to award the prize. The effect of full feedback (relative to no feedback) on the probability the prize is awarded is greater than that of a $1000 increase in the prize – a more than quadrupling of the average and median prize in the sample.

Table C.1: Correlations of contest outcomes with their characteristics

| | (1) Players | (2) Designs | (3) Designs/Player | (4) Awarded |
|---|---|---|---|---|
| Total Prize Value ($100s) | 13.314*** | 47.695*** | 0.050*** | 0.101*** |
| | (0.713) | (2.930) | (0.016) | (0.027) |
| Committed Value ($100s) | 2.469** | 8.674* | 0.038 | |
| | (1.205) | (4.932) | (0.025) | |
| Fraction Rated | -7.321*** | 15.195*** | 1.026*** | 1.121*** |
| | (0.813) | (3.098) | (0.041) | (0.102) |
| Contest Length | 0.537*** | 2.109*** | 0.013*** | 0.021** |
| | (0.073) | (0.283) | (0.004) | (0.010) |
| Words in Desc. (100s) | 0.130 | 3.228*** | 0.063*** | -0.143*** |
| | (0.092) | (0.449) | (0.006) | (0.013) |
| Attached Materials | -0.943*** | -1.884*** | 0.048*** | -0.021 |
| | (0.173) | (0.692) | (0.013) | (0.015) |
| Prize Committed | 1.398 | 4.539 | -0.007 | |
| | (3.559) | (14.552) | (0.087) | |
| Constant | -2.445 | -63.730*** | 1.916*** | 1.085*** |
| | (2.309) | (9.045) | (0.072) | (0.155) |
| N | 4294 | 4294 | 4294 | 3298 |
| $R^2$ | 0.57 | 0.54 | 0.22 | |

Notes: Table shows the estimated effect of contest attributes on overall participation and the probability that the prize is awarded, using the larger sample of Gross (2015). The final specification is estimated as a probit on contests without a committed prize. *, **, *** represent significance at the 0.1, 0.05, and 0.01 levels, respectively. Monthly fixed effects included but not shown. Robust SEs in parentheses.

**Estimationg the success function: details**

Recall the following model from the text: let $R_{ijk}$ denote the rating on design $i$ by player $j$ in contest $k$, and (in a slight abuse of notation) let $R_{ijk} = \emptyset$ when design $ijk$ is unrated. The value of each design, $\nu_{ijk}$, can then be written as follows:

$$\nu_{ijk} = \gamma_{\emptyset}\mathbb{1}(R_{ijk} = \emptyset) + \gamma_1\mathbb{1}(R_{ijk} = 1) + \ldots + \gamma_5\mathbb{1}(R_{ijk} = 5) + \varepsilon_{ijk} \equiv \psi_{ijk} + \varepsilon_{ijk} \tag{4}$$

As in the theoretical model, the sponsor is assumed to select as winner the design with the highest value. In estimating the $\gamma$ parameters, each sponsor's choice set of designs is assumed to satisfy I.I.A.; in principle, the submission of a design of any rating in a given contest will reduce competing designs' chances of winning proportionally. For contests with an uncommitted prize, the choice set also includes an outside option of not awarding the prize, with value normalized to zero. Letting $I_{jk}$ be the set of designs by player $j$ in contest $k$, and $I_k$ be the set of all designs in contest $k$, player $jk$'s probability of winning is:

$$Pr(j \text{ wins } k) = \frac{\sum_{i \in I_{jk}} e^{\psi_{ijk}}}{\sum_{i \in I_k} e^{\psi_{ik}} + \mathbb{1}(\text{Uncommitted prize})}$$

I use the sample of 496,401 designs in 4,294 contests from Gross (2015) to estimate this model by maximum likelihood. The results are reproduced in Table C.2.

Table C.2: Conditional logit of win-lose outcomes on ratings

| Model: Latent design value $\nu_{ijk} = \gamma_5 + \gamma_4 + \gamma_3 + \gamma_2 + \gamma_1 + \gamma_{\emptyset} + \varepsilon_{ijk}$ | | | | |
|---|---|---|---|---|
| **Fixed effect** | **Est.** | **S.E.** | **t-stat** | **Implied $\beta$ (eqs. 1-2)** |
| Rating==5 | 1.53 | 0.07 | 22.17 | 4.618 |
| Rating==4 | -0.96 | 0.06 | -15.35 | 0.383 |
| Rating==3 | -3.39 | 0.08 | -40.01 | 0.034 |
| Rating==2 | -5.20 | 0.17 | -30.16 | 0.006 |
| Rating==1 | -6.02 | 0.28 | -21.82 | 0.002 |
| No rating | -3.43 | 0.06 | -55.35 | 0.032 |

Notes: Table provides results from a conditional logit estimation of the win-lose outcome of each design as a function of its rating, using the larger sample of Gross (2015). Outside option is not awarding the prize, with utility normalized to zero. The results can be used to approximate a player's probability of winning a contest as a function of her ratings. As a measure of fit, the design predicted by this model as the odds-on favorite wins roughly 50 percent of contests. See Gross (2015) for further discussion.

Table C.3 below sheds more light on the source of the conditional logit estimates, which are difficult to interpret directly. The table shows a cross-tabulation of contests, by the highest rating granted (columns) and the rating of the winning design (rows). The table shows that sponsors typically select the highest-rated design as winner, especially when the highest rating is 4- or 5-stars, but sponsors also often select unrated designs or the outside option. Rarely are 1- or 2-star designs ever awarded.

Table C.3: Frequency of contests, by highest rating and winning rating

| *Rating of winner* | *Highest rating in contest* | | | | | | |
|---|---|---|---|---|---|---|---|
| | **Unrated** | **1-star** | **2-star** | **3-star** | **4-star** | **5-star** | **Total** |
| **Not awarded** | 66 | 4 | 12 | 92 | 202 | 85 | 461 |
| **Unrated** | 142 | 5 | 10 | 59 | 347 | 276 | 839 |
| **1-star** | . | . | . | 3 | 6 | 5 | 14 |
| **2-star** | . | . | 3 | 11 | 16 | 8 | 38 |
| **3-star** | . | . | . | 43 | 146 | 53 | 242 |
| **4-star** | . | . | . | . | 836 | 379 | 1,215 |
| **5-star** | . | . | . | . | . | 1,485 | 1,485 |
| **Total** | 208 | 9 | 25 | 208 | 1,553 | 2,291 | 4,294 |

Notes: Table shows the frequency of contests in the Gross (2015) sample by the highest rating granted and the rating of the winning design.

## Evidence that the samples are comparable

The dataset in Gross (2015) consists of nearly all logo design contests with open bidding completed on the platform between July 1, 2010 and June 30, 2012, excluding those with zero prizes, multiple prizes, mid-contest rule changes, or otherwise unusual behavior, and it includes nearly all of the same information as the sample in this paper – except for the designs themselves. Although this sample comes from a slightly earlier time period than the one in the present paper (which was collected in the fall of 2013), both cover periods well after the platform was created and its growth had begun to stabilize.

Table C.4 compares characteristics of contests in the two samples. The contests in the Gross (2015) sample period are on average slightly longer, offer larger prizes, and attract a bit more participation relative to the sample of the present paper, but otherwise, the two samples are similar on observables. These differences are mostly due to the presence of a handful of outlying large contests in the Gross (2015) data. Interestingly, although the total number of designs is on average higher in the Gross (2015) sample, the number of designs of each rating is on average the same; the difference in total designs is fully accounted for by an increase in unrated entries. The most notable difference between the two samples is in the fraction of contests with a committed prize (23 percent vs. 56 percent). This discrepancy is explained by the fact that prize commitment only became an option on the platform halfway through the Gross (2015) sample period. Interestingly, the fraction of contests awarded is nevertheless nearly the same in these two samples.

Table C.4: Comparing Samples: Contest characteristics

| | **Gross (2015)** | **This paper** |
|---|---|---|
| ***Sample size*** | *4,294* | *122* |
| Contest length (days) | 9.15 | 8.52 |
| Prize value (US$) | 295.22 | 247.57 |
| No. of players | 37.28 | 33.20 |
| No. of designs | 115.52 | 96.38 |
| 5-star designs | 3.41 | 2.59 |
| 4-star designs | 13.84 | 12.28 |
| 3-star designs | 22.16 | 22.16 |
| 2-star designs | 16.04 | 17.61 |
| 1-star designs | 10.94 | 12.11 |
| Unrated designs | 49.14 | 29.62 |
| Number rated | 66.38 | 66.75 |
| Fraction rated | 0.56 | 0.64 |
| Prize committed | 0.23 | 0.56 |
| Prize awarded | 0.89 | 0.85 |

Tables C.5 and C.6 compare the distribution of ratings and batches in the two samples. The tables demonstrate that individual behavior is consistent across samples: sponsors assign each rating, and players enter designs, at roughly the same frequency. The main differences between the two samples are thus isolated to a handful of the overall contest characteristics highlighted in the previous table.

Table C.5: Comparing Samples: Distribution of ratings

| | **Gross (2015)** | **This paper** |
|---|---|---|
| ***Sample size*** | *285,052* | *8,144* |
| 1 star (in percent) | 16.48 | 18.15 |
| 2 stars | 24.16 | 26.39 |
| 3 stars | 33.38 | 33.19 |
| 4 stars | 20.84 | 18.39 |
| 5 stars | 5.13 | 3.88 |
| | 100.00 | 100.00 |

Table C.6: Comparing Samples: Design batches by size of batch

| | **Gross (2015)** | **This paper** |
|---|---|---|
| ***Sample size*** | *335,016* | *8,072* |
| 1 design (in percent) | 72.46 | 71.84 |
| 2 designs | 17.04 | 18.62 |
| 3 designs | 5.75 | 5.57 |
| 4 designs | 2.50 | 2.19 |
| 5+ designs | 2.25 | 1.77 |
| | 100.00 | 100.00 |

# D Additional Evidence: Identification

## D.1 Timing of feedback unrelated to rating granted

This subsection provides further support for the empirical strategy of estimating the effects of information about relative performance on individual choices. The paper shows that players do not behave in any ways consistent with their being able to forecast forthcoming feedback: behavior is uncorrelated with forthcoming (future) ratings on their previous entries. But we can also show that the timing of feedback is itself difficult to predict, and that it offers no information on performance.

To do so, in Table D.1 I take the subsample of all rated designs and regress the lag between the time of submission and feedback on indicators for the rating granted. The dependent variable in Column (1) is this lag, in hours; Column (2), the lag as a percent of contest duration; and Column (3), an indicator for whether the design was rated before the contest ended. All specifications account for contest and player fixed effects, include the standard controls used in the paper, and cluster standard errors by contest. Across all specifications, I find no relationship between feedback lags and ratings.

Table D.1: Correlation of feedback lags with rating granted

| | (1) | (2) | (3) |
|---|---|---|---|
| | Lag (hours) | Lag (pct. of contest) | Rated before end? |
| Rating==5 | 1.473 | 0.007 | -0.022 |
| | (3.461) | (0.016) | (0.031) |
| Rating==4 | -2.453 | -0.013* | 0.010 |
| | (1.792) | (0.008) | (0.020) |
| Rating==3 | -0.759 | -0.004 | 0.007 |
| | (2.131) | (0.009) | (0.015) |
| Rating==2 | 1.130 | 0.005 | 0.006 |
| | (1.922) | (0.008) | (0.011) |
| Constant | 9.958 | 0.184** | 1.276*** |
| | (14.184) | (0.075) | (0.108) |
| N | 7388 | 7388 | 8144 |
| $R^2$ | 0.45 | 0.48 | 0.63 |
| Controls | Yes | Yes | Yes |
| Contest FEs | Yes | Yes | Yes |
| Player FEs | Yes | Yes | Yes |

Notes: Table illustrates tendency for designs of different ratings to be rated more or less quickly. The results suggest that sponsors are not quicker to rate their favorite designs. Dependent variable in Column (1) is the lag between submission and feedback, in hours; Column (2), this lag as a fraction of contest length; and Column (3), an indicator for whether a design receives feedback before the contest ends. All columns control for time of entry, the number of previous designs entered by the given player and competitors, and contest and player fixed effects. *, **, *** represent significance at the 0.1, 0.05, and 0.01 levels, respectively. SEs clustered by contest in parentheses.

## D.2 Written comments are rare

A distinct threat to identification arises if ratings are accompanied by unobserved, written feedback, and these comments provide explicit instruction that generates the patterns found in this paper.

To evaluate this possibility, I draw a new sample of contests from this platform in which written feedback was visible to the public, seemingly by glitch/error. This sample consists of contests from early in the platform's history (significantly preceding the sample period for this paper), but it can nevertheless shed light on the frequency and potential confounding effects of written comments.

Within this sample, sponsors provided written comments to fewer than 8 percent of submissions, though as Table D.2 shows, the frequency is substantially higher for designs rated 4- or 5-stars (20 percent) than for those with poor ratings. Comments take a range of flavors, with many echoing the rating given (e.g., "This is on the right track" or "Not what I'm looking for"), but some make a more explicit request or suggestion of content changes. Because the latter present the risk of a confound, I had individuals read every comment in this sample and determine whether the sponsor suggested specific changes.

Table D.2: Typical distribution of ratings on designs receiving comments

| | **All designs** | **w/ Comments** | | **w/ Instructive comments** | | |
|---|---|---|---|---|---|---|
| **Rating** | **in sample (1)** | **Number (2)** | **% of all (2)/(1)** | **Number (3)** | **% of all (3)/(1)** | **% of comm. (3)/(2)** |
| 5 stars | 719 | 141 | 19.61 | 90 | 12.52 | 63.83 |
| 4 stars | 3,045 | 551 | 18.10 | 401 | 13.17 | 72.78 |
| 3 stars | 5,334 | 762 | 14.29 | 553 | 10.37 | 72.57 |
| 2 stars | 5,205 | 519 | 9.97 | 369 | 7.09 | 71.10 |
| 1 star | 5,114 | 498 | 9.74 | 262 | 5.12 | 52.61 |
| Unrated | 25,753 | 1,066 | 4.14 | 665 | 2.58 | 62.38 |
| **Total** | **45,170** | **3,537** | **7.83** | **2,340** | **5.18** | **66.16** |

Notes: Table tabulates rated designs by rating, within the sample of designs for which written feedback was accessible. The sample was compiled under a separate data collection effort and comprises contests from early in the platform's history (significantly preceding the sample period for this paper), but it is nonetheless informative. The table provides the frequency of each rating within this sample as a whole, then among: (i) designs receiving written comments, and (ii) those receiving written comments deemed instructive. Though higher-rated designs are more likely to receive written comments, the risk of a confounding effect arises only if these comments instruct players to make specific changes. The table shows that instructive comments are disproportionately given to designs with middling ratings – where this commentary may prove most constructive. The net effect is that 5-star designs receive instructive comments at roughly the same rate as 4- and 3-star designs.

On average, roughly two-thirds of comments are instructive in this way. Among designs receiving any comment, the ones most likely to receive an instructive comment are those with middling ratings: ideas that are incomplete but can be improved. As a result, on net we find that 5-, 4-, and 3-star designs all receive instructive comments at comparable rates (10-13 percent, see Table D.2).

Table D.3 formalizes this result with a regression, which provides standard errors. Column (1) regresses an indicator for whether a design in this sample received written feedback on indicators for its rating. Column (2) replaces the dependent variable with an indicator for a design receiving an instructive comment; Column (3) repeats this model for the subsample of designs that received any comment. In all cases, standard errors are clustered by contest. With standard errors in hand, it can be seen that 5-, 4-, and 3-star designs receive instructive comments at statistically similar rates (Table D.3, Column 2).

Table D.3: Raw frequencies of written feedback, by rating

| | Commented on | Instructive | Instructive |
|---|---|---|---|
| | (1) | (2) | (3) |
| Rated 5-stars | 0.196*** | 0.125*** | 0.638*** |
| | (0.026) | (0.018) | (0.049) |
| Rated 4-stars | 0.181*** | 0.132*** | 0.728*** |
| | (0.014) | (0.011) | (0.025) |
| Rated 3-stars | 0.143*** | 0.104*** | 0.726*** |
| | (0.013) | (0.010) | (0.023) |
| Rated 2-stars | 0.100*** | 0.071*** | 0.711*** |
| | (0.013) | (0.010) | (0.024) |
| Rated 1-star | 0.097*** | 0.051*** | 0.526*** |
| | (0.019) | (0.010) | (0.029) |
| Unrated | 0.041*** | 0.026*** | 0.624*** |
| | (0.005) | (0.003) | (0.040) |
| N | 45170 | 45170 | 3537 |
| $R^2$ | 0.11 | 0.08 | 0.67 |
| Sample | All designs | All designs | Commented |

Notes: Table provides results from regressing an indicator for whether a design received any comment (Column 1) or received a constructive comment (Columns 2 and 3) on indicators for its rating. Column (3) restricts the sample to only designs that received any comment. The table reproduces the tabulations in Table D.2, but provides standard errors. *, **, *** represent significance at the 0.1, 0.05, and 0.01 levels, respectively. SEs clustered by contest in parentheses.

Given that only around 10 percent of 5-star designs receive instructive comments, and that 5-star designs receive instruction at similar rates to 4- and 3-star designs, these comments are unlikely to be generating the results of this paper, which are large and increase dramatically from 3- to 5-stars. As a final check, in Table D.4 I expand the specifications in the previous table to match the format of the main results in the paper, interacting with competition and adding fixed effects and controls. The patterns in the previous table persist, and – importantly – we see no evidence that 5-star designs receive comments of any kind at different rates in the presence versus absence of top-rated competition.

Table D.4: Adding measures of competition, interactions, and controls

| | Commented on | Instructive | Instructive |
|---|---|---|---|
| | (1) | (2) | (3) |
| Rated 5-stars | 0.148*** | 0.100*** | 0.088 |
| | (0.022) | (0.018) | (0.067) |
| * 1+ competing 5-stars | -0.013 | -0.004 | -0.203 |
| | (0.027) | (0.020) | (0.273) |
| * prize value ($100s) | 0.002 | -0.004 | -0.013 |
| | (0.008) | (0.006) | (0.069) |
| Rated 4-stars | 0.136*** | 0.108*** | 0.134*** |
| | (0.015) | (0.011) | (0.047) |
| Rated 3-stars | 0.092*** | 0.077*** | 0.133*** |
| | (0.015) | (0.012) | (0.041) |
| Rated 2-stars | 0.055*** | 0.047*** | 0.114** |
| | (0.016) | (0.012) | (0.046) |
| Rated 1-star | 0.045** | 0.020** | -0.066 |
| | (0.020) | (0.010) | (0.059) |
| One or more competing 5-stars | 0.014 | 0.008 | -0.039 |
| | (0.016) | (0.011) | (0.067) |
| Pct. of contest elapsed | -0.179*** | -0.130*** | -0.148** |
| | (0.014) | (0.011) | (0.061) |
| N | 45170 | 45170 | 3537 |
| $R^2$ | 0.26 | 0.20 | 0.45 |
| Controls | Yes | Yes | Yes |
| Contest FEs | Yes | Yes | Yes |
| Player FEs | Yes | Yes | Yes |
| Sample | All designs | All designs | Commented |

Notes: Table provides results from regressing an indicator for whether a design received any comment (Column 1) or received a constructive comment (Columns 2 and 3) on indicators for its rating as well as an interaction for the presence of top-rated competition, as in the paper. Column (3) restricts the sample to only designs that received any comment. All columns include contest and player fixed effects and the standard set of controls. *, **, *** represent significance at the 0.1, 0.05, and 0.01 levels, respectively. SEs clustered by contest in parentheses.

# E Robustness Checks (1)

The following tables provide variants of the tables in Section 3 estimating the effects of feedback and competition on experimentation, using the difference hash algorithm. These estimates serve as robustness checks to the principal empirical results of the paper, demonstrating that they are not sensitive to the procedure used to calculate similarity scores. Table E.1 is a robustness check on Table 4; Table E.2, on Table 5; Table E.3, on Table 6; and Table E.4, on Table 7. The results in these tables are quantitatively similar to those in the body of the paper, despite the use of a distinct similarity measure.

Table E.1: Similarity to any of player's previous designs (diff. hash)

| | (1) | (2) | (3) | (4) |
|---|---|---|---|---|
| Player's prior best rating==5 | 0.268*** | 0.253*** | 0.266*** | 0.248*** |
| | (0.086) | (0.087) | (0.086) | (0.085) |
| * 1+ competing 5-stars | -0.128** | -0.141** | -0.132** | -0.139** |
| | (0.058) | (0.056) | (0.059) | (0.056) |
| * prize value ($100s) | -0.037 | -0.047* | -0.036 | -0.049* |
| | (0.026) | (0.027) | (0.026) | (0.026) |
| Player's prior best rating==4 | 0.056*** | 0.025 | 0.054*** | 0.021 |
| | (0.020) | (0.021) | (0.020) | (0.021) |
| Player's prior best rating==3 | 0.027 | 0.011 | 0.025 | 0.007 |
| | (0.017) | (0.017) | (0.017) | (0.016) |
| Player's prior best rating==2 | -0.004 | -0.012 | -0.007 | -0.017 |
| | (0.020) | (0.020) | (0.020) | (0.020) |
| One or more competing 5-stars | -0.012 | -0.022 | -0.012 | -0.023 |
| | (0.023) | (0.024) | (0.023) | (0.024) |
| Pct. of contest elapsed | 0.041 | -0.018 | 0.039 | -0.020 |
| | (0.029) | (0.072) | (0.029) | (0.073) |
| Constant | 0.455*** | 0.454*** | 0.456*** | 0.449*** |
| | (0.141) | (0.160) | (0.141) | (0.161) |
| N | 5075 | 5075 | 5075 | 5075 |
| $R^2$ | 0.48 | 0.48 | 0.48 | 0.48 |
| Controls | No | Yes | No | Yes |
| Contest FEs | Yes | Yes | Yes | Yes |
| Player FEs | Yes | Yes | Yes | Yes |
| Forthcoming ratings | No | No | Yes | Yes |

Notes: Table shows the effects of feedback on originality. Observations are designs. Dependent variable is a continuous measure of a design's maximal similarity to previous entries in the same contest by the same player, taking values in [0,1], where a value of 1 indicates the design is identical to another. The mean value of this variable is 0.58 (s.d. 0.28). Columns (2) and (4) control for days remaining and number of previous designs by the player and her competitors. Columns (3) and (4) additionally control for the best *forthcoming* rating on the player's not-yet-rated designs. Similarity scores in this table are calculated using a difference hash algorithm. Preceding designs/ratings are defined to be those entered/provided at least 60 minutes prior to the given design. *, **, *** represent significance at the 0.1, 0.05, and 0.01 levels, respectively. Standard errors clustered by player in parentheses.

Table E.2: Similarity to player's best previously-rated designs & intra-batch similarity (diff. hash)

| | Designs | | Batches (uwtd.) | | Batches (wtd.) | |
|---|---|---|---|---|---|---|
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Player's prior best rating==5 | 0.246* | 0.242* | 0.225 | 0.246 | 0.235 | 0.260 |
| | (0.133) | (0.141) | (0.299) | (0.293) | (0.286) | (0.281) |
| * 1+ competing 5-stars | -0.169** | -0.177** | -0.328** | -0.324** | -0.313** | -0.308** |
| | (0.086) | (0.087) | (0.146) | (0.144) | (0.147) | (0.145) |
| * prize value ($100s) | -0.018 | -0.024 | -0.022 | -0.023 | -0.025 | -0.026 |
| | (0.038) | (0.042) | (0.093) | (0.092) | (0.087) | (0.085) |
| Player's prior best rating==4 | 0.067* | 0.049 | -0.015 | -0.003 | -0.013 | 0.004 |
| | (0.039) | (0.041) | (0.031) | (0.032) | (0.030) | (0.031) |
| Player's prior best rating==3 | 0.044 | 0.033 | 0.011 | 0.019 | 0.010 | 0.020 |
| | (0.038) | (0.039) | (0.034) | (0.035) | (0.031) | (0.032) |
| Player's prior best rating==2 | 0.014 | 0.007 | -0.019 | -0.012 | -0.022 | -0.014 |
| | (0.040) | (0.040) | (0.047) | (0.049) | (0.045) | (0.045) |
| One or more competing 5-stars | -0.010 | -0.019 | -0.017 | -0.018 | -0.015 | -0.017 |
| | (0.031) | (0.032) | (0.033) | (0.034) | (0.032) | (0.033) |
| Pct. of contest elapsed | -0.021 | -0.057 | -0.001 | -0.024 | 0.003 | 0.007 |
| | (0.038) | (0.104) | (0.039) | (0.085) | (0.038) | (0.079) |
| Constant | 0.862*** | 0.863*** | 0.643*** | 0.673*** | 0.670*** | 0.661*** |
| | (0.152) | (0.173) | (0.121) | (0.156) | (0.099) | (0.128) |
| N | 3871 | 3871 | 1987 | 1987 | 1987 | 1987 |
| $R^2$ | 0.53 | 0.53 | 0.59 | 0.59 | 0.59 | 0.59 |
| Controls | No | Yes | No | Yes | No | Yes |
| Contest FEs | Yes | Yes | Yes | Yes | Yes | Yes |
| Player FEs | Yes | Yes | Yes | Yes | Yes | Yes |

Notes: Table shows the effects of feedback on originality. Observations in Columns (1) and (2) are designs, and dependent variable is a continuous measure of a design's similarity to the highest-rated preceding entry by the same player, taking values in [0,1], where a value of 1 indicates the design is identical to another. The mean value of this variable is 0.52 (s.d. 0.30). Observations in Columns (3) to (6) are design batches, which are defined to be a set of designs by a single player entered into a contest in close proximity (15 minutes), and dependent variable is a continuous measure of intra-batch similarity, taking values in [0,1], where a value of 1 indicates that two designs in the batch are identical. The mean value of this variable is 0.72 (s.d. 0.27). Columns (5) and (6) weight the batch regressions by batch size. All columns control for the time of submission and number of previous designs by the player and her competitors. Similarity scores in this table are calculated using a difference hash algorithm. Preceding designs/ratings are defined to be those entered/provided at least 60 minutes prior to the given design. *, **, *** represent significance at the 0.1, 0.05, and 0.01 levels, respectively. Standard errors clustered by player in parentheses.

Table E.3: Change in similarity to player's best previously-rated designs (diff. hash)

| | (1) | (2) | (3) | (4) | (5) | (6) |
|---|---|---|---|---|---|---|
| Δ(Player's best rating==5) | 0.658*** | 0.679*** | 0.688** | 0.659*** | 0.693*** | 0.693*** |
| | (0.218) | (0.256) | (0.268) | (0.217) | (0.256) | (0.267) |
| * 1+ competing 5-stars | -0.349** | -0.375* | -0.364* | -0.350** | -0.379* | -0.368* |
| | (0.175) | (0.207) | (0.218) | (0.174) | (0.206) | (0.218) |
| * prize value ($100s) | -0.048 | -0.059 | -0.062 | -0.048 | -0.063 | -0.064 |
| | (0.046) | (0.055) | (0.057) | (0.046) | (0.055) | (0.057) |
| Δ(Player's best rating==4) | 0.262*** | 0.236*** | 0.232*** | 0.262*** | 0.237*** | 0.232*** |
| | (0.070) | (0.081) | (0.086) | (0.070) | (0.081) | (0.086) |
| Δ(Player's best rating==3) | 0.192*** | 0.170** | 0.162** | 0.192*** | 0.169** | 0.162** |
| | (0.062) | (0.073) | (0.077) | (0.063) | (0.073) | (0.077) |
| Δ(Player's best rating==2) | 0.133** | 0.110 | 0.104 | 0.131** | 0.110 | 0.104 |
| | (0.058) | (0.068) | (0.071) | (0.058) | (0.067) | (0.071) |
| One or more competing 5-stars | -0.002 | 0.002 | -0.001 | 0.001 | 0.000 | -0.001 |
| | (0.017) | (0.017) | (0.028) | (0.018) | (0.016) | (0.029) |
| Pct. of contest elapsed | -0.007 | -0.011 | -0.014 | -0.067 | -0.073 | -0.104 |
| | (0.020) | (0.029) | (0.036) | (0.062) | (0.064) | (0.110) |
| Constant | -0.009 | -0.008 | -0.238*** | 0.058 | 0.038 | -0.169** |
| | (0.010) | (0.017) | (0.036) | (0.050) | (0.047) | (0.072) |
| N | 2694 | 2694 | 2694 | 2694 | 2694 | 2694 |
| $R^2$ | 0.04 | 0.10 | 0.13 | 0.04 | 0.10 | 0.13 |
| Controls | No | No | No | Yes | Yes | Yes |
| Contest FEs | Yes | No | Yes | Yes | No | Yes |
| Player FEs | No | Yes | Yes | No | Yes | Yes |

Notes: Table shows the effects of feedback on originality. Observations are designs. Dependent variable is a continuous measure of the *change* in designs' similarity to the highest-rated preceding entry by the same player, taking values in [-1,1], where a value of 0 indicates that the player's current design is as similar to her best preceding design as was her previous design, and a value of 1 indicates that the player transitioned fully from experimenting to copying (and a value of -1, the converse). The mean value of this variable is -0.01 (s.d. 0.25). Columns (4) to (6) control for days remaining and number of previous designs by the player and competitors. Similarity scores in this table are calculated using a difference hash algorithm. Preceding designs/ratings are defined to be those entered/provided at least 60 minutes prior to the given design. *, **, *** represent significance at the 0.1, 0.05, and 0.01 levels, respectively. Standard errors clustered by player in parentheses.

Table E.4: Similarity to player's best not-yet-rated designs (placebo test; using diff. hash)

| | Similarity to forthcoming | | | Residual |
|---|---|---|---|---|
| | (1) | (2) | (3) | (4) |
| Player's best forthcoming rating==5 | 0.203 | 0.069 | 0.022 | 0.060 |
| | (0.241) | (0.116) | (0.127) | (0.119) |
| * 1+ competing 5-stars | -0.040 | -0.024 | 0.000 | -0.023 |
| | (0.145) | (0.073) | (0.080) | (0.079) |
| * prize value ($100s) | -0.064 | -0.009 | -0.001 | -0.006 |
| | (0.042) | (0.028) | (0.032) | (0.029) |
| Player's best forthcoming rating==4 | 0.023 | 0.051 | 0.059 | 0.045 |
| | (0.074) | (0.066) | (0.064) | (0.067) |
| Player's best forthcoming rating==3 | 0.043 | 0.069 | 0.069 | 0.055 |
| | (0.050) | (0.055) | (0.055) | (0.053) |
| Player's best forthcoming rating==2 | 0.031 | 0.025 | 0.026 | 0.025 |
| | (0.048) | (0.051) | (0.051) | (0.049) |
| One or more competing 5-stars | -0.077 | -0.089 | -0.087 | -0.085 |
| | (0.076) | (0.115) | (0.119) | (0.124) |
| Pct. of contest elapsed | -0.210 | 0.033 | 0.093 | 0.023 |
| | (0.261) | (0.442) | (0.422) | (0.469) |
| Constant | 0.826** | 0.491 | 0.515 | 0.060 |
| | (0.357) | (0.412) | (0.472) | (0.507) |
| N | 1147 | 577 | 577 | 577 |
| $R^2$ | 0.69 | 0.87 | 0.88 | 0.69 |
| Controls | Yes | Yes | Yes | Yes |
| Contest FEs | Yes | Yes | Yes | Yes |
| Player FEs | Yes | Yes | Yes | Yes |

Notes: Table provides a placebo test of the effects of future feedback on originality. Observations are designs. Dependent variable in Columns (1) to (3) is a continuous measure of a design's similarity to the best design that the player has previously entered that has yet to *but will eventually be* rated, taking values in [0,1], where a value of 1 indicates that the two designs are identical. The mean value of this variable is 0.50 (s.d. 0.29). Under the identifying assumption that future feedback is unpredictable, current choices should be unrelated to forthcoming ratings. Note that a given design's similarity to an earlier, unrated design can be incidental if they are both tweaks on a rated third design. To account for this possibility, Column (2) controls for the given and unrated designs' similarity to the best previously-rated design. Column (3) allows these controls to vary with the highest rating previously received. Dependent variable in Column (4) is the residual from a regression of the dependent variable in the previous columns on these controls. These residuals will be the subset of a given design's similarity to the unrated design that is not explained by jointly-occurring similarity to a third design. All columns control for days remaining and number of previous designs by the player and her competitors. Similarity scores in this table are calculated using a difference hash algorithm. Preceding designs/ratings are defined to be those entered/provided at least 60 minutes prior to the given design. *, **, *** represent significance at the 0.1, 0.05, and 0.01 levels, respectively. Standard errors clustered by player in parentheses.

# F Robustness Checks (2)

In additional robustness checks, I show that competition has a constant effect on high-performing players' tendency to experiment. Tables F.1 to F.3 demonstrate this result with the perceptual hash similarity measures, and Tables F.4 to F.6 do so with the difference hash measures. In all cases, I estimate differential effects for one vs. multiple top-rated, competing designs and find no differential effect.

Table F.1: Similarity to any of player's previous designs (p. hash)

| | (1) | (2) | (3) | (4) |
|---|---|---|---|---|
| Player's prior best rating==5 | 0.287*** | 0.280*** | 0.288*** | 0.278*** |
| | (0.086) | (0.088) | (0.085) | (0.088) |
| * 1+ competing 5-stars | -0.157* | -0.158* | -0.155* | -0.153* |
| | (0.093) | (0.089) | (0.094) | (0.089) |
| * 2+ competing 5-stars | 0.060 | 0.052 | 0.063 | 0.056 |
| | (0.111) | (0.098) | (0.110) | (0.096) |
| * prize value ($100s) | -0.012 | -0.022 | -0.014 | -0.026 |
| | (0.028) | (0.028) | (0.028) | (0.028) |
| Player's prior best rating==4 | 0.099*** | 0.076*** | 0.099*** | 0.074*** |
| | (0.017) | (0.017) | (0.017) | (0.018) |
| Player's prior best rating==3 | 0.039*** | 0.028* | 0.039*** | 0.026* |
| | (0.014) | (0.014) | (0.014) | (0.014) |
| Player's prior best rating==2 | -0.004 | -0.009 | -0.003 | -0.011 |
| | (0.020) | (0.020) | (0.020) | (0.020) |
| One or more competing 5-stars | 0.006 | 0.005 | 0.006 | 0.005 |
| | (0.032) | (0.032) | (0.032) | (0.032) |
| Two or more competing 5-stars | -0.024 | -0.035 | -0.024 | -0.035 |
| | (0.035) | (0.037) | (0.035) | (0.037) |
| Pct. of contest elapsed | 0.039 | -0.095 | 0.040 | -0.094 |
| | (0.028) | (0.073) | (0.028) | (0.073) |
| Constant | 0.292 | 0.408** | 0.290 | 0.400** |
| | (0.180) | (0.196) | (0.180) | (0.197) |
| N | 5075 | 5075 | 5075 | 5075 |
| $R^2$ | 0.47 | 0.47 | 0.47 | 0.47 |
| Controls | No | Yes | No | Yes |
| Contest FEs | Yes | Yes | Yes | Yes |
| Player FEs | Yes | Yes | Yes | Yes |
| Forthcoming ratings | No | No | Yes | Yes |

Notes: Table shows the effects of feedback on originality. Observations are designs. Dependent variable is a continuous measure of a design's maximal similarity to previous entries in the same contest by the same player, taking values in [0,1], where a value of 1 indicates the design is identical to another. The mean value of this variable is 0.32 (s.d. 0.27). Columns (2) and (4) control for days remaining and number of previous designs by the player and her competitors. Columns (3) and (4) additionally control for the best *forthcoming* rating on the player's not-yet-rated designs. Similarity scores in this table are calculated using a perceptual hash algorithm. Preceding designs/ratings are defined to be those entered/provided at least 60 minutes prior to the given design. *, **, *** represent significance at the 0.1, 0.05, and 0.01 levels, respectively. Standard errors clustered by player in parentheses.

Table F.2: Similarity to player's best previously-rated designs & intra-batch similarity (p. hash)

| | Designs | | Batches (uwtd.) | | Batches (wtd.) | |
|---|---|---|---|---|---|---|
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Player's prior best rating==5 | 0.358*** | 0.362*** | 0.213 | 0.240 | 0.245 | 0.285 |
| | (0.098) | (0.103) | (0.312) | (0.307) | (0.302) | (0.294) |
| * 1+ competing 5-stars | -0.225** | -0.226** | -0.437** | -0.433** | -0.475** | -0.469** |
| | (0.093) | (0.093) | (0.217) | (0.214) | (0.203) | (0.200) |
| * 2+ competing 5-stars | 0.025 | 0.023 | 0.180 | 0.177 | 0.237 | 0.238 |
| | (0.094) | (0.092) | (0.236) | (0.232) | (0.223) | (0.218) |
| * prize value ($100s) | -0.014 | -0.018 | 0.020 | 0.015 | 0.013 | 0.009 |
| | (0.031) | (0.033) | (0.099) | (0.099) | (0.095) | (0.093) |
| Player's prior best rating==4 | 0.122*** | 0.116*** | 0.055* | 0.066* | 0.065** | 0.086** |
| | (0.031) | (0.032) | (0.032) | (0.037) | (0.032) | (0.038) |
| Player's prior best rating==3 | 0.061** | 0.056** | 0.056 | 0.062* | 0.052 | 0.065* |
| | (0.028) | (0.028) | (0.035) | (0.037) | (0.035) | (0.037) |
| Player's prior best rating==2 | 0.026 | 0.024 | 0.023 | 0.029 | 0.009 | 0.020 |
| | (0.030) | (0.030) | (0.049) | (0.050) | (0.047) | (0.047) |
| One or more competing 5-stars | 0.002 | 0.002 | 0.083 | 0.086 | 0.076 | 0.078 |
| | (0.037) | (0.037) | (0.063) | (0.063) | (0.063) | (0.063) |
| Two or more competing 5-stars | 0.004 | 0.000 | -0.094 | -0.100 | -0.080 | -0.086 |
| | (0.042) | (0.044) | (0.072) | (0.076) | (0.074) | (0.076) |
| Pct. of contest elapsed | -0.019 | -0.103 | -0.016 | -0.102 | -0.005 | -0.064 |
| | (0.034) | (0.084) | (0.049) | (0.115) | (0.050) | (0.112) |
| Constant | 0.422** | 0.487*** | 0.392*** | 0.504*** | 0.385*** | 0.457*** |
| | (0.168) | (0.187) | (0.066) | (0.149) | (0.061) | (0.148) |
| N | 3871 | 3871 | 1987 | 1987 | 1987 | 1987 |
| $R^2$ | 0.53 | 0.53 | 0.58 | 0.58 | 0.58 | 0.58 |
| Controls | No | Yes | No | Yes | No | Yes |
| Contest FEs | Yes | Yes | Yes | Yes | Yes | Yes |
| Player FEs | Yes | Yes | Yes | Yes | Yes | Yes |

Notes: Table shows the effects of feedback on originality. Observations in Columns (1) and (2) are designs, and dependent variable is a continuous measure of a design's similarity to the highest-rated preceding entry by the same player, taking values in [0,1], where a value of 1 indicates the design is identical to another. The mean value of this variable is 0.28 (s.d. 0.27). Observations in Columns (3) to (6) are design batches, which are defined to be a set of designs by a single player entered into a contest in close proximity (15 minutes), and dependent variable is a continuous measure of intra-batch similarity, taking values in [0,1], where a value of 1 indicates that two designs in the batch are identical. The mean value of this variable is 0.48 (s.d. 0.32). Columns (5) and (6) weight the batch regressions by batch size. All columns control for the time of submission and number of previous designs by the player and her competitors. Similarity scores in this table are calculated using a perceptual hash algorithm. Preceding designs/ratings are defined to be those entered/provided at least 60 minutes prior to the given design. *, **, *** represent significance at the 0.1, 0.05, and 0.01 levels, respectively. Standard errors clustered by player in parentheses.

Table F.3: Change in similarity to player's best previously-rated designs (p. hash)

| | (1) | (2) | (3) | (4) | (5) | (6) |
|---|---|---|---|---|---|---|
| Δ(Player's best rating==5) | 0.879*** | 0.934*** | 0.921*** | 0.886*** | 0.934*** | 0.928*** |
| | (0.166) | (0.194) | (0.200) | (0.167) | (0.193) | (0.200) |
| * 1+ competing 5-stars | -0.487*** | -0.534*** | -0.504*** | -0.489*** | -0.532*** | -0.505*** |
| | (0.133) | (0.150) | (0.160) | (0.134) | (0.150) | (0.161) |
| * 2+ competing 5-stars | 0.099 | 0.162 | 0.108 | 0.099 | 0.158 | 0.107 |
| | (0.111) | (0.123) | (0.127) | (0.112) | (0.123) | (0.127) |
| * prize value ($100s) | -0.093** | -0.115*** | -0.108** | -0.095** | -0.114** | -0.110** |
| | (0.037) | (0.044) | (0.044) | (0.038) | (0.044) | (0.045) |
| Δ(Player's best rating==4) | 0.285*** | 0.274*** | 0.281*** | 0.287*** | 0.276*** | 0.284*** |
| | (0.065) | (0.071) | (0.077) | (0.065) | (0.071) | (0.077) |
| Δ(Player's best rating==3) | 0.154*** | 0.141** | 0.141** | 0.154*** | 0.142** | 0.142** |
| | (0.057) | (0.063) | (0.067) | (0.057) | (0.063) | (0.067) |
| Δ(Player's best rating==2) | 0.085* | 0.069 | 0.064 | 0.084* | 0.069 | 0.063 |
| | (0.046) | (0.052) | (0.056) | (0.046) | (0.052) | (0.056) |
| One or more competing 5-stars | 0.002 | -0.018 | -0.014 | 0.003 | -0.019 | -0.013 |
| | (0.026) | (0.031) | (0.044) | (0.026) | (0.032) | (0.043) |
| Two or more competing 5-stars | -0.006 | 0.021 | 0.031 | -0.006 | 0.022 | 0.029 |
| | (0.025) | (0.032) | (0.041) | (0.026) | (0.034) | (0.043) |
| Pct. of contest elapsed | 0.009 | 0.014 | -0.003 | -0.036 | 0.008 | -0.050 |
| | (0.018) | (0.023) | (0.029) | (0.043) | (0.048) | (0.074) |
| Constant | -0.017* | -0.021 | 0.029 | 0.024 | -0.014 | 0.058 |
| | (0.010) | (0.014) | (0.061) | (0.040) | (0.037) | (0.078) |
| N | 2694 | 2694 | 2694 | 2694 | 2694 | 2694 |
| $R^2$ | 0.05 | 0.11 | 0.14 | 0.05 | 0.11 | 0.14 |
| Controls | No | No | No | Yes | Yes | Yes |
| Contest FEs | Yes | No | Yes | Yes | No | Yes |
| Player FEs | No | Yes | Yes | No | Yes | Yes |

Notes: Table shows the effects of feedback on originality. Observations are designs. Dependent variable is a continuous measure of the *change* in designs' similarity to the highest-rated preceding entry by the same player, taking values in [-1,1], where a value of 0 indicates that the player's current design is as similar to her best preceding design as was her previous design, and a value of 1 indicates that the player transitioned fully from experimenting to copying (and a value of -1, the converse). The mean value of this variable is 0.00 (s.d. 0.23). Columns (4) to (6) control for days remaining and number of previous designs by the player and competitors. Similarity scores in this table are calculated using a perceptual hash algorithm. Preceding designs/ratings are defined to be those entered/provided at least 60 minutes prior to the given design. *, **, *** represent significance at the 0.1, 0.05, and 0.01 levels, respectively. Standard errors clustered by player in parentheses.

Table F.4: Similarity to any of player's previous designs (d. hash)

| | (1) | (2) | (3) | (4) |
|---|---|---|---|---|
| Player's prior best rating==5 | 0.267*** | 0.250*** | 0.266*** | 0.245*** |
| | (0.087) | (0.087) | (0.086) | (0.085) |
| * 1+ competing 5-stars | -0.103 | -0.110 | -0.105 | -0.108 |
| | (0.086) | (0.080) | (0.088) | (0.080) |
| * 2+ competing 5-stars | -0.033 | -0.039 | -0.036 | -0.040 |
| | (0.094) | (0.080) | (0.095) | (0.080) |
| * prize value ($100s) | -0.036 | -0.046* | -0.036 | -0.048* |
| | (0.026) | (0.027) | (0.026) | (0.026) |
| Player's prior best rating==4 | 0.058*** | 0.027 | 0.056*** | 0.023 |
| | (0.020) | (0.021) | (0.020) | (0.021) |
| Player's prior best rating==3 | 0.028* | 0.013 | 0.026 | 0.008 |
| | (0.017) | (0.017) | (0.017) | (0.017) |
| Player's prior best rating==2 | -0.004 | -0.012 | -0.007 | -0.017 |
| | (0.020) | (0.020) | (0.020) | (0.020) |
| One or more competing 5-stars | -0.043 | -0.044 | -0.043 | -0.045 |
| | (0.037) | (0.038) | (0.037) | (0.038) |
| Two or more competing 5-stars | 0.049 | 0.036 | 0.048 | 0.035 |
| | (0.037) | (0.040) | (0.037) | (0.040) |
| Pct. of contest elapsed | 0.035 | -0.020 | 0.034 | -0.022 |
| | (0.028) | (0.073) | (0.028) | (0.073) |
| Constant | 0.461*** | 0.460*** | 0.461*** | 0.454*** |
| | (0.141) | (0.160) | (0.141) | (0.161) |
| N | 5075 | 5075 | 5075 | 5075 |
| $R^2$ | 0.48 | 0.48 | 0.48 | 0.48 |
| Controls | No | Yes | No | Yes |
| Contest FEs | Yes | Yes | Yes | Yes |
| Player FEs | Yes | Yes | Yes | Yes |
| Forthcoming ratings | No | No | Yes | Yes |

Notes: Table shows the effects of feedback on originality. Observations are designs. Dependent variable is a continuous measure of a design's maximal similarity to previous entries in the same contest by the same player, taking values in [0,1], where a value of 1 indicates the design is identical to another. The mean value of this variable is 0.58 (s.d. 0.28). Columns (2) and (4) control for days remaining and number of previous designs by the player and her competitors. Columns (3) and (4) additionally control for the best *forthcoming* rating on the player's not-yet-rated designs. Similarity scores in this table are calculated using a difference hash algorithm. Preceding designs/ratings are defined to be those entered/provided at least 60 minutes prior to the given design. *, **, *** represent significance at the 0.1, 0.05, and 0.01 levels, respectively. Standard errors clustered by player in parentheses.

Table F.5: Similarity to player's best previously-rated designs & intra-batch similarity (d. hash)

| | Designs | | Batches (uwtd.) | | Batches (wtd.) | |
|---|---|---|---|---|---|---|
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Player's prior best rating==5 | 0.242* | 0.234* | 0.239 | 0.271 | 0.252 | 0.285 |
| | (0.134) | (0.141) | (0.317) | (0.312) | (0.300) | (0.295) |
| * 1+ competing 5-stars | -0.144 | -0.149 | -0.329** | -0.328** | -0.328** | -0.327** |
| | (0.140) | (0.138) | (0.144) | (0.142) | (0.132) | (0.131) |
| * 2+ competing 5-stars | -0.033 | -0.036 | 0.007 | 0.009 | 0.025 | 0.030 |
| | (0.143) | (0.134) | (0.189) | (0.186) | (0.173) | (0.172) |
| * prize value ($100s) | -0.017 | -0.022 | -0.026 | -0.031 | -0.030 | -0.033 |
| | (0.039) | (0.042) | (0.100) | (0.099) | (0.092) | (0.090) |
| Player's prior best rating==4 | 0.068* | 0.050 | -0.015 | -0.002 | -0.012 | 0.004 |
| | (0.039) | (0.041) | (0.031) | (0.032) | (0.030) | (0.031) |
| Player's prior best rating==3 | 0.044 | 0.033 | 0.012 | 0.019 | 0.011 | 0.020 |
| | (0.038) | (0.039) | (0.034) | (0.035) | (0.031) | (0.032) |
| Player's prior best rating==2 | 0.013 | 0.006 | -0.018 | -0.011 | -0.020 | -0.012 |
| | (0.040) | (0.040) | (0.047) | (0.048) | (0.044) | (0.045) |
| One or more competing 5-stars | -0.048 | -0.050 | 0.037 | 0.038 | 0.030 | 0.030 |
| | (0.045) | (0.045) | (0.042) | (0.042) | (0.038) | (0.038) |
| Two or more competing 5-stars | 0.059 | 0.052 | -0.089* | -0.096* | -0.076 | -0.082 |
| | (0.049) | (0.051) | (0.051) | (0.055) | (0.048) | (0.051) |
| Pct. of contest elapsed | -0.030 | -0.062 | 0.006 | -0.037 | 0.008 | -0.007 |
| | (0.037) | (0.105) | (0.040) | (0.085) | (0.039) | (0.080) |
| Constant | 0.866*** | 0.870*** | 0.636*** | 0.677*** | 0.665*** | 0.666*** |
| | (0.152) | (0.174) | (0.120) | (0.156) | (0.098) | (0.128) |
| N | 3871 | 3871 | 1987 | 1987 | 1987 | 1987 |
| $R^2$ | 0.53 | 0.54 | 0.59 | 0.59 | 0.59 | 0.59 |
| Controls | No | Yes | No | Yes | No | Yes |
| Contest FEs | Yes | Yes | Yes | Yes | Yes | Yes |
| Player FEs | Yes | Yes | Yes | Yes | Yes | Yes |

Notes: Table shows the effects of feedback on originality. Observations in Columns (1) and (2) are designs, and dependent variable is a continuous measure of a design's similarity to the highest-rated preceding entry by the same player, taking values in [0,1], where a value of 1 indicates the design is identical to another. The mean value of this variable is 0.52 (s.d. 0.30). Observations in Columns (3) to (6) are design batches, which are defined to be a set of designs by a single player entered into a contest in close proximity (15 minutes), and dependent variable is a continuous measure of intra-batch similarity, taking values in [0,1], where a value of 1 indicates that two designs in the batch are identical. The mean value of this variable is 0.72 (s.d. 0.27). Columns (5) and (6) weight the batch regressions by batch size. All columns control for the time of submission and number of previous designs by the player and her competitors. Similarity scores in this table are calculated using a difference hash algorithm. Preceding designs/ratings are defined to be those entered/provided at least 60 minutes prior to the given design. *, **, *** represent significance at the 0.1, 0.05, and 0.01 levels, respectively. Standard errors clustered by player in parentheses.

Table F.6: Change in similarity to player's best previously-rated designs (d. hash)

| | (1) | (2) | (3) | (4) | (5) | (6) |
|---|---|---|---|---|---|---|
| Δ(Player's best rating==5) | 0.657*** | 0.684*** | 0.695*** | 0.657*** | 0.695*** | 0.695*** |
| | (0.218) | (0.253) | (0.265) | (0.217) | (0.253) | (0.264) |
| * 1+ competing 5-stars | -0.401* | -0.452* | -0.428* | -0.400* | -0.456* | -0.429* |
| | (0.211) | (0.236) | (0.255) | (0.211) | (0.236) | (0.256) |
| * 2+ competing 5-stars | 0.070 | 0.108 | 0.090 | 0.068 | 0.109 | 0.088 |
| | (0.154) | (0.167) | (0.184) | (0.155) | (0.167) | (0.185) |
| * prize value ($100s) | -0.047 | -0.059 | -0.064 | -0.047 | -0.062 | -0.063 |
| | (0.046) | (0.053) | (0.055) | (0.045) | (0.053) | (0.055) |
| Δ(Player's best rating==4) | 0.264*** | 0.240*** | 0.236*** | 0.264*** | 0.241*** | 0.237*** |
| | (0.070) | (0.081) | (0.086) | (0.070) | (0.081) | (0.085) |
| Δ(Player's best rating==3) | 0.195*** | 0.174** | 0.166** | 0.194*** | 0.173** | 0.166** |
| | (0.063) | (0.073) | (0.077) | (0.063) | (0.074) | (0.077) |
| Δ(Player's best rating==2) | 0.135** | 0.115* | 0.108 | 0.133** | 0.114* | 0.108 |
| | (0.058) | (0.068) | (0.071) | (0.058) | (0.068) | (0.071) |
| One or more competing 5-stars | -0.011 | -0.018 | -0.028 | -0.008 | -0.018 | -0.026 |
| | (0.030) | (0.035) | (0.049) | (0.029) | (0.036) | (0.048) |
| Two or more competing 5-stars | 0.012 | 0.029 | 0.044 | 0.015 | 0.026 | 0.046 |
| | (0.027) | (0.037) | (0.046) | (0.028) | (0.038) | (0.048) |
| Pct. of contest elapsed | -0.009 | -0.013 | -0.023 | -0.067 | -0.072 | -0.107 |
| | (0.019) | (0.028) | (0.034) | (0.062) | (0.065) | (0.110) |
| Constant | -0.008 | -0.007 | -0.238*** | 0.059 | 0.040 | -0.171** |
| | (0.010) | (0.016) | (0.036) | (0.049) | (0.046) | (0.073) |
| N | 2694 | 2694 | 2694 | 2694 | 2694 | 2694 |
| $R^2$ | 0.04 | 0.10 | 0.13 | 0.04 | 0.10 | 0.13 |
| Controls | No | No | No | Yes | Yes | Yes |
| Contest FEs | Yes | No | Yes | Yes | No | Yes |
| Player FEs | No | Yes | Yes | No | Yes | Yes |

Notes: Table shows the effects of feedback on originality. Observations are designs. Dependent variable is a continuous measure of the *change* in designs' similarity to the highest-rated preceding entry by the same player, taking values in [-1,1], where a value of 0 indicates that the player's current design is as similar to her best preceding design as was her previous design, and a value of 1 indicates that the player transitioned fully from experimenting to copying (and a value of -1, the converse). The mean value of this variable is -0.01 (s.d. 0.25). Columns (4) to (6) control for days remaining and number of previous designs by the player and competitors. Similarity scores in this table are calculated using a difference hash algorithm. Preceding designs/ratings are defined to be those entered/provided at least 60 minutes prior to the given design. *, **, *** represent significance at the 0.1, 0.05, and 0.01 levels, respectively. Standard errors clustered by player in parentheses.

# G Robustness Checks (3)

The basic fact documented in this paper is the differential tendency for high-performing agents to try new ideas in the absence versus presence of other high-quality competition. Though the body of the paper focuses on variation in 5-star ratings, a natural question to ask is whether the same patterns arise for players with 4-star ratings, particularly when that is the highest rating in the contest. In other words: does heart of the matter have to do with having a 5-star rating, or having the best rating?

Table G.1 explores this question. The table runs regressions similar to those in Table 4, restricted to observations where there are not yet any 5-star designs in the same contest. The specifications regress the maximal similarity of the given design to previous entries by the same player on indicators for the player's highest rating, interacting the 4-star indicator with 4-star competition. Since the importance of the 4-star ratings grows towards the end of a contest, as it becomes clear that no higher ratings will be granted, the table breaks the results out for all submissions (Cols. 1 and 4), submissions in the 2nd half of a contest (Cols. 2 and 5), and submissions in the 4th quarter (Cols. 3 and 6). The latter three columns in the table add controls for a player's highest forthcoming rating, as in the body of the paper.

The evidence suggests the results are general to 4-on-4 competition. In Columns (1) to (3), players whose highest rating is 4-stars enter designs that are significantly more similar to their previous work than those with lower ratings, but these effects are reversed by 4-star competition, with both effects growing in magnitude over the course of a contest – though by the final quarter of a contest, reduced cell sizes cause the standard errors to rise, such that the estimates become marginally significant or not statistically different from zero. The results with controls, in the latter three columns, are identical in magnitude and statistical significance. Table G.2 repeats the exercise using the difference hash algorithm.

Table G.1: Similarity to any of player's previous designs: 4-vs-4 (p. hash)

| | (1) | (2) | (3) | (4) | (5) | (6) |
|---|---|---|---|---|---|---|
| Player's prior best rating==4 | 0.185*** | 0.382** | 0.564* | 0.184*** | 0.391** | 0.516* |
| | (0.069) | (0.190) | (0.299) | (0.069) | (0.196) | (0.283) |
| * 1+ competing 4-stars | -0.121** | -0.323* | -0.315 | -0.122** | -0.325* | -0.280 |
| | (0.051) | (0.181) | (0.248) | (0.051) | (0.185) | (0.226) |
| * prize value ($100s) | -0.011 | -0.006 | -0.045 | -0.012 | -0.010 | -0.039 |
| | (0.020) | (0.033) | (0.097) | (0.020) | (0.033) | (0.092) |
| Player's prior best rating==3 | 0.006 | 0.042 | -0.029 | 0.004 | 0.043 | -0.035 |
| | (0.023) | (0.041) | (0.088) | (0.024) | (0.041) | (0.090) |
| Player's prior best rating==2 | -0.016 | -0.046 | 0.040 | -0.017 | -0.042 | 0.050 |
| | (0.032) | (0.057) | (0.097) | (0.032) | (0.055) | (0.098) |
| One or more competing 4-stars | 0.060** | 0.076 | 0.091 | 0.059** | 0.071 | 0.098 |
| | (0.028) | (0.092) | (0.169) | (0.029) | (0.084) | (0.152) |
| Pct. of contest elapsed | -0.091 | -0.638** | 0.828 | -0.088 | -0.631** | 0.750 |
| | (0.119) | (0.274) | (0.755) | (0.121) | (0.271) | (0.756) |
| Constant | 0.457** | 0.948*** | -0.956 | 0.448** | 0.900*** | -0.410 |
| | (0.209) | (0.317) | (0.742) | (0.209) | (0.309) | (0.905) |
| N | 2926 | 1557 | 879 | 2926 | 1557 | 879 |
| $R^2$ | 0.52 | 0.60 | 0.67 | 0.52 | 0.60 | 0.68 |
| Controls | Yes | Yes | Yes | Yes | Yes | Yes |
| Contest FEs | Yes | Yes | Yes | Yes | Yes | Yes |
| Player FEs | Yes | Yes | Yes | Yes | Yes | Yes |
| Forthcoming | No | No | No | Yes | Yes | Yes |
| Restriction | All | 2nd half | 4th qtr | All | 2nd half | 4th qtr |

Notes: Table shows the effects of 4-star feedback and competition on originality when no player has a 5-star rating. Observations are designs. Dependent variable is a continuous measure of a design's maximal similarity to previous entries in the same contest by the same player, taking values in [0,1], where a value of 1 indicates the design is identical to another. All columns include contest and player fixed effects and control for the time of submission and number of previous designs by the player and her competitors. Columns (4) to (6) additionally control for the best *forthcoming* rating on the player's not-yet-rated designs. The second and third columns in each block restrict the sample to submissions in the second half or fourth quater of a contest, when the absence of 5-star ratings may be more meaningful and is increasingly likely to be final. Similarity scores in this table are calculated using a perceptual hash algorithm. *, **, *** represent significance at the 0.1, 0.05, and 0.01 levels, respectively. Standard errors clustered by player in parentheses.

Table G.2: Similarity to any of player's previous designs: 4-vs-4 (d. hash)

| | (1) | (2) | (3) | (4) | (5) | (6) |
|---|---|---|---|---|---|---|
| Player's prior best rating==4 | 0.087 | 0.221* | 0.745*** | 0.084 | 0.238** | 0.715*** |
| | (0.068) | (0.122) | (0.253) | (0.068) | (0.120) | (0.238) |
| * 1+ competing 4-stars | -0.105** | -0.190** | -0.304 | -0.107** | -0.201** | -0.295 |
| | (0.047) | (0.084) | (0.201) | (0.047) | (0.079) | (0.186) |
| * prize value ($100s) | 0.006 | -0.007 | -0.162* | 0.006 | -0.012 | -0.152* |
| | (0.020) | (0.037) | (0.084) | (0.020) | (0.036) | (0.080) |
| Player's prior best rating==3 | -0.014 | -0.009 | -0.048 | -0.020 | -0.010 | -0.047 |
| | (0.023) | (0.039) | (0.070) | (0.024) | (0.040) | (0.074) |
| Player's prior best rating==2 | -0.016 | 0.049 | 0.058 | -0.021 | 0.049 | 0.056 |
| | (0.026) | (0.053) | (0.097) | (0.026) | (0.050) | (0.097) |
| One or more competing 4-stars | 0.045* | 0.109 | 0.111 | 0.047* | 0.111 | 0.136 |
| | (0.025) | (0.081) | (0.177) | (0.025) | (0.077) | (0.168) |
| Pct. of contest elapsed | -0.006 | -0.394 | 1.045 | -0.010 | -0.412 | 1.056 |
| | (0.101) | (0.267) | (0.864) | (0.100) | (0.265) | (0.859) |
| Constant | 0.517*** | 1.033*** | -1.077 | 0.513*** | 1.021*** | -1.084 |
| | (0.168) | (0.294) | (0.834) | (0.169) | (0.292) | (0.825) |
| N | 2926 | 1557 | 879 | 2926 | 1557 | 879 |
| $R^2$ | 0.55 | 0.61 | 0.70 | 0.55 | 0.61 | 0.71 |
| Controls | Yes | Yes | Yes | Yes | Yes | Yes |
| Contest FEs | Yes | Yes | Yes | Yes | Yes | Yes |
| Player FEs | Yes | Yes | Yes | Yes | Yes | Yes |
| Forthcoming | No | No | No | Yes | Yes | Yes |
| Restriction | All | 2nd half | 4th qtr | All | 2nd half | 4th qtr |

Notes: Table shows the effects of 4-star feedback and competition on originality when no player has a 5-star rating. Observations are designs. Dependent variable is a continuous measure of a design's maximal similarity to previous entries in the same contest by the same player, taking values in [0,1], where a value of 1 indicates the design is identical to another. All columns include contest and player fixed effects and control for the time of submission and number of previous designs by the player and her competitors. Columns (4) to (6) additionally control for the best *forthcoming* rating on the player's not-yet-rated designs. The second and third columns in each block restrict the sample to submissions in the second half or fourth quater of a contest, when the absence of 5-star ratings may be more meaningful and is increasingly likely to be final. Similarity scores in this table are calculated using a difference hash algorithm. *, **, *** represent significance at the 0.1, 0.05, and 0.01 levels, respectively. Standard errors clustered by player in parentheses.

# H Imitation of Competing Designs

Though players can see a gallery of competing designs in the same contest, they see only the distribution of feedback these designs have received – not the ratings provided to specific, competing entries – and should therefore not be able to use this information to imitate highly-rated competitors. The regressions in Table H.1 test this assumption by examining players' tendency to imitate competitors.

Table H.1: Similarity and change in similarity to competitors' best previously-rated designs

| | (1) | (2) | | (3) | (4) |
|---|---|---|---|---|---|
| Competitors' best==5 | -0.045** | -0.002 | Δ(Competitors' best==5) | -0.001 | 0.002 |
| | -0.022 | -0.024 | | -0.055 | (0.056) |
| * 1+ own 5-stars | -0.017 | -0.02 | * 1+ own 5-stars | 0.021 | 0.019 |
| | -0.026 | -0.027 | | -0.026 | (0.026) |
| * prize value ($100s) | -0.003 | -0.017*** | * prize value ($100s) | -0.008 | -0.009 |
| | -0.004 | -0.004 | | -0.01 | (0.010) |
| Competitors' best==4 | -0.002 | 0.005 | Δ(Competitors' best==4) | 0.036 | 0.036 |
| | -0.018 | -0.019 | | -0.039 | (0.039) |
| Competitors' best==3 | 0.016 | 0.022 | Δ(Competitors' best==3) | 0.041 | 0.041 |
| | -0.018 | -0.018 | | -0.038 | (0.038) |
| Competitors' best==2 | 0.01 | 0.013 | Δ(Competitors' best==2) | 0.049 | 0.050 |
| | -0.021 | -0.021 | | -0.039 | (0.039) |
| One or more own 5-stars | 0.004 | 0.009 | One or more own 5-stars | 0.007 | 0.010 |
| | -0.027 | -0.028 | | -0.006 | (0.006) |
| Pct. of contest elapsed | 0.078*** | 0.056** | Pct. of contest elapsed | 0.006 | 0.002 |
| | -0.007 | -0.022 | | -0.006 | (0.018) |
| Constant | -0.006 | -0.001 | Constant | 0 | 0.004 |
| | -0.072 | -0.074 | | -0.011 | (0.019) |
| N | 9709 | 9709 | N | 6065 | 6065 |
| $R^2$ | 0.44 | 0.44 | $R^2$ | 0.11 | 0.11 |
| Controls | No | Yes | Controls | No | Yes |
| Contest FEs | Yes | Yes | Contest FEs | Yes | Yes |
| Player FEs | Yes | Yes | Player FEs | Yes | Yes |

Notes: Table provides a test of players' ability to discern the quality of, and then imitate, competing designs. Observations are designs. Dependent variable in Columns (1) and (2) is a continuous measure of the design's similarity to the highest-rated preceding entries by other players, taking values in [0,1], where a value of 1 indicates that the design is identical to another. The mean value in the sample is 0.14 (s.d. 0.10). Dependent variable in Columns (3) and (4) is a continuous measure of the change in designs' similarity to the highest-rated preceding entries by other players, taking values in [-1,1], where a value of 0 indicates that the player's current design is equally similar to the best competing design as was her previous design, and a value of 1 indicates that the player transitioned fully from experimenting to copying (and a value of -1, the converse). The mean value of this variable is 0.00 (s.d. 0.09). In general, players are provided only the distribution of ratings on competing designs; ratings of specific competing designs are not observed. Results in this table test whether players can nevertheless identify and imitate leading competition. Columns (2) and (4) control for days remaining and number of previous designs by the player and her competitors. Similarity scores in this table are calculated using a perceptual hash algorithm. Preceding designs/ratings are defined to be those entered/provided at least 60 minutes prior to the given design. *, **, *** represent significance at the 0.1, 0.05, and 0.01 levels, respectively. Standard errors clustered by player in parentheses.

The first two columns of the table provide estimates from regressions of similarity to the highest-rated design by competing players on indicators for its rating. As in previous specifications, the top-rating indicator is interacted with the prize and with an indicator for whether the player herself also has a top-rated design in the contest. The latter columns repeat the exercise with first-differenced variants of the same specifications. There is little evidence in this table that players imitate highly-rated competitors in any systematic way – likely because they are simply unable to identify which competing designs are highly-rated. In unreported results, I replace the left-hand side with imitation of *any* competing design and similarly find no effect. The results establish that "creativity" in the presence of competition is not just imitation of competitors' designs. Table H.2 provides counterpart estimates using the difference hash algorithm, which suggest that if anything, players tend to deviate *away* from competitors' highly-rated work.

Table H.2: Similarity and change in similarity to competitors' best previously-rated designs (diff. hash)

| | (1) | (2) | | (3) | (4) |
|---|---|---|---|---|---|
| Competitors' best==5 | -0.199*** | -0.127*** | Δ(Competitors' best==5) | -0.06 | -0.058 |
| | -0.046 | -0.048 | | -0.11 | (0.110) |
| * 1+ own 5-stars | -0.016 | -0.022 | * 1+ own 5-stars | -0.006 | -0.011 |
| | -0.034 | -0.035 | | -0.066 | (0.066) |
| * prize value ($100s) | 0.029*** | 0.005 | * prize value ($100s) | 0.001 | -0.001 |
| | -0.009 | -0.009 | | -0.02 | (0.020) |
| Competitors' best==4 | 0.018 | 0.033 | Δ(Competitors' best==4) | 0.07 | 0.070 |
| | -0.035 | -0.035 | | -0.084 | (0.085) |
| Competitors' best==3 | 0.032 | 0.045 | Δ(Competitors' best==3) | 0.069 | 0.070 |
| | -0.034 | -0.034 | | -0.084 | (0.084) |
| Competitors' best==2 | -0.057 | -0.052 | Δ(Competitors' best==2) | 0.014 | 0.016 |
| | -0.04 | -0.04 | | -0.085 | (0.085) |
| One or more own 5-stars | -0.005 | 0.005 | One or more own 5-stars | 0.015 | 0.020 |
| | -0.027 | -0.03 | | -0.012 | (0.013) |
| Pct. of contest elapsed | 0.133*** | 0.138*** | Pct. of contest elapsed | -0.004 | 0.021 |
| | -0.014 | -0.04 | | -0.01 | (0.027) |
| Constant | 0.431*** | 0.045 | Constant | 0.04 | 0.017 |
| | -0.103 | -0.101 | | -0.066 | (0.071) |
| N | 9709 | 9709 | N | 6065 | 6065 |
| $R^2$ | 0.54 | 0.54 | $R^2$ | 0.14 | 0.15 |
| Controls | No | Yes | Controls | No | Yes |
| Contest FEs | Yes | Yes | Contest FEs | Yes | Yes |
| Player FEs | Yes | Yes | Player FEs | Yes | Yes |

Notes: Table provides a test of players' ability to discern the quality of, and then imitate, competing designs. Observations are designs. Dependent variable in Columns (1) and (2) is a continuous measure of the design's similarity to the highest-rated preceding entries by other players, taking values in [0,1], where a value of 1 indicates that the design is identical to another. The mean value in the sample is 0.33 (s.d. 0.21). Dependent variable in Columns (3) and (4) is a continuous measure of the change in designs' similarity to the highest-rated preceding entries by other players, taking values in [-1,1], where a value of 0 indicates that the player's current design is equally similar to the best competing design as was her previous design, and a value of 1 indicates that the player transitioned fully from experimenting to copying (and a value of -1, the converse). The mean value of this variable is 0.00 (s.d. 0.15). In general, players are provided only the distribution of ratings on competing designs; ratings of specific competing designs are not observed. Results in this table test whether players can nevertheless identify and imitate leading competition. Columns (2) and (4) control for days remaining and number of previous designs by the player and her competitors. Similarity scores in this table are calculated using a difference hash algorithm. Preceding designs/ratings are defined to be those entered/provided at least 60 minutes prior to the given design. *, **, *** represent significance at the 0.1, 0.05, and 0.01 levels, respectively. Standard errors clustered by player in parentheses.

# I Robustness Checks (4)

Though the body of the paper focuses on incumbents' creativity, a natural question is whether similar effects are present for new entrants. Measuring the originality of a player's first submission is mechanically more challenging, as there are no prior own submissions in the same contest against which to compare. Indeed, the concept of creativity itself in this case is not well-defined, as there is no clear precedent against which to compare. One possibility is to compare against that player's work in other contests, but each project is sufficiently distinct that we will rarely see high values of similarity in this measure. Another option is to compare against competitors' work in the same contest, but the evidence from the paper shows that this similarity is almost never high in the first place (see Table 3).

To make headway on this question, I invoke the batch-level measure of originality. Recall that players sometimes enter designs in short bursts, which this papers terms "batches", and within these batches can choose to make their entries similar to or different from each other. Occasionally, we see players submit a batch of designs in quick succession when they first enter a contest – presenting an opportunity to evaluate intra-batch experimentation by entrants facing different levels of competition.

Table I.1 regresses intra-batch similarity for initial batches of 2+ designs on indicators for the highest-rated competition that was present at the time. The first two columns do so for initial batches by any player, while the latter columns explore heterogeneity across players with different levels of experience: Columns (3) and (4) restrict to those who have never won a contest; Columns (5) and (6) restrict to those who have. Even-numbered columns add the standard controls used throughout the paper.

Qualitative evidence of an effect is present across all specifications: intra-batch similarity decreases monotonically with higher-quality competition. The magnitudes are substantial: with low-rated (1-star) competition, intra-batch similarity is on average 0.78 (the constant), whereas with high-rated (5-star) competition it is a full standard deviation (0.31) lower. The pattern is especially pronounced for less-experienced entrants, though it is also present (with smaller magnitudes) for the more experienced players. However, we must note that the restriction to contest-players who enter with batch submissions reduces the sample size to a point where standard errors are too large to statistically rule out zero.

Table I.1: Experimentation within initial batches (p. hash)

| | (1) | (2) | (3) | (4) | (5) | (6) |
|---|---|---|---|---|---|---|
| Best competing rating==5 | -0.314 | -0.326 | -0.855 | -0.839 | -0.122 | -0.135 |
| | (0.302) | (0.309) | (0.552) | (0.532) | (0.355) | (0.390) |
| Best competing rating==4 | -0.241 | -0.240 | -0.790 | -0.538 | -0.026 | -0.039 |
| | (0.294) | (0.293) | (0.559) | (0.775) | (0.337) | (0.356) |
| Best competing rating==3 | -0.047 | -0.048 | -0.566 | -0.529 | 0.148 | 0.145 |
| | (0.310) | (0.312) | (0.584) | (0.556) | (0.333) | (0.340) |
| Best competing rating==2 | 0.160 | 0.164 | -0.204 | 0.102 | 0.283 | 0.278 |
| | (0.410) | (0.414) | (0.559) | (0.855) | (0.491) | (0.500) |
| Constant | 0.781** | 0.776** | 1.089 | 0.814 | 0.568 | 0.567 |
| | (0.370) | (0.377) | (0.818) | (0.930) | (0.413) | (0.416) |
| N | 826 | 826 | 329 | 329 | 496 | 496 |
| $R^2$ | 0.83 | 0.83 | 0.97 | 0.97 | 0.84 | 0.84 |
| Controls | No | Yes | No | Yes | No | Yes |
| Contest FEs | Yes | Yes | Yes | Yes | Yes | Yes |
| Player FEs | Yes | Yes | Yes | Yes | Yes | Yes |
| Experience | All | All | Inexp. | Inexp. | Exp. | Exp. |

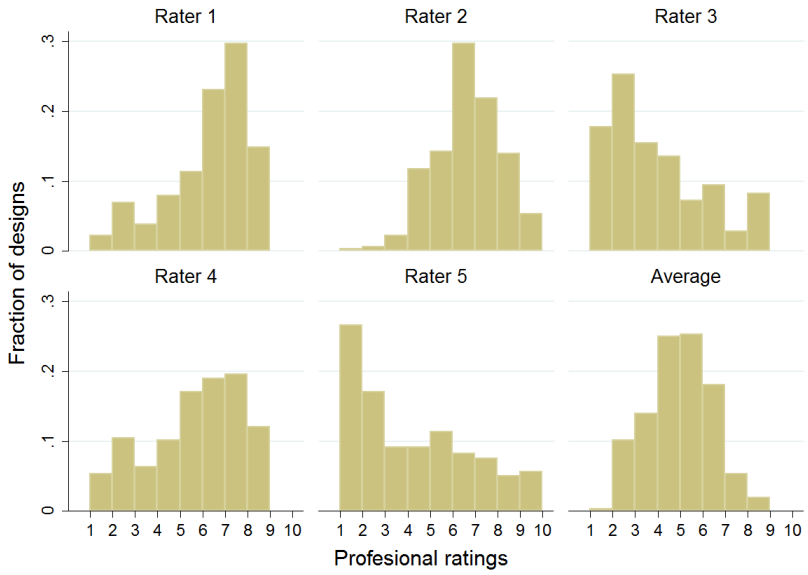
Notes: Table shows new entrants' tendency to experiment within initial batches as a function of the competition observed at the time. The sample is restricted to players who enter a contest with two or more submissions in quick succession, and the dependent variable measures the similarity within these batches, taking values in [0,1], where a value of 1 indicates that two designs in the batch are identical. Columns (3) and (4) restrict the sample to inexperienced players, while Columns (5) and (6) restrict to more experienced players, to highlight heterogeneity. Even-numbered columns add controls for the time of submission and number of previous designs entered by the player's competitors. Similarity scores in this table are calculated using a perceptual hash algorithm. *, **, *** represent significance at the 0.1, 0.05, and 0.01 levels, respectively. Standard errors clustered by player in parentheses.

# J Collection of Professional Ratings

The panelists that participated in the ratings exercise were recruited through the author's personal and professional networks and hired at their freelance rates. All have formal training and experience in commercial graphic design, and they represent a diverse swath of the profession: three panelists work at advertising agencies, and two others are employed in-house for a client and primarily as a freelancer (respectively).

Ratings were collected though a web-based application created and managed on Amazon Mechanical Turk. Designs were presented in random order and panelists were limited to 100 ratings per day. With each design, the panelist was provided the project title and client industry (as they appear in the design brief in the source data) and instructed to rate the "quality and appropriateness" of the given logo on a scale of 1 to 10. Panelists were asked to rate each logo "objectively, on its own merits" and not to "rate logos relative to others." Figure J.1 provides the distribution of ratings from each of the five panelists and the average.

Figure J.1: Panelists' ratings on subsample of sponsors' top-rated designs




Notes: Figure shows the distribution of professionals' ratings on all 316 designs in the dataset that received the top rating from contest sponsors. Professional graphic designers were hired at regular rates to participate in this task. Each professional designer provided independent ratings on every design in the sample rated 5 stars by a contest sponsor. Ratings were solicited on a scale of 1-10, in random order, with a limit of 100 ratings per day.

It can be seen in the figure that one panelist ("Rater 5") amassed over a quarter of her ratings at the lower bound, raising questions about the reliability of these assessments: it is unclear what the panelist intended with these ratings, why such a high proportion was given the lowest rating, and whether the panelist would have chosen an even lower rating had the option been available. The panelist's tendency to assign very low ratings became apparent after the first day of her participation, and in light of the anomaly, the decision to omit this panelist's ratings from the analysis was made at that time. The results of the paper are nevertheless robust to including ratings from this panelist that lie above the lower bound.